

Short Sales and Price Discovery in the Hong Kong Stock Market



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Declaration

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To my parents and beloved wife.

Abstract

This thesis explores the effect of short sales on the process of price discovery and the determinants of the speed of price adjustment in the Hong Kong stock market over the period 2001-2011 using an ultra-high frequency dataset.

The thesis extends the VAR model of [Hasbrouck \(1991\)](#) using four estimation methods: OLS and WLS in which the models are estimated on an equation-by-equation basis; VAR and WVAR in which correlations between residuals in both equations are taken into account. To deal with conditional heteroscedasticity present in high frequency data, the thesis also uses GARCH and BEKK models. More detailed analyses cover key determinants of the speed of price adjustment: market conditions, firm size, trading volume and short sellers' activities. The thesis employs a number of non-standard statistical procedures.

The results show that price discovery is affected by short selling for a substantial minority of stocks, but not all. For those affected, short sales accelerate the speed of price adjustment by decreasing the autocorrelations in quote returns and trades. A smaller absolute value of autocorrelation indicates a more efficient process of price discovery. Trade continuity (autocorrelation in quote returns) is affected most when stocks are added to (removed from) the list of stocks eligible for short sales. The results of the more detailed analysis are as follows. Short sales improve price efficiency regardless of market condition. Stocks adjust their prices more quickly in a down market than an up market. The stock prices of larger firms adjust more quickly to new information than smaller ones. Medium-sized (small-sized) firms

are the most affected when they are added to (removed from) the short selling list. Higher trading volumes are associated with quicker speed of incorporation of new information into prices. Short sellers play an important role in the process of price discovery.

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Chapter 1

Introduction

1.1 Overview

Short selling refers to the practice of selling stocks without owning them by borrowing them from a lender and then returning them at a future date by purchasing them back in the market at that time. Short sellers can benefit from buying back the borrowed stocks if prices go down but suffer losses if prices go up. Short sales provide an alternative trading option for traders with negative expectation of future performance as traders may face difficulties in selling stocks directly if they have bad news.

Since the US Securities and Exchange Commission (SEC) took emergency regulatory action in respect of short sales, banning short sales in the securities of financial institutions temporarily in the recent global financial crisis, the benefits and costs of this ban have become a topic of debate and discussion among investors, regulators and academics. The common belief is that short sales may be problematic when the market lacks confidence and can be used as a tool to mislead other participants resulting in a downward manipulation. In such circumstances, a manipulator short sells the shares of a company and then spreads rumours about the pessimistic prospects of the company. This could be harmful to the company, investors and the entire market. In order to stabilise the market and rebuild investors' confidence in

the crisis period, market regulators take various regulatory actions on short sales including restricting naked short selling¹ but permitting covered short selling or prohibiting short sales of some particular securities such as financial stocks.

However, there are divergent views, particularly from researchers, about the benefits of bans or constraints on short sales. Many researchers argue that the responses of market regulators are unprecedented and the actions to ban or limit short sales are out of line with the objectives and principles of market regulation if regulators aim to ensure an efficient market in the midst of the crisis. It is believed that short selling is the reflection of investors' negative expectation of price changes in the future and not all information will be fully reflected on prices if short sales are banned or restricted. In the presence of short sales constraints price transmissions are asymmetric, which can cause overvaluation (Miller, 1977) and a delay in price adjustment to new information (Diamond & Verrecchia, 1987). Given heterogeneous beliefs among investors and no restrictions on short sales, Miller (1977) believes that an equilibrium stock price can be jointly set by optimistic investors who buy long and pessimistic investors who sell short. Stocks will be overpriced when there are binding short sales constraints as pessimists are restricted from trading on their beliefs. Many empirical studies find subsequent significant negative stock returns upon the lifting of short sales bans/constraints, which largely support the overvaluation view (Bris *et al.*, 2007; Chang *et al.*, 2007, 2014; Diether *et al.*, 2009; Figlewski, 1981; Jones & Lamont, 2002).

In the seminal work of Miller (1977), investors are assumed not to acquire information from stock prices. Using a rational expectations model², Diamond & Verrecchia (1987) provide an alternative view and argue that rational investors can recognise the existence of short sales constraints. Therefore, they are able to adjust the potential bias caused by short sales constraints so that observed stock

¹Unlike traditional (covered) short sales, naked short sales occur when the short seller sells the security first without ever borrowing the security.

²People with rational expectations will make decisions based on their rational outlook, all current available information, and past experiences. The current economy expectations are equivalent to what the future economy will be.

prices will represent an unbiased expectation based on all available information and no overpricing of stocks will exist. In addition, their model indicates that the prices of stocks without the eligibility of short sales will have a delayed speed of adjustment to unfavourable private information as short sellers' activities are limited in the market. Due to the lack of data and to difficulties in characterising the speed of price adjustment, only a few papers empirically test the prediction of [Diamond & Verrecchia \(1987\)](#), namely that short sales constraints slow down the process of price discovery ([Beber & Pagano, 2013](#); [Boehmer & Wu, 2013](#); [Bris et al., 2007](#); [Chang et al., 2014](#); [Chen & Rhee, 2010](#); [Saffi & Sigurdsson, 2010](#)). Another key empirical issue is to determine an appropriate measure of short sales constraints. Most studies use indirect measurements as the proxies for short sales constraints including short interest ([Figlewski, 1981](#)); breadth of ownership³ ([Chen et al., 2002](#)); institutional ownership ([Nagel, 2005](#)); costs of short sales ([Jones & Lamont, 2002](#); [Saffi & Sigurdsson, 2010](#)); option introductions ([Danielsen & Sorescu, 2001](#)). Besides these, [Chang et al. \(2007\)](#) use institutional changes and provide a direct examination on the effect of short sales constraints in the Hong Kong stock market. Stocks in the Hong Kong market are allowed for short sales only when they are added to the designated list (called the D-list). This list is revised on a quarterly basis, with stocks being deleted from it as well as added to it. This rare institutional environment allowed [Chang et al. \(2007\)](#) to trace changes in price movements before and after stocks are added to or removed from the list for short sales.

Taken together, the main purpose of this thesis is to empirically explore the effect of short sales on the process of price discovery as well as the determinants of the speed of price adjustment during both addition and deletion events in the Hong Kong stock market.

³Breadth is defined as the number of investors with long positions in a particular stock. Lower breadth of a stock is associated with more investors sitting on the sidelines with pessimistic opinions not registered in the stock's price.

1.2 Institutional Background

1.2.1 Short Sales in Hong Kong

The Hong Kong Stock Exchange [HKEx henceforth] introduced short selling in January 1994. Different from the US market where most of the stocks are shortable, stocks on the HKEx are allowed for short selling only when they are added to the D-list by meeting certain requirements⁴. The selection criteria for inclusion in the D-list for short selling include: (i) all constituent stocks of the underlying indices with financial derivatives written and traded on the HKEx; (ii) underlying stocks of individual stock options and futures; (iii) stocks with a minimum public flotation of HK\$10 billion, a minimum market capitalisation⁵ of HK\$3 billion, or a minimum annual turnover ratio⁶ of 50%. Stocks that meet at least one requirement for at least 3 months can be added to the D-list which is revised quarterly. The HKEx provides daily dissemination of short sales information including daily short selling trading volume.

The HKEx is an order driven market without market makers. Orders are submitted to the Automatic Order Matching and Execution System (AMS) and executed by matching the price and in order of time of arrival. The regulations for stock borrowing and lending require the traders to show their collateral which is a minimum of 105% of the borrowed stocks' market value. The practise for short sales, except short sales for stock futures, follows the up-tick rule that a short sale cannot be made below the best current ask price. Naked or uncovered short sales without stock borrowing arrangements are forbidden by the exchange and are considered to be a criminal offence. When placing short sales orders in the AMS system, all transactions are identified to the HKEx by the broker and the information is available to the public through the limit order book by flagging the short orders. A ledger with specific details about each short selling order is required to be available

⁴The latest selection criteria for short selling is the updated version in July 2012 and full details can be found at http://www.hkex.com.hk/eng/market/sec_tradinfo/regshortsell.htm.

⁵Market capitalisation is the total number of issued shares times the latest share price.

⁶Annual turnover is the aggregate turnover during the preceding 12 months.

to the HKEx at any time.

1.2.2 Simultaneous Entry to the Hang Seng Index

The Hang Seng Index [HSI henceforth] was launched on 24 November 1969 and it measures the performance of largest and most liquid companies listed in Hong Kong. The number of constituents is fixed at 50 for the index. To be eligible for selection, a company: *(i)* must be among those companies that constitute the top 90% of the total market capitalisation of all primary listed shares on the HKEx (market capitalisation is expressed as an average of the past 12 months); *(ii)* must be among those companies that constitute the top 90% of the total turnover of all primary listed shares on the HKEx (turnover is aggregated and individually assessed for eight quarterly sub-periods for the past 24 months); and *(iii)* should normally have a listing history of 24 months on the HKEx or meet the requirements of the guidelines for handling large-cap stocks listed for less than 24 months. From the many eligible candidates, final selections are based on the following: *(i)* the market capitalisation and turnover ranking of the company; *(ii)* the representation of the relevant sub-sector within the HSI directly reflecting that of the market; and *(iii)* the financial performance of the company.

[Insert Table 1.1 about here]

Table 1.1 shown above reports the number of changes to the D-list with respect to simultaneous entry to the HSI. As shown in the table, there are 1,296 addition and 810 deletion events during the 10-year study period. Regarding the simultaneous entry to the HSI and the D-list, there is no stock being added to/removed from the D-list and the HSI on the same effective date. Furthermore, there are only 3 out of 1,296 stock/events (0.23%) being added to the HSI around (one month before and after) the effective date of being shorable. Therefore, the effect of entering into the index can be negligible and the changes in the speed of price adjustment

are mainly attributed to the changes of the D-list not the index.

1.3 Motivations, Research Questions and Summary

As discussed earlier, direct evidence on the asymmetric speed of price adjustment is scarcer due to the limitation of data availability and the difficulty of measuring the speed. To fill this gap, Chapter 2 extends a bivariate vector autoregressive [VAR henceforth] model of [Hasbrouck \(1991\)](#) and aims to offer a direct examination of the effect of short sales on the speed of price adjustment based on ultra-high frequency trade-by-trade data over a 10-year period from 2001 to 2011 in the Hong Kong stock market. This VAR approach has been applied to high-frequency data to measure informational content of stock trades and the price efficiency. [Chen & Rhee \(2010\)](#) use the VAR model containing quote returns and signed trade indicators to measure the respective speed of price adjustment between shortable and non-shortable stocks. They find that stocks have greater autocorrelations in quote returns and less trade continuity indicating a quicker speed of price adjustment when they are added to the short selling list. On the contrary, [Boehmer & Wu \(2013\)](#) argue that an efficient process of price discovery should follow a random walk in association with a smaller autocorrelations. They find that more shorting flows speed up the incorporation of new information into prices by weakening the autocorrelations in quote returns.

Chapter 2 contributes to the existing literature in the following ways. First, this chapter extends the VAR model by adding a constant to both equations in the model. This modification removes a potential internal inconsistency in the standard implementation. Secondly, different from [Chen & Rhee \(2010\)](#)'s model estimation on a daily basis, the VAR model in this chapter is estimated for each stock/event based on the transactions in a 60-day period before and after the addition or deletion event. Thirdly, this chapter focusses not only on individual model parameters and summaries thereof, but also on their dynamics. Consideration of

changes provides deeper insights into the effect of short sales on the process of price discovery during both addition and deletion events. Fourthly, this chapter also considers the variable time duration of consecutive trades during the model estimation as information is conveyed by trades as well as by time (Diamond & Verrecchia, 1987; Easley & O'Hara, 1992). Moreover, a *bona fide* VAR model as well as its time-weighted form are also employed with consideration of cross-sectional correlation of the residuals. Thus, there are four estimation methods: namely ordinary least squares, weighted least squares, VAR and weighted VAR (henceforth denoted OLS, WLS, VAR and WVAR respectively). Finally, the chapter employs a number of non-standard statistical procedures required to deal the effect of unequal stock specific volatility in the hypothesis testing procedures.

The findings from Chapter 2 show that a small number of stocks do exhibit significant changes including both increases and decreases in the magnitude of estimated parameters. The overall results indicate that short sales accelerate the speed of price adjustment by decreasing the autocorrelations in both quote returns and trades in addition events; these findings being consistent with those from Boehmer & Wu (2013), namely that short sales contribute to a more efficient process of price discovery with a smaller autocorrelations in quote returns which is more associated with a random walk. The study of parameter dynamics reports that there is a great consistency for the majority of model parameters; that is, parameters which, for example, are statistically significantly negative before an event tend to remain in that category after it. Comparing OLS and WLS, it is noted that the VAR and WVAR models have a better ability to capture residual correlations as well as the effect of short sales on model parameters. The WVAR model is considered to be the preferred model as it allows for variable times between trades and is preferable for theoretical reasons.

Following the studies in Chapter 2 and empirical evidence from other studies in financial economics using high frequency data, it is believed that high frequency quotations and trades in the VAR model invariably exhibit heterogeneity of variance. Chapter 3 therefore carries an investigation into the effect of short sales on

the process of price discovery taking into account time-varying volatility and covariance. The contribution of this chapter is summarised as follows. To the best of our knowledge, the study in Chapter 3 is the first to examine changes in the speed of price adjustment affected by short sales with consideration of heteroscedasticity. The Generalised ARCH [GARCH henceforth] model developed by [Bollerslev \(1986\)](#) is applied to each equation of the VAR model to capture the time-varying variance from individual quote returns or signed volumes. The VAR model is also estimated using the BEKK model of [Engle & Kroner \(1995\)](#), which takes account of possible volatility spillover between quote returns and signed volumes as well as time-varying volatilities. The BEKK model is both parsimonious and ensures the positive semi-definiteness of the time-varying covariance matrix.

The results of Chapter 3 are as follows. First, compared with the results under the least squares based models of Chapter 2, more stocks exhibit significant changes including both increases and decreases in the absolute value of model parameters under the GARCH and BEKK models. Overall, short sales accelerate the speed of price adjustment when time-varying variances and covariances are considered, but are consistent with the results based on models of Chapter 2. Secondly, combining the results from Chapters 2 & 3, it is found that trade continuity (autocorrelation in quote returns) is the most affected during addition (deletion) events. Thirdly, the combined results from six estimation models reveal that the VAR models estimated simultaneously and taking heterogeneous variances and covariances into account (WVAR and BEKK) are more powerful to capture the changes in the speed of price adjustment than the other four models.

The purpose of Chapter 4 is to explore how the speed of price adjustment changes by controlling market condition, firm size, trading volume and short sellers' activities. This chapter makes several contributions. First, it introduces an extension to the VAR model interacted with a dummy variable of current and one-day lagged market returns to examine the speed of price adjustment under different market conditions. Furthermore, the model interacted with the dummy variable under the GARCH framework is also used to deal with the issue of heteroscedasticity. Secondly, the estimation results are divided into two periods, crisis and non-crisis,

in order to examine the effect of short sales on price efficiency during the extreme market conditions of 2007-2009. Thirdly, the models are re-estimated using by splitting that data into two sets: days in the 60-day period with positive market returns and non-positive market returns, respectively. Lastly, this chapter compares the differences in the speed of price adjustment directly according to stocks' market capitalisation, trading volume and shorting activities.

The findings in Chapter 4 are as follows. Short sales enhance price efficiency regardless of market conditions and stocks are faster in response to new information when the market is bearish. By looking at the recent global financial crisis, it is found that short sales contribute to the efficient process of price discovery only in the non-crisis period. Mixed results during the crisis period indicate that the role of short sales in extreme market conditions is ambiguous. For the size effect, it is reported that large firms react more quickly to new information than small firms during both addition and deletion events. Medium (small) firms are the most affected during the addition (deletion) events. For the volume effect, stocks with more trading volumes are more likely to be affected during the addition events. The results for the deletion events are mixed as it shows that stocks with highest and lowest trading volumes are the most affected ones. For both events, it is noted that stocks with higher trading volumes have faster speed of price adjustment than those with lower trading volumes and the results are consistent under the models when the time duration is not considered. For the effect of short sellers, the results from using short interest and shorting flow as explanatory variables indicate that the speed of price adjustment is faster when short sellers are more active. It therefore supports the view that short sellers play a prominent role to improve price efficiency in the market.

1.4 Structure of the Thesis

This thesis consists of four further chapters, as follows. Chapter 2 contains the tests of hypothesis for short sales and price discovery in the Hong Kong stock market

using the VAR model based on least squares type methods. Chapter 3 explores the effect of short sales by estimating the VAR model under GARCH and BEKK frameworks. Chapter 4 conducts the detailed analysis of the impact of market conditions, firm size, trading volume and short sellers' activities on the speed of price adjustment. Chapter 5 summarises the findings of the thesis and discusses limitations and directions for future research. The tables and the appendices are at the end of each chapter.

**Table 1.1: The Records of Addition and Deletion Events
of the D-list and the HSI**

The table reports the revision history of short selling list on the HKEx and the corresponding records for the HSI during the study period of 2001-2011. Columns are as follows: 1. Announcement date; 2. Effective date; 3. No. of additions; 4. No. of deletions; 5. No. of stocks on the list after each revision; 6. No. of additions entry to the HSI; 7. No. of deletions quit from the HSI.

1	2	3	4	5	6	7	1	2	3	4	5	6	7
03/05/2001	14/05/2001	6	0	280	1	0	14/05/2007	21/05/2007	29	14	451	0	0
07/08/2001	20/08/2001	9	11	278	1	0	18/05/2007	21/05/2007	1	0	452	0	0
27/11/2001	03/12/2001	17	85	210	0	0	22/05/2007	29/05/2007	1	0	453	0	0
05/02/2002	25/02/2002	7	14	203	0	0	27/06/2007	04/07/2007	1	0	454	0	0
08/05/2002	21/05/2002	11	6	208	0	0	16/07/2007	17/07/2007	1	0	455	0	0
17/07/2002	29/07/2002	24	5	227	0	0	06/08/2007	13/08/2007	134	9	580	0	0
28/11/2002	29/11/2002	6	15	218	1	0	20/08/2007	27/08/2007	1	0	581	0	0
21/01/2003	27/01/2003	5	7	216	0	0	19/11/2007	26/11/2007	64	23	622	0	0
07/05/2003	19/05/2003	18	7	227	0	0	07/12/2007	14/12/2007	2	0	624	0	0
14/07/2003	21/07/2003	1	16	212	0	0	11/12/2007	14/12/2007	1	0	625	0	0
27/10/2003	03/11/2003	36	5	243	0	0	11/02/2008	18/02/2008	33	41	617	0	0
05/01/2004	06/01/2004	1	0	244	0	0	11/03/2008	13/03/2008	1	0	618	0	0
03/02/2004	10/02/2004	29	3	270	0	0	05/05/2008	13/05/2008	22	47	593	0	0
06/04/2004	07/04/2004	1	0	271	0	0	14/05/2008	15/05/2008	1	0	594	0	0
20/04/2004	27/04/2004	26	4	293	0	0	02/06/2008	03/06/2008	5	0	599	0	0
25/06/2004	01/07/2004	1	0	294	0	0	31/07/2008	07/08/2008	10	51	558	0	0
05/07/2004	09/07/2004	1	0	295	0	0	07/11/2008	14/11/2008	6	144	420	0	0
26/07/2004	02/08/2004	8	21	282	0	0	05/02/2009	12/02/2009	25	27	418	0	0
01/11/2004	08/11/2004	9	11	280	0	0	07/05/2009	14/05/2009	13	22	409	0	0
28/01/2005	07/02/2005	15	7	288	0	0	03/07/2009	10/07/2009	1	0	410	0	0
28/02/2005	01/03/2005	2	0	290	0	0	29/07/2009	05/08/2009	49	16	443	0	0
09/05/2005	17/05/2005	37	9	318	0	0	29/10/2009	05/11/2009	58	11	490	0	0

Table continued on the following page.

Table 1.1: Continued from previous page

1	2	3	4	5	6	7	1	2	3	4	5	6	7
04/07/2005	08/07/2005	1	0	319	0	0	11/11/2009	18/11/2009	1	0	491	0	0
11/07/2005	15/07/2005	1	0	320	0	0	30/11/2009	03/12/2009	1	0	492	0	0
04/08/2005	15/08/2005	14	12	322	0	0	08/12/2009	15/12/2009	1	0	493	0	0
26/08/2005	05/09/2005	1	0	323	0	0	21/12/2009	24/12/2009	1	0	494	0	0
24/10/2005	28/10/2005	1	0	324	0	0	25/01/2010	01/02/2010	65	8	551	0	0
18/11/2005	17/11/2005	11	7	328	0	0	22/02/2010	01/03/2010	1	0	552	0	0
14/02/2006	20/02/2006	10	8	330	0	0	03/03/2010	10/03/2010	1	0	553	0	0
27/02/2006	01/03/2006	2	0	332	0	0	18/03/2010	25/03/2010	1	0	554	0	0
19/05/2006	29/05/2006	23	17	338	0	0	03/05/2010	10/05/2010	59	12	601	0	0
29/05/2006	02/06/2006	1	0	339	0	0	12/07/2010	16/07/2010	1	0	602	0	0
01/06/2006	02/06/2006	1	0	340	0	0	28/07/2010	04/08/2010	40	19	623	0	0
18/08/2006	25/08/2006	38	10	368	0	0	23/08/2010	30/08/2010	1	0	624	0	0
24/08/2006	01/09/2006	1	0	369	0	0	25/10/2010	29/10/2010	47	18	653	0	0
19/10/2006	23/10/2006	1	0	370	0	0	11/11/2010	15/11/2010	1	0	654	0	0
19/10/2006	27/10/2006	1	0	371	0	0	17/11/2010	22/11/2010	2	0	656	0	0
24/11/2006	01/12/2006	55	9	417	0	0	17/12/2010	20/12/2010	1	0	657	0	0
26/02/2007	05/03/2007	30	24	423	0	0	23/12/2010	30/12/2010	1	0	658	0	0
13/03/2007	14/03/2007	1	0	424	0	0	12/01/2011	28/01/2011	1	0	659	0	0
12/04/2007	19/04/2007	5	0	429	0	0	25/01/2011	01/02/2011	1	0	660	0	0
25/04/2007	26/04/2007	6	0	435	0	0	18/02/2011	25/02/2011	70	17	713	0	0
25/04/2007	27/04/2007	1	0	436	0	0	17/05/2011	24/05/2011	65	18	760	0	0

Chapter 2

Tests of Hypothesis for Short Sales and Price Discovery using the VAR Model

2.1 Introduction

The impact of short sales on market efficiency is a much debated topic in finance and the debate has been reinvigorated by the financial crisis of 2007-09. Many countries⁷ and capital markets around the world imposed restrictions during that period, or even banned short selling (Beber & Pagano, 2013). In a bear market, short sellers are criticised for exacerbating the market decline, causing market panics and manipulating stock prices. Regulators believe that stricter regulations on short sales can help to protect the market's integrity and quality and rebuild investors' confidence. In sharp contrast, many researchers voice opposition to short selling restrictions. They believe that short selling is the reflection of some market participants' negative expectation of future price changes. In the presence of short sales constraints, price transmissions are asymmetric, which can cause overvalua-

⁷After Lehman Brother's bankruptcy in late September 2008, short sales on financial stocks are temporarily prohibited in the United States (from September to October in 2008) and in the United Kingdom (from September 2008 to January 2009), respectively. Countries such as Germany, Spain and France also adopted temporary short sales bans in that period.

tion (Miller, 1977) and a delay in price adjustment to new information (Diamond & Verrecchia, 1987). Existing studies tend to focus on the relationship between short sales and overvaluation (Bris *et al.*, 2007; Chang *et al.*, 2007, 2014; Diether *et al.*, 2009; Figlewski, 1981; Jones & Lamont, 2002). However, only a few papers examine the speed of price adjustment to new information. This is apparently due to the lack of data and to difficulties in characterising the speed of price adjustment. An important paper concerned with price adjustment is due to Chen & Rhee (2010) who use a bivariate VAR model (Hasbrouck, 1991) to measure the respective speed of price adjustment between shortable and non-shortable stocks. Their model includes quote-midpoint changes and signed trade indicators, which are used to describe autocorrelations in quote revision and trade continuity respectively. They find that shortable stocks have a faster speed of price adjustment to new information by exhibiting more negative autocorrelation in mid-quote returns and less positive trade continuity. Boehmer & Wu (2013) define the mid-quote return as the return on the mid-price for two successive quotations in the VAR model to capture price efficiency and they believe that an efficient price process should follow a random walk. Their results show that a greater shorting flow (that is, a larger number of short selling orders) is associated with smaller pricing errors and smaller mid-quote return autocorrelations, which indicate faster price discovery.

This chapter extends the VAR model of Hasbrouck (1991) and offers a direct examination of the impact of short sales on the speed of price adjustment to new information using high-frequency data over a 10-year period from May 2001 to May 2011 on the HKEx. The HKEx only allows stocks which meet certain requirements to be available for short selling. Furthermore, stocks may be added to and deleted from the short selling list (known as the D-list), which is revised on a quarterly basis. Addition and deletion events may occur multiple times for a given stock during the study period. The D-list and the implied set of addition and deletion events offers the means to examine the effect of short-selling on price discovery.

This chapter also makes six methodological contributions to the existing literature. First, the changes in quotes and trades dynamics in the bivariate VAR model of Hasbrouck (1991) capture the information revealed by trades. The model has been

widely used to examine the effects of trade-related information (Chen & Rhee, 2010; Chung *et al.*, 2005; Dufour & Engle, 2000). This chapter notes an internal inconsistency in the standard implementation of the VAR model which may lead to a violation of its assumptions. To solve this potential problem, the VAR model is modified by adding a constant term to both equations. Secondly, Chen & Rhee (2010) estimate the VAR model on a daily basis. Examination of the data used in this chapter reveals that during the 60-day period before and after addition and deletion events over a 10-year study period, there are respectively 22,246 days out of 124,202 days (17.9%) and 30,566 days out of 78,338 days (39.0%) with less than 20 transactions. Estimating a small number of transactions on a daily basis could cause multi-collinearity and other statistical problems. Therefore, in this chapter, and differently from Chen & Rhee (2010), the VAR model is estimated for each stock for each addition and deletion event using all transactions in a 60-day period both before and after the events. Thirdly, Chen & Rhee (2010) focus on the estimated coefficients for lagged quote revisions and signed trade indicators. This chapter analyses the dynamics not only of the individual parameters but also the sums of parameters in the model. This gives a clearer and more detailed picture of how the price discovery process is affected when a stock is either added to or deleted from the D-list. Fourthly, this chapter also examines the informational role of trades by taking consideration of the time duration between two consecutive transactions. Information is conveyed by trades as well as by time (Diamond & Verrecchia, 1987; Easley & O'Hara, 1992). Informed traders will always trade unless they do not have stocks and information or there are short-sale constraints in the market. Therefore, it is reasonable to assume that variations of trade time duration are associated with the behaviour of informed traders. Using the standard assumption that volatility is proportional to the square root of the length of a time interval provides a better understanding of the information impact of a trade on price. Fifth, in addition to the use of least squares and weighted least squares on an equation by equation basis, this chapter also uses the *bona fide* VAR model in which the cross-sectional correlation of the residual terms is considered. Lastly, hypothesis tests in this chapter explicitly recognize that the estimated standard errors of comparable estimated coefficients in the VAR model for different stocks are not the same. They are proportional to the estimated residual or stock specific standard errors. Tests of significance based on standard *t*-tests are therefore not

appropriate as they are not robust to variance heterogeneity and may lead to misleading inferences. Instead this chapter uses procedures based on a robust version of the t -test originally due to [Scott & Smith \(1971\)](#) [SS test henceforth].

The main results are summarised as follows. First, different from those in [Chen & Rhee \(2010\)](#), there is a significant decrease (increase) in quote autocorrelations which measure the speed of price adjustment for stocks during the addition (deletion) events. An efficient price process implies that return of quote midpoints follow a random walk, thus a smaller absolute value of autocorrelations caused by short sales indicates a faster speed of price adjustment. Quote revision autocorrelations, which are used as a proxy for speed of price adjustment in the previous empirical studies ([Boehmer & Wu, 2013](#); [Chen & Rhee, 2010](#)), are observed to decrease (increase) during the addition (deletion) events. [Chen & Rhee \(2010\)](#) suggest that stronger quote reversals led by weaker trade continuity reveal a quicker speed of price adjustment to new information while [Boehmer & Wu \(2013\)](#) argue that weaker autocorrelations in quote revision indicates greater price efficiency under random walk theory. In this chapter, the autocorrelations in quote revision become weaker (stronger) for addition (deletion) events and this finding on speed of price adjustment is consistent with those from [Boehmer & Wu \(2013\)](#). It is also found that trade continuity becomes weaker when stocks are allowed for short selling. This suggests that, in general, short sales contribute to price discovery for addition and deletion events. Secondly, a study of the dynamics indicates that parameters in the model generally remain consistent as far as sign is concerned; that is, an estimated model parameter which is significantly negative (positive) before the addition/deletion event will remain negative (positive) after the event. Thirdly, the results of statistical test on the difference in frequency of changes suggest that significant changes in the price impact of trades are less visible when time durations are considered in the estimation process and more significant changes in the autocorrelations in quote returns and trades are captured if two equations in the model are estimated simultaneously by VAR or by WVAR. Finally, it is found that WVAR model with consideration of time duration between trades has a stronger ability to capture the residual correlations in the model and the effect of short sales

on price efficiency during the addition and the deletion events.

The chapter is organised as follows. Section 2.2 presents a brief literature review of short sales and their impact on the speed of price adjustment. Section 2.3 introduces the extended version of the VAR model and the non-standard hypothesis tests that are used in the chapter. Section 2.4 describes the data in this chapter as well as the data filtering procedures. The results are discussed in Section 2.5. Section 2.6 summarises the main conclusions and planned further work. There are three separate appendices. Appendices A and B contain details of the model and a related statistical test. Appendix C contains detailed tables of results.

2.2 Literature Review

The effect of short selling on stock prices is debated by regulators, market participants and researchers alike. Short selling has been blamed by regulators and the media for triggering or exaggerating market declines especially during a financial crisis. However, theoretical research, including that of [Miller \(1977\)](#) and [Diamond & Verrecchia \(1987\)](#), questions the effectiveness of the regulations of short-sales constraints and bans and argue that short selling has a positive impact on stock market prices. This section presents a short review of the impact of short-sale constraints on stock price efficiency and the empirical evidence.

In an efficient price discovery process, stock prices fully reflect all current and past information and adjust to new information instantaneously ([Fama, 1991](#)). However, short sale constraints hinder the participation of investors with pessimistic opinions in the market and cause an asymmetric price transmission process. [Miller \(1977\)](#) therefore argues that stock prices have a tendency to reflect a more optimistic view which can lead to overvaluation. [Diamond & Verrecchia \(1987\)](#) suggest that short sales constraints lead to slower speed of price adjustment to new information, especially to bad news.

Direct evidence on the asymmetric speed of price adjustment is scarcer owing to the limitation of data availability and the difficulty of measuring the speed. By using differences in short selling regulations globally as different short selling constraints, [Bris *et al.* \(2007\)](#) use cross-sectional and time-series information in 46 stock markets around the world with two measures for price efficiency. The first measure, following [Morck *et al.* \(2000\)](#), is to calculate the difference in the R^2 coefficients (a downside-minus-upside R^2) measured as the relative co-movement of an individual stock with the market, depending on the sign of the market returns on weekly basis. The second measure is based on the cross-autocorrelations between individual stock returns and the signed market returns with 1-week's lag. [Bris *et al.* \(2007\)](#) find that the downside R^2 decreases more in countries where short sales are practised, which causes a drop in the difference in the downside-minus-upside R^2 . They also find short sales decrease the total cross-autocorrelation and make them significantly lower for dual-listed stocks. This accords with the hypothesis that there is a faster price discovery for dual-listed stocks which may be shortable in another market ([Foerster & Karolyi, 1999](#)) or when there is no market friction ([Hou & Moskowitz, 2005](#)). Following [Bris *et al.* \(2007\)](#)'s approaches, [Beber & Pagano \(2013\)](#) explore short sales regulations in 30 countries during the 2007-09 financial crisis. They also find that the speed of price adjustment is slowed down by short-sales bans, especially when the market is bearish in the 30 countries covered in their study.

Different from [Bris *et al.* \(2007\)](#), [Saffi & Sigurdsson \(2010\)](#) use lending supply and loan fees as more precise proxies for short sales restrictions to explain cross-sectional differences in price efficiency from 26 countries. These two short-sales proxies overcome the problems caused by country-level information and enable them to examine the role of short selling even in countries without short-sales constraints. In addition to the cross-autocorrelation applied in [Bris *et al.* \(2007\)](#), [Saffi & Sigurdsson \(2010\)](#) include the variance ratios of stocks returns ([Lo & MacKinlay, 1988](#)) and two measures of price delay ([Hou & Moskowitz, 2005](#)) which are based on the delay of price-response between stock returns and the lagged returns of a local market index as well as the contemporaneous world index returns. [Saffi & Sigurdsson \(2010\)](#) confirm that a high level of lending supply and small loan fees are related to an increase in the speed of price adjustment to the information

and suggest that the implication of short-sales constraints destroy price efficiency, which stands against to the decisions of market regulators.

Going beyond the previous work summarised above, [Boehmer & Wu \(2013\)](#) add two different measures on a higher frequency basis, the pricing error and the absolute value of intraday return autocorrelations, to capture the relative informational efficiency of prices. The pricing error in [Hasbrouck \(1993\)](#)'s VAR model captures the difference between the actual price and the fundamental value. The quote mid-points, considered as the market's best estimates of the equilibrium value of the stock at the times, should follow a random walk with smaller autocorrelation in quote returns if the price discovery is efficient. Following [Chordia *et al.* \(2005\)](#), the absolute value of return autocorrelations is estimated at a 30-minute interval. As a result, they conclude that NYSE-listed stocks with a greater flow of short sales have a smaller pricing error and that returns follow a random walk more closely. This indicates that price discovery process is more efficient when short sellers are more active in the market.

Based on the trade-by-trade high frequency data over a 3-year period, [Chen & Rhee \(2010\)](#) estimate [Hasbrouck \(1991\)](#)'s VAR model daily for the 3-month period before and after the addition or deletion event to test the speed of price adjustment to new firm-specific information in the Hong Kong stock market. The HKEx offers an ideal setting where short selling constraints are directly captured by their add/remove status on the D-list based on certain selection criteria. [Chen & Rhee \(2010\)](#) believe that stronger positive trade continuity is interpreted as a slower adjustment to new information ([Hasbrouck, 1991](#); [Hasbrouck & Ho, 1987](#)) while weaker negative quote reversal is considered as a delayed price adjustment due to less revision in beliefs accompanied by greater positive correlated order flow ([Madhavan *et al.*, 1997](#)). Contrary to [Boehmer & Wu \(2013\)](#), in the context of price adjustment speed, [Chen & Rhee \(2010\)](#) argue that more negative autocorrelation in quote returns reflects a faster speed of price adjustment to new information. In their results, they point out that quote reversals increase and trade autocorrelations decrease significantly and simultaneously when stocks are allowed for short selling. In addition, [Chen & Rhee \(2010\)](#) employ [Dimson \(1979\)](#)'s beta regression

to compare the speed of price adjustment to market-wide information between shortable and non-shortable stocks. Their findings provide evidence that shortable stock responds faster not only to firm-specific but also market-wide information.

2.3 The Model and Hypothesis Tests

This section has three parts. Section 2.3.1 describes the bivariate VAR model used in the chapter and defines the variables that are used. It then summarises the four versions of the model that are considered in this chapter. Section 2.3.2 describes a test that is used to check the internal consistency of the VAR model. Section 2.3.3 describes the various tests of hypothesis that are carried out.

2.3.1 The VAR Model

This chapter uses a bivariate VAR model to examine the speed of price adjustment to new information and to contrast the speed before and after the change in stock shortable status. [Hasbrouck \(1991\)](#) reports that transactions are informative and can cause a persistent and permanent impact on prices. The information in a new trade contains the revised beliefs of traders about prices. It thus provides the opportunity for trades to adjust prices accordingly. In conventional regression notation and for a given stock the model used in this chapter is

$$\begin{aligned}
 R_t &= \alpha_0 + \sum_{i=1}^5 \alpha_i R_{t-i} + \sum_{i=0}^5 \beta_i X_{t-i} + \varepsilon_{Rt}, \\
 X_t &= \gamma_0 + \sum_{i=1}^5 \gamma_i R_{t-i} + \sum_{i=1}^5 \delta_i X_{t-i} + \varepsilon_{Xt}.
 \end{aligned}
 \tag{2.1}$$

In applications reported in the literature, the VAR model used in this chapter and the following chapters is truncated at five lags as in [Hasbrouck \(1991\)](#) and [Chen & Rhee \(2010\)](#). Thus it includes the constants and that gives 23 parameters in all in the model. [Hasbrouck \(1991\)](#) determines the number of lags by examining autocorrelations and cross-autocorrelations between quote returns and trades. The results show that the coefficients of higher lags are not statistically significant and five is chosen as the optimal lag length as a practical empirical matter. Different from both Hasbrouck's and Chen and Rhee's in that the model in this chapter includes a constant term in each equation. This is because the model without constant terms can lead to internal inconsistencies. These may be removed by the inclusion of constant terms, which also leads to a statistical test for the consistency of the model. Study of the empirical results in both [Hasbrouck \(1991\)](#) and [Chung *et al.* \(2005\)](#), neither of whom include constants, shows that there are inconsistencies in their models. Details are available in Appendix A. It is not possible to check model consistency for [Chen & Rhee \(2010\)](#) as they do not report the necessary level of details. The concept of model consistency may not be applied to [Dufour & Engle \(2000\)](#) as they use a different model. The statistical test, which is also in part a specification test for the model, is described in Section 2.3.2 below.

The variable denoted R is the difference in the natural logarithm of the mid-quotes for two successive transactions.

$$R_t = \ln M_t - \ln M_{t-1}, M_t = (Q_t^b + Q_t^a)/2,$$

where $Q_t^{b,a}$ are the bid and ask quotes at time t . This so-called quote return will be referred to in the rest of this chapter as return. The variable denoted X is a trade measure. The measure used in this chapter is the signed volume of the trade. According to [Hasbrouck \(1991\)](#), trade X_t takes place after bid and ask quotes $Q_{t-1}^{b,a}$. The new bid and ask quotes $Q_t^{b,a}$ will be revised based on the occurred trade X_t and then more trades will follow. Hence, the variables $Q_t^{b,a}$ and X_t are not determined simultaneously even though $Q_t^{b,a}$ and X_t carry the same time subscript

t . The classification of Lee & Ready (1991) is used to define the trade sign. A buyer initiated trade is assigned a value of one if the trade price is higher than the prevailing mid-quote price before the trade. A seller initiated trade is assigned a value of minus one if the trade price is lower than the prevailing mid-quote price. If the trade price equals the prevailing mid-quote price but is higher (lower) than the previous trade price, it is classified as buyer-initiated (seller-initiated). If the trade price equals both the prevailing mid-quote price and the previous trade price, the trade sign is undetermined and is assigned a value of zero⁸. For convenience, the quote return variable denoted R is rescaled by multiplying 10^3 and the trade variable denoted X is rescaled by dividing by 10^4 .

In the VAR model described above at (2.1) the coefficients α_i and δ_i are autocorrelations in quote revision and signed trade volume respectively. The coefficients β_i indicate the impact on quote revision subsequent to each trade and the coefficient β_0 allows the immediate impact of contemporaneous trade on the quote revision. The coefficients γ_i capture Granger causality of lagged quote revision on trades. The structure of the non-standard VAR model⁹ assumes that the market contains not only information on all lagged returns and lagged trades but also information on contemporaneous trades available at time t . The residual term ε_R in the return equation at (2.1) represents the non-trade public information, while ε_X in the trade equation represents private information from unexpected trades. The non-standard VAR model with the current trade variable in the return equation captures the whole process of how private and public information is incorporated into asset prices gradually through trades by the two residual terms ε_R and ε_X .

This chapter considers four versions of the VAR model. First, the general assumption is that the 2-vector of residuals $\varepsilon_t^T = (\varepsilon_{R,t}, \varepsilon_{X,t})$ has a bivariate normal

⁸Alternative trade sign measure, which is used in some papers (Chen & Rhee, 2010; Chung *et al.*, 2005; Dufour & Engle, 2000) but not here, is that X is a discrete variable taking values equal to ± 1 or 0.

⁹By contrast, the standard VAR model contains no contemporaneous dependent variables but only lags of the dependent variables and the number of lags are the same.

distribution with zero mean vector and constant covariance matrix with no serial correlation. It is further assumed in this chapter that the contemporaneous correlation between ε_R and ε_X is zero. In this case the parameters of the two models may be estimated separately by OLS. Analogous to the market model, it is also assumed that there is no residual correlation between stocks. Secondly, an innovation adopted in this chapter is to take into account the fact that the time intervals between successive trades are not equal. In general, variations in the trade duration could provide information to market participants. Long durations are presumably due to bad news or no news (Diamond & Verrecchia, 1987; Easley & O'Hara, 1992) while trade durations are likely associated with the underlying intuition that informed investors always trade as quickly as possible and as much as possible once they have available information (Easley & O'Hara, 1992). Putting more (less) weight to the trades with less (more) time duration captures the speed of price adjustment to new information at an alternative horizon with explanatory power of inter-trade time. This chapter therefore adopts the standard assumption that volatility is proportional to the square root of the time interval. Variables in the model, quote revisions and signed trade volume are weighted by the reciprocal of $\sqrt{\Delta t}$, the square root of the inter-trade time durations, thus resulting in weighted least squares, denoted by WLS.

The chapter also includes results based on two *bona fide* VAR models, in which contemporaneous correlation between ε_R and ε_X is allowed. The two models are estimated on an unweighted and weighed basis as described above and are denoted VAR and WVAR respectively.

The parameters in the equations for R and X are referred collectively as Θ . Estimated parameters are denoted as $\hat{\alpha}_0$ and so on, collectively $\hat{\Theta}$. A subscript j is used to denote an individual stock. Before and after an addition or deletion event are denoted with subscripts B and A . These are used only when required. The estimated covariance matrix of $\hat{\Theta}$ is denoted by $\Sigma_{\hat{\Theta}}$. Sub-matrices of $\Sigma_{\hat{\Theta}}$ which are required for some of the tests described below are denoted $\Sigma_{\hat{\alpha}\hat{\alpha}}$ and so on.

2.3.2 A Model Consistency Check

A consistency check for the fitted model may be derived by taking expected values of the individual equations in (2.1) as described in Appendix B. After rearrangement and elimination of the expected value of X , this gives

$$\Omega = \left\{ 1 - \sum_{i=1}^5 \alpha_i - \left[\left(\sum_{i=0}^5 \beta_i \right) \left(\sum_{i=1}^5 \gamma_i \right) / \left(1 - \sum_{i=1}^5 \delta_i \right) \right] \right\} \mu_R - \alpha_0 - \gamma_0 \left(\sum_{i=0}^5 \beta_i \right) / \left(1 - \sum_{i=1}^5 \delta_i \right) = 0, \quad (2.2)$$

where μ_R is the unconditional expected value of R . Given the assumptions of the model, the OLS estimates of the parameters are also maximum likelihood (ML) estimates. The ML estimate of Ω is given by substituting the estimates of the elements of Θ in equation (2.2). The asymptotic variance of $\hat{\Omega}$ is given using standard methods, as described in Appendix B. The null hypothesis $H_0 : \Omega = 0$ may be tested by the Z statistic

$$\hat{\Omega} / \sqrt{\text{var}(\hat{\Omega})}, \quad (2.3)$$

which has a standard normal distribution under the null hypothesis. Rejection of this hypothesis would lead to concerns about the specification of the VAR model at (2.1).

2.3.3 Hypothesis Tests

The objective of research in this area is to test the parameters in the model for statistical significance. In addition to standard regression tests, it is important to determine whether the events of addition to or deletion from the short selling list result in statistically significant changes to the model parameters. For example, the null hypothesis for one such test would be

$$H_0 : \alpha_{j,After} = \alpha_{j,Before}, \quad (2.4)$$

along with a suitable alternative, which may be either one or two-sided. A subscript j is used to denote an individual stock. Similar tests are performed for the other parameters in the model at (2.1). If the null hypothesis is rejected, the inference that may be drawn is that the ability or inability to short sell a stock results in changes to the process of price discovery. In addition, it is also common practice to test for the significance of linear combinations of model parameters, namely

$$A = \sum_{i=1}^5 \alpha_i, B = \sum_{i=0}^5 \beta_i, \Gamma = \sum_{i=1}^5 \gamma_i, \Delta = \sum_{i=1}^5 \delta_i, \quad (2.5)$$

and to test differences before and after the event. Both of these types of test are complicated by the fact that the residual variance may also be changed by the event. It is therefore unwise to use standard t -tests which assume a constant variance.

In addition to tests of the parameters for an individual stock and a single event, it is also normal practice to examine the sums (or equivalently averages) of each model parameter before and after addition and deletion events. Thus, for n stocks, equation (2.4) becomes

$$H_0 : \sum_{j=1}^n \alpha_{j,After} = \sum_{j=1}^n \alpha_{j,Before}. \quad (2.6)$$

The corresponding null hypothesis for the linear combination, the sums, denoted as A in (2.5) is

$$H_0 : \sum_{j=1}^n A_{j,After} = \sum_{j=1}^n A_{j,Before}, \quad (2.7)$$

with similar hypotheses for the other combinations. Again, standard t -tests may not be used because, even under the assumptions of the model at (2.1), the estimated variances of the elements of $\hat{\Theta}$ are different for each stock.

Tests of Single Parameters

In this section and the next, to simplify the notation an arbitrary parameter in the regression model at (2.1) is denoted by α and its standard deviation by σ . Estimated parameters are denoted with the $\hat{\cdot}$ symbol. The standard test of $H_0 : \alpha = 0$ against one or two-sided alternatives is the t -test

$$t = \sqrt{n}\bar{\hat{\alpha}}/\bar{\hat{\sigma}}. \quad (2.8)$$

An exercise that is carried out in other papers in this area of research (Chen & Rhee, 2010) is to examine the parameter α using tests based on the estimated values $\hat{\alpha}_j$. Loosely, this is to investigate if the mean value of α is significantly different from zero. Using the notation

$$\bar{\hat{\alpha}} = n^{-1} \sum_{j=1}^n \hat{\alpha}_j, \quad \bar{\hat{\sigma}}^2 = (n-1)^{-1} \sum_{j=1}^n (\hat{\alpha}_j - \bar{\hat{\alpha}})^2,$$

the usual t -test is $t = \sqrt{n}\bar{\hat{\alpha}}/\bar{\hat{\sigma}}$, for which the null hypothesis is $H_0 : \alpha = 0$. The implicit assumption in this test is that the alternative hypothesis is $H_1 : \alpha \neq 0$. This implies in turn that a non-zero value of α is constant for all stock/events,

which is a strong assumption. It is, for example, equivalent to assuming that all betas are equal in the market model. It also seems to be an exercise of limited value since the tests carried out under (2.8) invariably indicate that some stock/events have an estimated α that is not significantly different from zero. In addition to these difficulties, the standard t -test above assumes that the variance of $\hat{\alpha}_j$ is a constant over all stock/events. Since σ is proportional to the residual volatility for the stock/event, it is clear that this assumption will be violated. A more formal test uses the property that the estimated variance of $\bar{\hat{\alpha}}$ is

$$\text{var}(\bar{\hat{\alpha}}) = n^{-2} \sum_{j=1}^n \hat{\sigma}_j^2.$$

Following [Scott & Smith \(1971\)](#), a test statistics for $H_0 : \alpha = 0$ is

$$Z = \bar{\hat{\alpha}} / \sqrt{\text{var}(\bar{\hat{\alpha}})}, \quad (2.9)$$

which is compared to the critical values of the standard normal distribution.

Tests of Multiple Parameters

Similar tests of the sums A , B , Γ and Δ , in equations (2.6) and (2.7), may be conducted using the methods described above. Using α to denote the vector of parameters α_i ($i = 1, \dots, 5$), the variances of the sums of parameters α_i ($\Sigma_{\hat{\alpha}\hat{\alpha}}$) or the other sums can be computed directly from the estimated covariance matrix $\Sigma_{\hat{\Theta}}$.

Tests of the Differences in Parameters

The setup for testing the difference in α before and after the event is as follows. The null hypothesis is $H_0 : \lambda = \alpha_{After} - \alpha_{Before} = 0$ against the two-sided alternative $H_1 : \lambda \neq 0$. Under the standard assumption

$$\hat{\lambda} = \hat{\alpha}_{After} - \hat{\alpha}_{Before} \sim N(\lambda, \sigma_{After}^2 + \sigma_{Before}^2),$$

suggests a conventional t -test for the difference between two means. However, the standard test is not robust to differences in variance before and after. A better test follows [Scott & Smith \(1971\)](#) as above and is

$$Z = (\hat{\alpha}_{After} - \hat{\alpha}_{Before}) / \sqrt{\hat{\sigma}_{After}^2 + \hat{\sigma}_{Before}^2}, \quad (2.10)$$

which has a standard normal distribution. Tests of the sum of the λ_i over the n stock/events may be performed using the same procedures in the preceding sections, noting that the estimated variance of $\hat{\lambda}$ is computed by $\hat{\sigma}_{After}^2 + \hat{\sigma}_{Before}^2$. For further technical details about equations (2.9) and (2.10), see the discussion at the end of Section 2.3.1.

2.4 Data Description

This section describes the data used for this chapter as well as the procedures for data cleaning. The study in this chapter covers a period of 10 years from May 2001 to May 2011. During this period, the D-list for short selling has been revised 86 times including 41 quarterly announcements and 45 single additions for IPO firms, with a total of 1,296 addition events and 810 deletion events. An addition (deletion) event is defined when an individual stocks is added to (removed from)

the D-list for short selling from the effective date. When an individual stock re-enters (re-quits) the D-list, it is regarded as another addition (deletion) event. The changes are summarised in Table 2.1 including the records of additions and deletions from the D-list. In May 2011 there were 760 stocks on the D-list.

[Insert Table 2.1 about here]

The unit trust, exchange traded funds, mutual funds, investment companies and stocks traded on the Growth Enterprise Market (GEM) are excluded from the analysis. Addition events for IPO firms are excluded due to the inadequacy of observations in the pre-event period. Stocks with options and futures written are excluded as option and future trading can be considered as an alternative to short selling but at a lower cost (Diamond & Verrecchia, 1987). The sample selection procedure above reduces the final data set to 1,127 addition events and 728 deletion events.

The transaction data and bid and ask prices are obtained from Trade Record and Bid and Ask Record from the HKEx, respectively. The Bid and Ask Record provides the best five bid and ask prices and quantities in the limit order book. Bid and Ask Record provides only changes which means that a data value is available only if it is different from the one previously provided. Prior to January 2008, the information is recorded at 30-second intervals with timestamp accuracy of one second¹⁰. Since 2008, it has been recorded for every change with timestamp accuracy of 0.001 second. The Trade Record provides transaction prices and volumes with a time stamp to the nearest second.

All the trades occurring outside the normal opening hours are excluded. This covers the period before 9:00 a.m. and after 4:00 p.m. and those during the pre-opening

¹⁰This chapter assumes that bid and ask information does not change within the 30 second intervals.

session (between 9:00 a.m. and 9:30 a.m.) as the trading mechanism in the pre-opening session is different from the regular trading session¹¹. To avoid overnight news arrival, the first quote return of each day is treated as a missing value for all lagged variables. Trades with the same time stamp and the same price are treated as one trade by adding up the volume corresponding to individual trades. For trades with the same time stamp but the different price, (i) calculate the signed volume weighted average price and the absolute value of the sum of signed volume for multiple trades, (ii) assign a new trade sign to the aggregated volume by the classification of Lee & Ready (1991). Anomalous data caused by systematic errors, such as zero bid and ask price and negative spreads are discarded. The estimation of the model for each stock for each event is based on all transactions in the 60-day period before and after the event date. All univariate regressions are estimated for each equation separately with the White heteroscedasticity correction for standard errors (White, 1980). The *bona fide* VAR model is estimated as a whole system using the maximum likelihood method with consideration of contemporaneous correlation between residuals in two equations. All model estimations are on an unweighted and weighted basis and are denoted as OLS, WLS, VAR and WVAR respectively.

Combining and processing the two datasets: Trade Record and Bid and Ask Record is a complex and time-consuming task due to the huge size of data (nearly 95GB for Trade Record and 1.5TB for Bid and Ask Record) on the high frequency basis over a 10-year period. Thus, this study uses powerful database software, such as SQL, to process the records. First, trade price and quantity as well as the current bid and ask prices are extracted from the datasets, respectively. Secondly, the transaction data and bid and ask prices are merged together and then sorted by time by using SQL. The nearest mid-quotes before and after each transaction are determined as the prevailing and the subsequent mid-quotes of a trade by using Perl. The prevailing mid-quotes are used to define the trade sign under the classification of Lee & Ready (1991). The prevailing and the subsequent mid-quotes are used to

¹¹Detailed information about the trading mechanism in the pre-opening and regular trading sessions can be found at http://www.hkex.com.hk/eng/market/sec_tradinfra/tradmech.htm.

calculate the quote returns. The resulting database is nearly 2GB in size. An example of the data format is shown in Appendix D.

2.5 Results

The results described in this section use 60-day transaction data (before and after the event) for model estimation and are based on 1,127 addition events and 728 deletion events. As the D-list for short sales is reviewed by the HKEx on the quarterly basis to add or remove stocks from the list, it means that 60-day (3 months) period is the appropriate period for testing the effect of short sales in the speed of price adjustment during each individual event. Moreover, a 60-day period gives a sufficient estimation window for comparison before and after an event with consideration of the availability of lending shares for short sales. Previous studies also suggest that the 60-day trading window is sufficient to allow reasonably precise estimation of the parameters (Easley & O'Hara, 1992; Easley *et al.*, 1997). The SS test is used throughout in place of the standard t -test and the significance level is 1%. All tests of significance in the thesis are at the 1% level. First, Section 2.5.1 provides a statistical summary for the stocks in the data set during the 10-year period. Section 2.5.2 reports validation tests under the OLS, WLS, VAR, and WVAR models using the procedure described in Section 2.3.2. Section 2.5.3 presents the results about changes in model parameters as a result of addition/deletion event. Sections 2.5.4 demonstrates the results for dynamics and p-values of model parameters for both addition/deletion events. Sections 2.5.5 reports the Z scores for tests on difference in model parameters for both events. Sections 2.5.6 reports results on model selection. From Section 2.5.3 onwards, detailed results are presented for OLS and summarised for the other three models as the findings of the other three models are qualitatively similar to those by OLS. A full set of detailed results for the WLS, VAR and WVAR models is available in Appendix C.

2.5.1 Statistics Summary

Table 2.2 presents a summary statistics for all sample stocks for the period 2001-2011.

[Insert Table 2.2 about here]

As shown in Table 2.2, the number of trades for addition stocks decreases from 10,654 to 8,650 after they become shortable, while the number of trades for deletion stocks trade stay at similar level after they are taken from the D-list. The number of trades for addition stocks in both pre- and post- event window are much higher than those for deletion stocks. It suggests that addition stocks are traded more actively during the 60-day period. Stocks for the addition events are traded generally at a higher average price than those for the deletion events in both the before and after periods (3.77 vs. 2.26 and 3.85 vs. 2.29, respectively). Quote returns for both addition and deletion stocks stay similar to the average level. The average trade durations for addition stocks in the before and after event are 369 and 409 seconds, while for deletion stocks are 705 and 762 seconds respectively. This is consistent with the fact that the trading for addition stocks is more frequent. The average aggregated 60-day trading volume for addition stocks is more than that for deletion stocks in the before and after periods (1,088,996 vs. 410,697 and 631,793 vs. 478,416, respectively). It shows that stocks for the addition events have higher level of interest and liquidity in the market than those for the deletion events. The signed volume indicates the average aggregated 60-day trading volume with trade sign. The signed volume decreases for addition stocks and increases for deletion stocks. This implies that more trades initiated by a seller are triggered when shocks are added to the list for short sales and there are fewer trades from sellers when stocks are removed from the list.

2.5.2 Model Validation

The bivariate VAR model used in this chapter is modified by adding a constant in each equation. A consistency check of the parameter in the modified model is conducted using the procedure described in Section 2.3.2. The Z statistic defined at (2.3) in Section 2.3.2 has a standard normal distribution under the null hypothesis. Rejection of the null hypothesis could indicate inconsistency of the model specification at (2.1). Table 2.3 contains the results for the OLS, WLS, VAR and WVAR models. The results show that for both addition and deletion events, the models are validated according to the test for almost all stocks. Compared with the results for OLS and WLS, all stocks are estimated using VAR and WVAR with a validated model. The overall Z scores, shown in the last column of Table 2.3 are computed for all sample stocks. As the table shows, the Z scores for all four methods are very small thus supporting the view that the model at (2.1) is well specified for this data set.

[Insert Table 2.3 about here]

2.5.3 Changes in Model Parameters

To examine whether short selling improves price efficiency, the estimated parameters from the models are compared before and after the addition/deletion events. For the OLS model, Table 2.4 shows the number of parameter estimates for the addition events which are significantly less than zero, significantly greater than zero or which fail to achieve statistical significance. The table has four vertical sections and two horizontal panels. The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the addition event. The second vertical section shows the corresponding results after the addition event. The section headed “*Chsq*” shows the value of a

Chi-squared statistic that is used to test the difference between the observed frequencies in the before and the after period for each variable.

The results in the Chi-squared score column, show that none of the observed sets of frequencies for each parameter exhibit significant differences. The last section presents number of differences in parameters during the addition events. It shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the addition events. Differences are calculated by absolute value of model parameters between after and before the addition event. Cells in which there is a substantial number of statistically significant changes are shown in bold. Using 112 (or 10% of the total number of 1,127 addition events) as a threshold shows that there is a substantial number of changes in α_1 (the first lagged quote autocorrelation), β_0 (the price impact of the contemporaneous trade), β_1 and the sum B . Overall, the aggregated price impact of trades is affected by short selling for nearly half of addition stocks, among which there is generally a stronger price impact of trades. However, there is only a small number of addition stocks experiencing significant change in quote autocorrelations and trade continuity during the events.

[Insert Tables 2.4 and 2.5 about here]

The corresponding results for the deletion events are reported in Table 2.5. The results show that the majority of individual parameters do not change significantly. Using 72 stocks (10% of the total number of 728 deletion events) as a threshold indicates that there is a significant number of changes for α_1 , the contemporaneous parameter β_0 and the sum B . The results show that for these stocks, there is generally a greater quote autocorrelation and a weaker price impact of trades when they are deleted from the short selling list.

Similar results based on the WLS model are omitted here, but are available in Appendix C. Briefly, WLS reduces the number of estimated parameters which are significantly different from zero and, as a consequence, the number of significant

differences during both addition and deletion events. Qualitatively, the findings are similar to those reported for the OLS model.

The corresponding results for the VAR model are shown in Tables 2.6 and 2.7. In Table 2.6, for the addition events, there are significant changes in the parameters α_1 , β_0 , β_1 and the sum B , as is the case for the OLS model. In addition, there is a significant number of changes in α_2 (quote revision autocorrelations), β_2 (price impact of trades), δ_1 (trade continuity), the sum A and the sum Δ . The results indicate that the effect of short sales during the addition events spreads to more parameters when the model at (2.1) is estimated simultaneously. Table 2.7 shows the corresponding results for the deletion events under the VAR model. There are more significant changes under the VAR model than the case for the deletion events under OLS, but fewer than for the addition events. The corresponding results for the WVAR model are in Appendix C.

[Insert Tables 2.6 and 2.7 about here]

So far, the results for changes due to both addition and deletion events show that only some model parameters are affected and that the changes in quote autocorrelation and trade continuity apply to a minority of stocks. For the stocks that are significantly affected, the parameters that change as a result of an addition or deletion event are those that measure quote revision autocorrelations, price impact of trades and trade autocorrelations. However, the overall effect of short selling is not clear as significant changes can be positive or negative. Table 2.8 contains a summary of the differences between the four models for the addition events. The table is a summary of more detailed results, which are available in Appendix C. The table has four panels. In panel (i), the results from the WLS model are compared with those from OLS. The titles of panels (ii) to (iv) indicate the other model comparisons. The table has three vertical sections: parameters before the addition event, after it and the differences. The rows of the table corresponding to those parameters for which the number of significant differences (either significant increase or decrease in the magnitude of parameters) is at least 112, that is about

10% of the 1,127 addition events. The interpretation of the table entries is best illustrated by an example.

In Table 2.4, the OLS model for the addition events, there are 1,055 stock/events for which the estimates of the parameter δ_1 do not change significantly. For the corresponding entry for the VAR model in Table 2.6, there are 929 such events. Thus, in panel (iii) of Table 2.8 the entry in the corresponding cell is $-126 = 929 - 1,055$; that is, 126 less stock/events exhibit no change in the estimated value of δ_1 under VAR than under OLS. There is a corresponding increase in the number of negative and positive changes. Panel (iii) shows that, where the number of changes is greater than (or very close to) the threshold, the corresponding cells for other parameters are also negative. This then suggests that that the OLS model is more conservative than VAR; the number of estimated parameters and changes to parameters which are significantly different from zero is increased.

[Insert Table 2.8 about here]

In panels (iii) and (iv) of Table 2.8, the entries in the corresponding cells are negative, thus suggesting that both the VAR and WVAR models are more sensitive than their least squares counterparts; the number of estimated parameters and changes to parameters which are significantly different from zero increases compared to the results for OLS and WLS. The results in panel (i) suggest that the WLS model is more conservative than OLS; the results in panel (ii) are mixed.

[Insert Table 2.9 about here]

Table 2.9 shows the corresponding results for the deletion events. Table 2.9 has the same rows as Table 2.8 to facilitate comparison, but the threshold is 72 events. The third vertical section of the table shows that the number of changes in the estimated parameters is lower for the deletion events. Where there are a significant number of changes, however, the two VAR models are, as above, more sensitive

than their least squares counterparts. WLS is somewhat more conservative than OLS and the results are once again mixed for panel (ii).

The difference comparison results summarised above for both events show that less stocks exhibit significant changes especially in the price impact of trades when inter-trade time durations are used while in general there are more stocks with significant changes in the autocorrelations in quote returns and trades if the two model equations are estimated simultaneously.

2.5.4 Parameter Dynamics

A study of parameter dynamics is carried out to check the consistency in the sign of model parameters before and after addition and deletion events. Table 2.10 shows parameter dynamics for both addition and deletion events for the OLS model. Tables for individual model parameters are available in Appendix C. The table has four panels which show the results summed over all model parameters A , B , Γ and Δ . The horizontal and vertical directions of each significance category denote before and after the event, respectively. The table is divided vertically into two sections for addition and deletion events. The predominant cells are shown in bold. In the first row of the addition section, 83.85% of stock/events have an estimated value of parameter A (quote revision autocorrelations) which is less than zero at the 1% level of significance both before and after the addition event. In the deletion section, the last row of the table 3.57% of sample stocks have an estimated value of parameter Δ (trade continuity) which is positive and significant at the 1% level before the deletion event and is positive and significant at the 5% level after the event. These results show that the majority of the sums of estimated model parameters, with the exception of the sum Γ , have consistent signs both before and after addition and deletion events. Thus, even for stocks that exhibit significant changes in individual estimated parameter values, the general performance of returns according to the OLS model remains the same in most cases.

[Insert Table 2.10 about here]

The corresponding results for WLS, VAR and WVAR are omitted, but a summary is provided in Table 2.11 with details in Appendix C. As noted in Table 2.10, the dominant cells are all in the diagonals of each sub-table, indicating consistent sign before and after the event. Table 2.11 reports the corresponding percentages for all four models. The summary broadly confirms the results in this section. It is interesting to note that there is less parameter consistency when weighted models are used.

[Insert Table 2.11 about here]

2.5.5 Z Scores for Tests of Difference in Model Parameters

To examine the effect of short selling, Table 2.12 summarises the Z scores for tests of the differences in estimated model parameters before and after addition and deletion events for the OLS, WLS, VAR and WVAR models. The Z scores for the differences are computed using the SS test statistic at (2.10). For both addition and deletion sections, there are four sub-sections corresponding to the test results for the four models. In this table, Z scores shown in bold are significantly different from zero at the 1% level.

[Insert Table 2.12 about here]

For the addition events, the parameter β_0 which captures the contemporaneous price impact of trades increases in value at the 1% significance level when the model is estimated by WLS. It therefore indicates that, when time duration is considered, on average the quote midpoint rises (falls) more immediately subsequent

to a purchase (sell) order if stocks are shortable. For the lagged variables β_1, \dots, β_5 , few significant changes are observed under any of the four models. The change in the estimated aggregated price impact parameter B is also significantly greater than zero under the WLS and WVAR models. It implies that short sales strengthen the price impact of trades with main contribution from current trades under the consideration of trade duration. For quote revision, it is found that the sum A which measures the aggregated autocorrelation in quote revision significantly decreases during the addition events under WVAR. The sum B measuring the price impact of trades increases after stocks being added to the D-list by using WLS and WVAR. There are statistically significant changes in the aggregate values of Γ (Granger causality) for four models however the changes are not consistent as those are positive for OLS and VAR and negative if the estimations are weighted. The aggregated trade continuity, the sum Δ , decreases and the significance level is at the 1% significance for all models. The overall results for the addition events show that the trades become less correlated and this pattern keeps the same under four estimation models. The quote autocorrelation also decreases when it is estimated by WVAR. The price impact of trades has been enhanced by short sales if the time duration is considered. The effect of short sales on Granger causality is not conclusive.

For the deletion events, less parameters experience significant changes. For quote revision, it is found that the estimated values of the sum A including several corresponding lagged parameters α increase at the 1% significance level by OLS. There are no significant changes for lagged parameters α for other three models. Overall, the aggregated quote autocorrelations become stronger if stocks are removed from the short-selling list in three out of four models. The aggregated trade continuity Δ is not observed to have a significant change in value under four models, thus indicating that trade continuity is not affected significantly during the deletion events. Similar as the addition events, there is a significant decrease in the aggregated Granger causality when there is consideration of inter-trade times. By contrast, there are significant increases when OLS and VAR are used. Therefore, the effect of short sales on Granger causality remains unclear. Different from the addition events, the changes in price impact of trades stay neutral which implies that short

sales do not affect price impact of trades during the deletion events.

In sum, it is found that the decrease in trade continuity is the most consistent change during the addition events while the increase in quote autocorrelations stands out from the crowd during the deletion events. When stocks become shortable, trades continuity experiences less autocorrelations. When stocks lose their eligibility of short sales, quote autocorrelations become stronger. The price impact of trades gets enhanced by short sales during the addition events when time duration is counted. Lower autocorrelations in both trade continuity and quote revision for addition stocks and greater autocorrelations in quote revision for deletion stocks suggest that the speed of price adjustment is accelerated by short sales.

2.5.6 Model Selection

This section provides some indications of the best model based on the correlation test of residuals and the summary of model estimations.

Formal model selection between OLS and VAR models can be made using a likelihood ratio test which is asymptotically equivalent to a test of the correlation between the two time series of residuals from the VAR model. Table 2.13 reports the residual correlation tests under four models. The first vertical section presents the number of stock/events with no less than 100 transactions before and after each addition and deletion event. Each cell in the second vertical section shows the percentage of stock/events for which the residual correlation is significant at 1%. The results show that no residual correlations is observed under OLS and WLS while there is a significant residual correlation for the majority of stocks under the VAR and WVAR models. It indicates that the ability to capture residual correlations between quote and trade equations is hugely better under the VAR and WVAR models.

[Insert Table 2.13 about here]

Table 2.14 provides a summary of model estimations for addition and deletion events. The significant changes are counted only if the autocorrelations of parameters measuring the speed of price adjustment decrease (increase) during the addition (deletion) events. The table shows the percentage of changes at the specified level of significance under each model. For instance, for the addition events, under the VAR model, 22.63% of stocks/events experience a faster speed of price adjustment at the 1% significance level. The figures in the table indicate that the VAR and WVAR models have better abilities to capture the changes during both events.

[Insert Table 2.14 about here]

A test of WLS vs. OLS or WVAR vs. VAR is non-standard and there is no direct test to the best of our knowledge for comparison between models with and without consideration of time duration. According to the evidences shown above, the WVAR model is considered as the preferred model as it has a better ability to capture the effect of short sales and the weighted method is more theoretically sound.

2.6 Conclusions

The institutional environment in Hong Kong offers a unique setting to investigate the effect of short sales constraints. By using the intraday data for a 10-year period, this chapter investigates the speed of price adjustment for stocks before and after the change of their eligibility of short selling. This work uses a bivariate VAR model based on four different estimation methods, namely OLS, WLS, VAR and WVAR. From a technical perspective, examining the averages of individual parameters requires non-standard approaches. This is because, even under the assumptions of the model, the estimated values of a given parameter do not have constant variance. Therefore, in addition to some standard tests commonly carried

out for regression analysis, this chapter also applies some non-standard statistical tests to facilitate a more robust investigation of the estimation results.

The main conclusions of this chapter are as follows. There is a subset of stocks for which certain parameters in the model do exhibit significant changes which implies that eligibility for short selling or loss thereof does aid price discovery, as shown by tests carried out for individual stocks before and after addition/deletion events. The parameters which measure trade continuity (quote autocorrelation) are the most affected during the addition (deletion) events. The decreased (increased) autocorrelations during the addition (deletion) events indicates that more (less) quote midpoints are likely to follow a random walk. The results are consistent with those from [Boehmer & Wu \(2013\)](#). It is also observed that trade continuity becomes stronger when stocks are removed from the D-list. Both significant changes in quote revision and trade continuity indicates greater price discovery by short sales. The study of parameter dynamics indicates that model parameters largely remain consistent. For example a negative parameter that is statistically significant will generally remain in the same category, even if the estimated value changes as a result of the addition/deletion event.

For model selection, it is found that VAR and WVAR models are better in capturing the residual correlations and the effect of short sales rather than OLS and WLS during both events. The WVAR model is considered as the preferred model to investigate price efficiency with short sales as the time-weighted model is more theoretically sound.

High frequency returns invariably exhibit heterogeneity of variance. Even though the return variables in this chapter have a non-standard definition, it is conjectured that heterogeneity of variance may have an effect on other model parameters and, if so, on the subsequent inferences. The next step in research in this area is to employ GARCH models.

**Table 2.1: The Records of Addition and Deletion Events
of the D-list on the HKEx**

The table reports the revision history of short selling list on the HKEx during the study period of 2001-2011. Columns are as follows: 1. Announcement date; 2. Effective date; 3. No. of additions; 4. No. of deletions; 5. No. of stocks on the list after each revision.

1	2	3	4	5	1	2	3	4	5
03/05/2001	14/05/2001	6	0	280	14/05/2007	21/05/2007	29	14	451
07/08/2001	20/08/2001	9	11	278	18/05/2007	21/05/2007	1	0	452
27/11/2001	03/12/2001	17	85	210	22/05/2007	29/05/2007	1	0	453
05/02/2002	25/02/2002	7	14	203	27/06/2007	04/07/2007	1	0	454
08/05/2002	21/05/2002	11	6	208	16/07/2007	17/07/2007	1	0	455
17/07/2002	29/07/2002	24	5	227	06/08/2007	13/08/2007	134	9	580
28/11/2002	29/11/2002	6	15	218	20/08/2007	27/08/2007	1	0	581
21/01/2003	27/01/2003	5	7	216	19/11/2007	26/11/2007	64	23	622
07/05/2003	19/05/2003	18	7	227	07/12/2007	14/12/2007	2	0	624
14/07/2003	21/07/2003	1	16	212	11/12/2007	14/12/2007	1	0	625
27/10/2003	03/11/2003	36	5	243	11/02/2008	18/02/2008	33	41	617
05/01/2004	06/01/2004	1	0	244	11/03/2008	13/03/2008	1	0	618
03/02/2004	10/02/2004	29	3	270	05/05/2008	13/05/2008	22	47	593
06/04/2004	07/04/2004	1	0	271	14/05/2008	15/05/2008	1	0	594
20/04/2004	27/04/2004	26	4	293	02/06/2008	03/06/2008	5	0	599
25/06/2004	01/07/2004	1	0	294	31/07/2008	07/08/2008	10	51	558
05/07/2004	09/07/2004	1	0	295	07/11/2008	14/11/2008	6	144	420
26/07/2004	02/08/2004	8	21	282	05/02/2009	12/02/2009	25	27	418
01/11/2004	08/11/2004	9	11	280	07/05/2009	14/05/2009	13	22	409
28/01/2005	07/02/2005	15	7	288	03/07/2009	10/07/2009	1	0	410
28/02/2005	01/03/2005	2	0	290	29/07/2009	05/08/2009	49	16	443
09/05/2005	17/05/2005	37	9	318	29/10/2009	05/11/2009	58	11	490

Table continued on the following page.

Table 2.1: Continued from previous page

1	2	3	4	5	1	2	3	4	5
04/07/2005	08/07/2005	1	0	319	11/11/2009	18/11/2009	1	0	491
11/07/2005	15/07/2005	1	0	320	30/11/2009	03/12/2009	1	0	492
04/08/2005	15/08/2005	14	12	322	08/12/2009	15/12/2009	1	0	493
26/08/2005	05/09/2005	1	0	323	21/12/2009	24/12/2009	1	0	494
24/10/2005	28/10/2005	1	0	324	25/01/2010	01/02/2010	65	8	551
18/11/2005	17/11/2005	11	7	328	22/02/2010	01/03/2010	1	0	552
14/02/2006	20/02/2006	10	8	330	03/03/2010	10/03/2010	1	0	553
27/02/2006	01/03/2006	2	0	332	18/03/2010	25/03/2010	1	0	554
19/05/2006	29/05/2006	23	17	338	03/05/2010	10/05/2010	59	12	601
29/05/2006	02/06/2006	1	0	339	12/07/2010	16/07/2010	1	0	602
01/06/2006	02/06/2006	1	0	340	28/07/2010	04/08/2010	40	19	623
18/08/2006	25/08/2006	38	10	368	23/08/2010	30/08/2010	1	0	624
24/08/2006	01/09/2006	1	0	369	25/10/2010	29/10/2010	47	18	653
19/10/2006	23/10/2006	1	0	370	11/11/2010	15/11/2010	1	0	654
19/10/2006	27/10/2006	1	0	371	17/11/2010	22/11/2010	2	0	656
24/11/2006	01/12/2006	55	9	417	17/12/2010	20/12/2010	1	0	657
26/02/2007	05/03/2007	30	24	423	23/12/2010	30/12/2010	1	0	658
13/03/2007	14/03/2007	1	0	424	12/01/2011	28/01/2011	1	0	659
12/04/2007	19/04/2007	5	0	429	25/01/2011	01/02/2011	1	0	660
25/04/2007	26/04/2007	6	0	435	18/02/2011	25/02/2011	70	17	713
25/04/2007	27/04/2007	1	0	436	17/05/2011	24/05/2011	65	18	760

Table 2.2: Summary Statistics

This table describes the data including trades and quotes for 1,127 addition stocks and 728 deletion stocks in a 60-day period before and after the event. Midquote is the average of bid and ask quotes after the trade. Quote Return is the log midquote change. Trade Duration is measured in seconds. Volume is the sum of the volume per transaction in a 60-day period. Signed Volume is the sum of the volume with its trade sign in a 60-day period (1 for a buy; -1 for a sale; and 0 otherwise). Each variable is calculated for each stock for each event in a 60-day period and compute the average across the stocks.

	Addition		Deletion	
No. of Stocks	1,127		728	
	Before	After	Before	After
No. of Trades	10,654	8,650	3,495	3,722
	Trade Price			
Mean	3.77	3.85	2.26	2.29
Max	4.47	4.55	2.74	2.70
Min	3.09	3.16	1.84	1.92
Std. Dev.	0.78	0.76	0.57	0.62
	Midquote			
Mean	3.77	3.85	2.26	2.29
Max	4.47	4.55	2.77	2.72
Min	3.09	3.15	1.85	1.91
Std. Dev.	0.78	0.76	0.57	0.63
	Quote Return			
Mean	0.000%	-0.001%	-0.003%	-0.004%
Max	5.099%	5.388%	10.656%	7.702%
Min	-5.164%	-5.525%	-11.294%	-8.530%
Std. Dev.	0.610%	0.661%	1.498%	1.467%
	Trade Duration			
Mean	369	409	705	762
Max	10,890	11,319	13,395	13,539
Min	1	1	1	1
Std. Dev.	1,298	1,389	1,945	2,019

Table continued on the following page.

Table 2.2: Continued from previous page

	Addition		Deletion	
	Before	After	Before	After
	Volume (1,000 Shares)			
Mean	1,088,996	631,793	410,697	478,416
Max	183,732,080	42,180,880	26,394,260	53,770,050
Min	67	55	404	256
Std. Dev.	6,192,789	1,900,829	1,403,265	2,428,428
	Signed Volume (1,000 Shares)			
Mean	46,686	23,381	6,056	27,330
Max	9,738,720	2,357,200	7,694,430	8,728,600
Min	-1,528,908	-1,624,540	-1,943,920	-784,525
Std. Dev.	377,875	166,297	322,513	458,057

Table 2.3: Model Validation

The table shows the number of occasions for which the model consistency test based on the equation (2.2) yields results at the specified level of significance under the OLS, WLS, VAR, and WVAR models. The Z scores shown in the last column are computed for all sample stocks using the Scott-Smith test.

		$\leq 0.5\%$	$\leq 1\%$	$\leq 2.5\%$	$\leq 5\%$	$\leq 10\%$	$\geq 90\%$	$\geq 95\%$	$\geq 97.5\%$	$\geq 99\%$	$\geq 99.5\%$	Z
OLS	Addition	0	0	0	0	0	0	0	0	0	0	-0.512
	After	0	0	0	0	0	1	1	1	0	0	0.136
	Deletion	0	0	0	0	0	0	0	0	0	0	0.000
	After	0	0	0	0	0	0	0	0	0	0	-0.162
WLS	Before	1	1	1	1	1	0	0	0	0	0	-0.295
	After	1	1	1	1	2	0	0	0	0	0	-0.116
	Deletion	1	1	1	1	1	0	0	0	0	0	-0.432
	After	1	1	2	3	1	1	1	1	1	0	-0.134
VAR	Before	0	0	0	0	0	0	0	0	0	0	-0.157
	After	0	0	0	0	0	0	0	0	0	0	0.055
	Deletion	0	0	0	0	0	0	0	0	0	0	0.065
	After	0	0	0	0	0	0	0	0	0	0	0.010
WVAR	Before	0	0	0	0	0	0	0	0	0	0	-0.108
	After	0	0	0	0	0	0	0	0	0	0	0.042
	Deletion	0	0	0	0	0	0	0	0	0	0	0.185
	After	0	0	0	0	0	0	0	0	0	0	0.062

**Table 2.4: Frequency of Changes in
Model Parameters for Addition Events (OLS)**

Based on the model defined at (2.1) using 60 days of trade data by OLS. The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the addition event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the addition event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the addition events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the addition events. Differences are calculated by absolute value of model parameters between after and before the addition events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 1,127) are shown in bold.

	Before Addition			After Addition			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	76	1027	24	84	1012	31	0.496	7	1110	10
α_1	1054	73	0	1044	82	1	0.456	98	907	122
α_2	919	208	0	838	289	0	0.000	61	1013	53
α_3	668	454	5	590	536	1	0.001	32	1074	21
α_4	463	661	3	384	742	1	0.001	17	1094	16
α_5	281	840	6	238	885	4	0.077	9	1112	6
β_0	1	103	1023	0	126	1001	0.170	184	673	270
β_1	1	206	920	1	275	851	0.002	87	868	172
β_2	3	419	705	0	502	625	0.000	35	1011	81
β_3	5	653	469	6	726	395	0.006	20	1051	56
β_4	3	842	282	2	866	259	0.469	10	1091	26
β_5	10	935	182	8	969	150	0.141	7	1101	19
γ_0	143	790	194	129	840	158	0.051	62	1019	46
γ_1	262	739	126	274	773	80	0.004	73	1009	45
γ_2	157	915	55	144	948	35	0.061	26	1079	22
γ_3	76	1009	42	96	1000	31	0.134	14	1102	11
γ_4	55	1033	39	40	1059	28	0.106	9	1116	2
γ_5	13	1066	48	25	1072	30	0.019	10	1114	3
δ_1	1	148	978	1	186	940	0.079	45	1055	27
δ_2	0	345	782	3	424	700	0.000	15	1097	15
δ_3	0	540	587	1	626	500	0.001	17	1097	13
δ_4	0	668	459	1	740	386	0.004	10	1114	3
δ_5	0	742	385	0	796	331	0.051	10	1113	4
A	1023	103	1	1006	121	0	0.274	89	950	88
B	3	52	1072	1	65	1061	0.286	205	589	333
Γ	186	816	125	186	861	80	0.004	57	1038	32
Δ	0	72	1055	1	87	1039	0.281	64	1019	44

**Table 2.5: Frequency of Changes in
Model Parameters for Deletion Events (OLS)**

Based on the model defined at (2.1) using 60 days of trade data by OLS. The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the deletion event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the deletion event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the deletion events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the deletion events. Differences are calculated by absolute value of model parameters between after and before the deletion events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 728) are shown in bold.

	Before Deletion			After Deletion			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	30	681	17	26	685	17	0.862	3	723	2
α_1	648	80	0	657	71	0	0.741	39	613	76
α_2	471	257	0	504	224	0	0.184	17	670	41
α_3	312	415	1	342	383	3	0.160	11	698	19
α_4	159	568	1	216	512	0	0.002	7	708	13
α_5	80	647	1	108	618	2	0.075	3	720	5
β_0	0	80	648	0	97	631	0.395	168	463	97
β_1	0	313	415	2	291	435	0.195	54	642	32
β_2	1	498	229	4	468	256	0.120	21	690	17
β_3	3	607	118	3	585	140	0.319	7	716	5
β_4	2	678	48	3	669	56	0.645	5	719	4
β_5	7	698	23	7	696	25	0.958	2	722	4
γ_0	108	555	65	68	589	71	0.006	19	682	27
γ_1	186	517	25	215	493	20	0.200	22	651	55
γ_2	72	648	8	100	622	6	0.068	6	705	17
γ_3	51	671	6	59	662	7	0.698	2	717	9
γ_4	23	700	5	28	694	6	0.738	1	726	1
γ_5	9	715	4	8	714	6	0.795	0	726	2
δ_1	0	224	504	1	207	520	0.383	16	689	23
δ_2	1	412	315	1	399	328	0.790	3	717	8
δ_3	2	540	186	1	506	221	0.108	1	716	11
δ_4	1	605	122	4	591	133	0.295	1	725	2
δ_5	1	627	100	2	632	94	0.764	3	720	5
<i>A</i>	594	134	0	607	121	0	0.669	26	644	58
<i>B</i>	0	64	664	0	111	617	0.001	158	479	91
Γ	130	579	19	141	572	15	0.619	16	681	31
Δ	0	115	613	1	129	598	0.370	18	678	32

**Table 2.6: Frequency of Changes in
Model Parameters for Addition Events (VAR)**

Based on the model defined at (2.1) using 60 days of trade data by VAR. The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the addition event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the addition event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the addition events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the addition events. Differences are calculated by absolute value of model parameters between after and before the addition events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 1,127) are shown in bold.

	Before Addition			After Addition			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	68	1044	15	76	1034	17	0.734	6	1108	13
α_1	1047	79	1	1041	86	0	0.518	184	773	170
α_2	939	188	0	896	230	1	0.044	123	899	105
α_3	741	386	0	669	448	10	0.000	80	992	55
α_4	534	584	9	450	666	11	0.002	48	1035	44
α_5	350	765	12	291	822	14	0.022	20	1083	24
β_0	0	92	1035	1	108	1018	0.298	187	655	285
β_1	2	165	960	1	223	903	0.005	141	733	253
β_2	3	414	710	0	489	638	0.001	50	949	128
β_3	5	657	465	8	700	419	0.108	35	1010	82
β_4	4	818	305	4	851	272	0.281	14	1071	42
β_5	10	922	195	19	936	172	0.114	8	1085	34
γ_0	155	784	188	151	821	155	0.130	55	1021	51
γ_1	271	725	131	253	781	93	0.010	80	996	51
γ_2	142	919	66	135	955	37	0.011	39	1063	25
γ_3	68	1016	43	79	1017	31	0.250	18	1098	11
γ_4	47	1037	43	38	1047	42	0.603	12	1110	5
γ_5	9	1068	50	14	1083	30	0.045	14	1108	5
δ_1	1	144	982	0	172	955	0.145	116	929	82
δ_2	3	212	912	3	288	836	0.001	93	974	60
δ_3	1	384	742	1	466	660	0.002	66	1014	47
δ_4	10	531	586	4	612	511	0.001	58	1026	43
δ_5	5	616	506	5	666	456	0.103	39	1048	40
A	1069	58	0	1060	66	1	0.460	174	778	175
B	1	60	1066	0	65	1062	0.547	208	597	322
Γ	180	820	127	171	877	79	0.001	70	1022	35
Δ	0	38	1089	0	52	1075	0.322	181	823	123

**Table 2.7: Frequency of Changes in
Model Parameters for Deletion Events (VAR)**

Based on the model defined at (2.1) using 60 days of trade data by VAR. The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the deletion event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the deletion event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the deletion events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the deletion events. Differences are calculated by absolute value of model parameters between after and before the deletion events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 728) are shown in bold.

	Before Deletion			After Deletion			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	23	693	12	29	689	10	0.642	1	727	0
α_1	602	125	1	598	129	1	0.963	67	551	110
α_2	462	266	0	486	242	0	0.419	47	612	69
α_3	332	392	4	342	384	2	0.638	34	662	32
α_4	184	538	6	228	499	1	0.008	16	682	30
α_5	124	598	6	128	597	3	0.587	7	710	11
β_0	0	70	658	0	89	639	0.280	177	447	104
β_1	0	270	458	3	269	456	0.222	70	607	51
β_2	5	489	234	3	463	262	0.248	35	675	18
β_3	2	613	113	3	591	134	0.303	7	712	9
β_4	8	667	53	5	674	49	0.642	6	717	5
β_5	7	703	18	9	701	18	0.881	2	724	2
γ_0	125	542	61	85	573	70	0.011	19	679	30
γ_1	191	514	23	200	508	20	0.798	30	634	64
γ_2	73	645	10	85	634	9	0.589	6	699	23
γ_3	38	680	10	45	681	2	0.052	3	718	7
γ_4	20	698	10	21	700	7	0.757	4	722	2
γ_5	8	714	6	11	710	7	0.755	4	722	2
δ_1	0	311	417	0	286	442	0.412	26	677	25
δ_2	4	305	419	1	291	436	0.291	27	672	29
δ_3	5	454	269	2	438	288	0.329	27	673	28
δ_4	3	533	192	4	524	200	0.826	17	696	15
δ_5	5	568	155	5	567	156	0.998	11	702	15
A	658	69	1	658	70	0	0.604	75	525	128
B	0	65	663	0	96	632	0.035	159	477	92
Γ	116	590	22	125	585	18	0.685	16	683	29
Δ	0	66	662	0	82	646	0.382	66	574	88

**Table 2.8: Summary of the Differences in
Frequency of Changes in Model Parameters for Addition Events**

The table reports a summary of the difference in frequency of changes in model parameters for the addition events. The significance level is defined as the 1%.

	Before Addition			After Addition			Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve	sig. dec.	no sig. change	sig. inc.
(i) OLS-WLS									
β_0	1	113	-114	1	143	-144	-78	165	-87
β_1	1	285	-286	1	280	-281	-52	145	-93
B	-1	61	-60	-1	88	-87	-76	176	-100
(ii) VAR-WVAR									
β_0	0	97	-97	0	140	-140	-83	174	-91
δ_1	-1	-41	42	4	-45	41	123	-220	97
δ_2	3	84	-87	14	44	-58	40	-115	75
B	0	48	-48	0	88	-88	-78	163	-85
Δ	0	29	-29	2	41	-43	70	-124	54
(iii) OLS-VAR									
α_1	-7	6	1	-3	4	-1	86	-134	48
α_2	20	-20	0	58	-59	1	62	-114	52
β_1	1	-41	40	0	-52	52	54	-135	81
δ_1	0	-4	4	-1	-14	15	71	-126	55
δ_2	3	-133	130	0	-136	136	78	-123	45
A	46	-45	-1	54	-55	1	85	-172	87
Δ	0	-34	34	-1	-35	36	117	-196	79
(iv) WLS-WVAR									
α_1	-66	64	2	-54	52	2	68	-120	52
β_1	1	-209	208	3	-196	193	93	-239	146
β_2	7	-183	176	5	-187	182	39	-134	95
γ_1	47	-122	75	37	-124	87	82	-128	46
δ_1	0	-200	200	4	-213	209	206	-357	151
δ_2	5	-247	242	15	-275	260	120	-243	123
δ_3	4	-248	244	9	-272	263	85	-167	82
δ_4	4	-247	243	7	-229	222	73	-154	81
δ_5	9	-223	214	11	-197	186	59	-131	72
A	83	-84	1	149	-154	5	94	-165	71
Δ	0	-65	65	2	-63	61	191	-331	140

**Table 2.9: Summary of the Differences in
Frequency of Changes in Model Parameters for Deletion Events**

The table reports a summary of the difference in frequency of changes in model parameters for the deletion events. The significance level is defined as the 1%.

	Before Deletion			After Deletion			Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve	sig. dec.	no sig. change	sig. inc.
(i) OLS-WLS									
β_0	0	124	-124	1	123	-124	-83	118	-35
B	0	120	-120	2	76	-78	-61	90	-29
(ii) VAR-WVAR									
β_0	0	114	-114	1	100	-101	-82	117	-35
δ_1	4	-141	137	2	-116	114	84	-170	86
A	-166	161	5	-166	162	4	-21	73	-52
B	0	113	-113	1	88	-89	-66	98	-32
(iii) OLS-VAR									
A	64	-65	1	51	-51	0	49	-119	70
Δ	0	-49	49	-1	-47	48	48	-104	56
(iv) WLS-WVAR									
α_1	-25	24	1	-6	5	1	36	-86	50
β_1	3	-141	138	4	-130	126	54	-98	44
δ_1	3	-152	149	1	-160	159	91	-174	83
δ_2	12	-176	164	2	-171	169	52	-106	54
A	112	-118	6	94	-98	4	41	-93	52
Δ	2	-96	94	0	-91	91	61	-155	94

Table 2.10: Parameter Dynamics (OLS)

Based on the model defined at (2.1) using 60 days of trade data by OLS. Each panel shows the results of estimates of the sums of the α s, β s, γ s and δ s respectively. Rows of each panel are before the event and columns are after it. “N” and “P” denotes negative and positive respectively. All other changes at a significance level of which is greater than 5% are classified as neutral. Bold entries in the table are percentages which are substantial. More detailed tables which cover individual parameters are available on request.

	Addition					Deletion				
	N (1%)	N (5%)	Neutral	P (5%)	P (1%)	N (1%)	N (5%)	Neutral	P (5%)	P (1%)
<i>A</i>										
N (1%)	83.85	2.84	4.08	0.00	0.00	70.88	5.49	5.22	0.00	0.00
N (5%)	3.19	0.44	1.24	0.00	0.00	5.63	1.10	1.92	0.00	0.00
Neutral	2.22	0.53	1.51	0.00	0.00	6.87	1.24	1.65	0.00	0.00
P (5%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
P (1%)	0.00	0.00	0.09	0.00	0.00	0.00	0.00	0.00	0.00	0.00
<i>B</i>										
N (1%)	0.00	0.00	0.09	0.09	0.09	0.00	0.00	0.00	0.00	0.00
N (5%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
Neutral	0.09	0.00	0.98	0.80	1.33	0.00	0.00	2.75	1.24	2.34
P (5%)	0.00	0.00	0.44	0.35	0.62	0.00	0.00	0.96	0.14	1.37
P (1%)	0.00	0.00	1.95	1.06	92.10	0.00	0.00	6.32	3.85	81.04
<i>Γ</i>										
N (1%)	7.99	1.69	6.65	0.09	0.09	8.52	0.96	8.38	0.00	0.00
N (5%)	1.33	1.15	3.64	0.18	0.09	2.47	1.51	6.32	0.14	0.27
Neutral	6.65	4.88	45.87	1.33	2.40	8.24	7.01	47.94	1.65	1.24
P (5%)	0.35	0.18	3.19	0.80	0.35	0.14	0.14	1.92	0.27	0.27
P (1%)	0.18	0.27	6.03	0.44	4.17	0.00	0.14	1.92	0.27	0.27
<i>Δ</i>										
N (1%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
N (5%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
Neutral	0.09	0.00	0.62	0.35	2.66	0.00	0.14	2.61	1.24	5.49
P (5%)	0.00	0.00	0.44	0.18	2.04	0.00	0.00	1.24	1.24	3.85
P (1%)	0.00	0.00	4.26	1.86	87.49	0.14	0.00	7.69	3.57	72.80

Table 2.11: Summary of Parameter Dynamics

The tables shows the percentages in the specified dominant cells in the OLS dynamics in Table 2.10 and the comparable percentages for the WLS, VAR and WVAR models.

		OLS	WLS	VAR	WVAR
Parameter	Cell	Addition			
<i>A</i>	N (1%)-N (1%)	83.85	59.54	91.13	71.78
<i>B</i>	P (1%)-P (1%)	92.10	81.63	91.57	81.46
Γ	Neutral-Neutral	45.87	44.54	48.09	37.53
Δ	P (1%)-P (1%)	87.49	77.99	93.17	87.58
		Deletion			
<i>A</i>	N (1%)-N (1%)	70.88	37.23	83.52	53.02
<i>B</i>	P (1%)-P (1%)	81.04	60.85	82.28	62.23
Γ	Neutral-Neutral	47.94	55.91	53.16	52.61
Δ	P (1%)-P (1%)	72.80	56.87	81.59	75.14

Table 2.12: Z Scores for Tests of Difference in Model Parameters

Based on the model defined at (2.1) using 60 days of trade data by four estimation methods. Differences are calculated by absolute value of model parameters between after and before both events (After minus Before). Z scores for difference in individual parameters and sums are computed as described in Section 2.3.3. Bold format denotes significance at the 1% level.

	Addition				Deletion			
	OLS	WLS	VAR	WVAR	OLS	WLS	VAR	WVAR
α_0	-0.113	2.202	1.014	-0.609	1.325	-0.756	-1.025	-0.297
α_1	-0.678	1.246	-0.006	0.001	4.866	0.839	0.005	0.000
α_2	-2.404	-1.220	-0.008	0.000	2.374	0.174	0.004	0.000
α_3	-1.311	-3.348	-0.020	-0.006	2.088	1.132	0.000	0.000
α_4	0.147	-0.306	-0.015	-0.003	3.810	0.205	0.005	0.000
α_5	1.249	2.060	0.004	-0.001	2.616	-0.906	0.002	0.000
β_0	2.488	3.219	1.947	2.492	1.335	2.380	0.460	1.787
β_1	-3.281	0.844	0.334	-0.547	0.660	1.882	2.395	1.715
β_2	0.695	-2.560	2.743	1.578	1.531	-0.595	0.691	0.254
β_3	0.813	1.263	1.805	1.364	0.668	0.852	0.039	-0.390
β_4	-0.192	1.750	-0.199	-0.410	1.277	1.009	0.261	0.025
β_5	-1.720	0.556	0.107	1.035	0.363	1.686	0.234	0.197
γ_0	-1.241	-0.510	-1.392	-0.413	-2.276	2.681	-1.739	2.867
γ_1	6.854	-5.951	8.273	-10.457	5.395	-3.174	5.470	-5.177
γ_2	3.212	-3.422	3.549	-4.582	4.251	-2.488	4.093	-3.316
γ_3	3.793	-3.480	4.136	-6.859	4.302	-0.934	3.719	-1.233
γ_4	2.665	-2.313	2.469	-3.550	2.584	-1.520	2.319	-1.603
γ_5	-1.021	-2.702	-1.619	-4.685	1.552	0.275	1.299	-0.134
δ_1	-0.766	-2.409	-0.138	-5.163	0.931	-0.045	0.380	-0.176
δ_2	-3.541	-2.058	-3.292	-4.937	0.367	1.275	-0.081	2.976
δ_3	-2.535	-1.030	-4.872	-3.010	1.780	-0.461	2.702	0.039
δ_4	0.605	-1.640	-0.752	-2.919	-1.267	-0.302	-1.039	0.438
δ_5	-0.600	-0.940	-1.291	-1.638	-0.536	-1.509	1.084	-1.202
A	-1.195	-0.620	-2.177	-3.211	4.923	0.617	5.394	3.347
B	1.386	3.487	2.132	2.753	1.001	2.355	1.105	1.471
Γ	7.068	-7.355	7.788	-12.257	6.188	-3.374	5.806	-4.377
Δ	-3.484	-4.489	-5.772	-9.804	0.754	-0.354	2.565	1.032

Table 2.13: Residual Correlation Test

Based on the model defined at (2.1) using 60 days of trade data by OLS, WLS, VAR and WVAR, the table reports the residual correlation test under four models. The first vertical section presents the number of stock/events with no less than 100 transactions before and after each addition and deletion event. The second section shows the percentage for which the residual correlation at the significance level of 1% under the number of observations shown in the first vertical section for each addition and deletion event.

		No. of Observations				Significant Correlation			
		OLS	WLS	VAR	WVAR	OLS	WLS	VAR	WVAR
Addition	Before	1,114	1,114	1,124	1,124	0.00%	0.00%	93.59%	92.62%
	After	1,108	1,108	1,125	1,125	0.00%	0.00%	92.71%	90.84%
Deletion	Before	716	716	726	726	0.00%	0.00%	92.84%	90.08%
	After	700	700	724	724	0.00%	0.00%	93.09%	91.71%

Table 2.14: Summary of Model Estimations

The table reports the summary of model estimation using 60 days of trade data by OLS, WLS, VAR and WVAR for addition and deletion events. There are 1,127 addition stocks and 728 deletion stocks by using these four models. The table shows the percentage of changes at the specified level of significance under each model. These significant changes are counted only if the autocorrelations of parameters measuring the speed of price adjustment decrease (increase) during the addition (deletion) events.

	OLS	WLS	VAR	WVAR
Addition				
$\leq 0.1\%$	7.19%	3.99%	16.68%	20.41%
$\leq 1.0\%$	11.98%	8.43%	22.63%	26.00%
$\leq 5.0\%$	20.85%	15.35%	30.43%	29.64%
$\leq 10.0\%$	25.73%	21.47%	33.72%	33.63%
Deletion				
$\leq 0.1\%$	4.81%	3.30%	15.25%	14.29%
$\leq 1.0\%$	11.13%	7.01%	23.35%	21.29%
$\leq 5.0\%$	19.09%	13.74%	31.18%	27.20%
$\leq 10.0\%$	22.94%	19.51%	33.52%	31.04%

Appendix A

Consistency of the VAR Model

[Hasbrouck \(1991\)](#) adopts the bivariate VAR model to examine the impact of trade-related information on prices. The model is

$$R_t = \sum_{i=1}^5 \alpha_i R_{t-i} + \sum_{i=0}^5 \beta_i X_{t-i} + \varepsilon_{Rt}, \quad (\text{A.1})$$

$$X_t = \sum_{i=1}^5 \gamma_i R_{t-i} + \sum_{i=1}^5 \delta_i X_{t-i} + \varepsilon_{Xt}. \quad (\text{A.2})$$

Using the notation $E(R_t) = \mu$ and $E(X_t) = \theta$ and assuming that $E(\varepsilon_R) = E(\varepsilon_X) = 0$ gives

$$\begin{aligned} \mu &= \sum_{i=1}^5 \alpha_i \mu + \sum_{i=0}^5 \beta_i \theta, \\ \theta &= \sum_{i=1}^5 \gamma_i \mu + \sum_{i=1}^5 \delta_i \theta. \end{aligned} \quad (\text{A.3})$$

Elimination of μ and θ gives

$$\left(1 - \sum_{i=1}^5 \alpha_i\right) \left(1 - \sum_{i=1}^5 \delta_i\right) - \sum_{i=0}^5 \beta_i \sum_{i=1}^5 \gamma_i = 0. \quad (\text{A.4})$$

Using the empirical results from [Hasbrouck \(1991\)](#),

$$\sum_{i=1}^5 \alpha_i = -0.166, \quad \sum_{i=0}^5 \beta_i = 0.023, \quad \sum_{i=1}^5 \gamma_i = -5.856, \quad \sum_{i=1}^5 \delta_i = 0.42,$$

or those from [Chung *et al.* \(2005\)](#),

$$\sum_{i=1}^5 \alpha_i = 0.0097, \quad \sum_{i=0}^5 \beta_i = 0.0915, \quad \sum_{i=1}^5 \gamma_i = -1.06, \quad \sum_{i=1}^5 \delta_i = 0.4507,$$

and substituting them into (A.4) does not lead to zero, which may therefore be considered to be a contradiction. This may be resolved by specifying a constant term in each of the models at (A.1) and (A.2), as follows

$$\begin{aligned} R_t &= \alpha_0 + \sum_{i=1}^5 \alpha_i R_{t-i} + \sum_{i=0}^5 \beta_i X_{t-i} + \varepsilon_{R,t}, \\ X_t &= \gamma_0 + \sum_{i=1}^5 \gamma_i R_{t-i} + \sum_{i=1}^5 \delta_i X_{t-i} + \varepsilon_{X,t}. \end{aligned} \quad (\text{A.5})$$

Taking expected values into (A.5) gives

$$\left(\left(1 - \sum_{i=1}^5 \alpha_i \right) \left(1 - \sum_{i=1}^5 \delta_i \right) - \sum_{i=0}^5 \beta_i \sum_{i=1}^5 \gamma_i \right) \mu = \left(1 - \sum_{i=1}^5 \delta_i \right) \alpha_0 + \sum_{i=0}^5 \beta_i \gamma_0,$$

which removes the inconsistency.

Appendix B

Test of Significance of Ω for the VAR Model Consistency

The definition of Ω is written as

$$\frac{\alpha_0 + \gamma_0 \left(\sum_{i=0}^5 \beta_i \right) / \left(1 - \sum_{i=1}^5 \delta_i \right)}{\left\{ 1 - \sum_{i=1}^5 \alpha_i - \left(\sum_{i=0}^5 \beta_i \right) \left(\sum_{i=1}^5 \gamma_i \right) / \left(1 - \sum_{i=1}^5 \delta_i \right) \right\}} = \mu_R = \mu(\Theta).$$

The variance of $\mu(\hat{\Theta})$ is $\varphi^T \Sigma_{\hat{\Theta}} \varphi$ where φ is the vector of first derivatives of $\mu(\Theta)$ with respect to the 23 elements of Θ and evaluated at $\hat{\Theta}$. Denoting the individual elements by φ_i , the vector is as

$$\begin{aligned} \varphi_1 &= \Phi_0, \varphi_k = \Phi_k, k = 2, \dots, 6; \varphi_k = \Phi_2, k = 7, \dots, 12; \\ \varphi_{13} &= \Phi_3, \varphi_k = \Phi_4, k = 14, \dots, 18; \varphi_k = \Phi_5, k = 19, \dots, 23; \end{aligned}$$

where

$$\Phi_0 = 1 - \sum_{i=1}^5 \alpha_i - \left(\sum_{i=0}^5 \beta_i \right) \left(\sum_{i=1}^5 \gamma_i \right) / \left(1 - \sum_{i=1}^5 \delta_i \right),$$

$$\Phi_1 = -\mu_R,$$

$$\Phi_2 = -\mu_R \left(\sum_{i=1}^5 \gamma_i \right) / \left(1 - \sum_{i=1}^5 \delta_i \right) - \gamma_0 / \left(1 - \sum_{i=1}^5 \delta_i \right),$$

$$\Phi_3 = -\sum_{i=0}^5 \beta_i / \left(1 - \sum_{i=1}^5 \delta_i \right),$$

$$\Phi_4 = -\mu_R \left(\sum_{i=0}^5 \beta_i \right) / \left(1 - \sum_{i=1}^5 \delta_i \right),$$

$$\Phi_5 = -\mu_R \left(\sum_{i=1}^5 \gamma_i \right) \left(\sum_{i=0}^5 \beta_i \right) / \left(1 - \sum_{i=1}^5 \delta_i \right)^2 - \gamma_0 \left(\sum_{i=0}^5 \beta_i \right) / \left(1 - \sum_{i=1}^5 \delta_i \right)^2.$$

Appendix C

Other Tables

**Table C.1: Frequency of Changes in
Model Parameters for Addition Events (WLS)**

Based on the model defined at (2.1) using 60 days of trade data by WLS. The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the addition event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the addition event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the addition events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the addition events. Differences are calculated by absolute value of model parameters between after and before the addition events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 1,127) are shown in bold.

	Before Addition			After Addition			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	22	1051	54	27	1059	41	0.314	6	1109	12
α_1	968	159	0	935	192	0	0.159	58	981	88
α_2	695	432	0	592	535	0	0.000	38	1056	33
α_3	421	704	2	309	809	9	0.000	19	1100	8
α_4	265	858	4	179	943	5	0.000	9	1115	3
α_5	137	988	2	104	1021	2	0.080	5	1117	5
β_0	2	216	909	1	269	857	0.022	106	838	183
β_1	2	491	634	2	555	570	0.026	35	1013	79
β_2	2	753	372	3	799	325	0.094	14	1072	41
β_3	5	904	218	1	930	196	0.122	6	1107	14
β_4	4	989	134	4	1023	100	0.063	3	1116	8
β_5	8	1016	103	3	1060	64	0.002	2	1115	10
γ_0	63	742	322	75	785	267	0.025	98	982	47
γ_1	148	673	306	149	728	250	0.020	70	1021	36
γ_2	69	938	120	60	992	75	0.002	26	1085	16
γ_3	25	1041	61	26	1060	41	0.128	15	1103	9
γ_4	15	1063	49	10	1092	25	0.010	10	1112	5
γ_5	7	1082	38	10	1089	28	0.356	8	1118	1
δ_1	0	303	824	0	340	787	0.226	33	1066	28
δ_2	1	543	583	2	607	518	0.021	13	1102	12
δ_3	3	697	427	1	778	348	0.001	14	1099	14
δ_4	1	828	298	1	868	258	0.148	8	1114	5
δ_5	2	851	274	1	884	242	0.229	14	1108	5
A	848	278	1	765	361	1	0.001	53	1029	45
B	2	113	1012	0	153	974	0.013	129	765	233
Γ	81	783	263	76	835	216	0.040	72	1027	28
Δ	0	132	995	0	156	971	0.318	60	1030	37

**Table C.2: Frequency of Changes in
Model Parameters for Deletion Events (WLS)**

Based on the model defined at (2.1) using 60 days of trade data by WLS. The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the deletion event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the deletion event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the deletion events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the deletion events. Differences are calculated by absolute value of model parameters between after and before the deletion events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 728) are shown in bold.

	Before Deletion			After Deletion			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	13	682	33	16	681	31	0.830	7	716	5
α_1	526	202	0	516	212	0	0.845	20	673	35
α_2	285	441	2	298	430	0	0.297	11	704	13
α_3	117	608	3	141	586	1	0.162	3	722	3
α_4	58	666	4	74	651	3	0.324	2	724	2
α_5	30	694	4	33	693	2	0.667	2	724	2
β_0	0	204	524	1	220	507	0.390	85	581	62
β_1	1	506	221	3	487	238	0.369	17	700	11
β_2	3	643	82	3	629	96	0.534	2	722	4
β_3	3	687	38	3	685	40	0.973	2	724	2
β_4	4	715	9	4	705	19	0.162	2	723	3
β_5	7	712	9	8	709	11	0.872	2	723	3
γ_0	73	547	108	45	562	121	0.023	25	671	32
γ_1	76	543	109	81	548	99	0.718	19	672	37
γ_2	20	677	31	35	667	26	0.100	4	715	9
γ_3	11	705	12	16	695	17	0.395	2	723	3
γ_4	6	711	11	10	708	10	0.590	1	725	2
γ_5	4	718	6	4	714	10	0.603	0	724	4
δ_1	1	322	405	1	330	397	0.915	19	681	28
δ_2	4	517	207	4	500	224	0.620	6	719	3
δ_3	1	613	114	1	600	127	0.657	1	717	10
δ_4	1	657	70	3	641	84	0.291	1	722	5
δ_5	0	667	61	2	652	74	0.181	1	723	4
A	380	348	0	398	330	0	0.639	13	691	24
B	0	184	544	2	187	539	0.359	97	569	62
Γ	45	602	81	45	613	70	0.637	11	685	32
Δ	0	188	540	1	208	519	0.297	29	666	33

**Table C.3: Frequency of Changes in
Model Parameters for Addition Events (WVAR)**

Based on the model defined at (2.1) using 60 days of trade data by WVAR. The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the addition event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the addition event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the addition events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the addition events. Differences are calculated by absolute value of model parameters between after and before the addition events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 1,127) are shown in bold.

	Before Addition			After Addition			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	18	1073	36	30	1053	44	0.136	5	1110	12
α_1	902	223	2	881	244	2	0.551	126	861	140
α_2	715	406	6	666	457	4	0.076	90	967	70
α_3	474	644	9	356	743	28	0.000	51	1045	31
α_4	277	825	25	220	880	27	0.015	33	1066	28
α_5	150	947	30	112	991	24	0.028	15	1092	20
β_0	0	189	938	1	248	878	0.004	104	829	194
β_1	3	282	842	5	359	763	0.001	128	774	225
β_2	9	570	548	8	612	507	0.208	53	938	136
β_3	10	751	366	12	795	320	0.104	29	1019	79
β_4	12	867	248	13	909	205	0.078	19	1069	39
β_5	15	915	197	24	950	153	0.016	16	1056	55
γ_0	62	743	322	76	796	255	0.004	96	987	44
γ_1	195	551	381	186	604	337	0.069	152	893	82
γ_2	105	834	188	87	890	150	0.020	83	999	45
γ_3	39	955	133	28	1000	99	0.020	57	1048	22
γ_4	25	989	113	21	1026	80	0.036	44	1063	20
γ_5	11	1015	101	15	1043	69	0.030	45	1074	8
δ_1	0	103	1024	4	127	996	0.032	239	709	179
δ_2	6	296	825	17	332	778	0.013	133	859	135
δ_3	7	449	671	10	506	611	0.034	99	932	96
δ_4	5	581	541	8	639	480	0.029	81	960	86
δ_5	11	628	488	12	687	428	0.037	73	977	77
A	931	194	2	914	207	6	0.276	147	864	116
B	1	108	1018	0	153	974	0.008	130	760	237
Γ	104	691	332	99	746	282	0.043	147	918	62
Δ	0	67	1060	2	93	1032	0.037	251	699	177

**Table C.4: Frequency of Changes in
Model Parameters for Deletion Events (WVAR)**

Based on the model defined at (2.1) using 60 days of trade data by WVAR. The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the deletion event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the deletion event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the deletion events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the deletion events. Differences are calculated by absolute value of model parameters between after and before the deletion events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 728) are shown in bold.

	Before Deletion			After Deletion			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	12	687	29	13	683	32	0.905	6	716	6
α_1	501	226	1	510	217	1	0.877	56	587	85
α_2	313	413	2	341	384	3	0.293	30	662	36
α_3	169	549	10	174	546	8	0.859	19	691	18
α_4	86	625	17	93	624	11	0.458	9	707	12
α_5	48	669	11	43	671	14	0.727	8	712	8
β_0	0	184	544	1	189	538	0.577	95	564	69
β_1	4	365	359	7	357	364	0.625	71	602	55
β_2	3	552	173	11	541	176	0.095	35	666	27
β_3	13	624	91	14	620	94	0.952	18	693	17
β_4	16	671	41	12	669	47	0.612	21	696	11
β_5	15	683	30	14	683	31	0.975	12	711	5
γ_0	79	551	98	48	569	111	0.013	22	680	26
γ_1	91	470	167	104	483	141	0.198	42	615	71
γ_2	31	625	72	41	626	61	0.317	17	686	25
γ_3	17	671	40	27	662	39	0.309	5	702	21
γ_4	11	689	28	18	680	30	0.403	7	708	13
γ_5	10	695	23	10	691	27	0.847	7	711	10
δ_1	4	170	554	2	170	556	0.715	110	507	111
δ_2	16	341	371	6	329	393	0.067	58	613	57
δ_3	9	482	237	6	471	251	0.569	39	651	38
δ_4	13	537	178	13	520	195	0.592	27	664	37
δ_5	10	553	165	15	559	154	0.494	24	678	26
A	492	230	6	492	232	4	0.815	54	598	76
B	0	178	550	1	184	543	0.564	93	575	60
Γ	44	576	108	57	567	104	0.403	29	649	50
Δ	2	92	634	1	117	610	0.151	90	511	127

**Table C.5: Difference in Frequency of Changes in
Model Parameters for Addition Events (OLS vs. WLS)**

The table reports the difference in frequency of changes in model parameters between OLS and WLS for the addition events. The 1% is defined as the significance level. Bold format denotes the significant difference including increase and decrease in the magnitude of parameters which is greater than or equal to 10% of the sample size, 1,127.

	Before Addition			After Addition			Difference (After minus Before)		
	-ive	ntrl	sig. dec.	sig. dec.	sig. dec.	+ve	sig. dec.	no sig. change	sig. inc.
α_0	-54	24	30	-57	47	10	-1	-1	2
α_1	-86	86	0	-109	110	-1	-40	74	-34
α_2	-224	224	0	-246	246	0	-23	43	-20
α_3	-247	250	-3	-281	273	8	-13	26	-13
α_4	-198	197	1	-205	201	4	-8	21	-13
α_5	-144	148	-4	-134	136	-2	-4	5	-1
β_0	1	113	-114	1	143	-144	-78	165	-87
β_1	1	285	-286	1	280	-281	-52	145	-93
β_2	-1	334	-333	3	297	-300	-21	61	-40
β_3	0	251	-251	-5	204	-199	-14	56	-42
β_4	1	147	-148	2	157	-159	-7	25	-18
β_5	-2	81	-79	-5	91	-86	-5	14	-9
γ_0	-80	-48	128	-54	-55	109	36	-37	1
γ_1	-114	-66	180	-125	-45	170	-3	12	-9
γ_2	-88	23	65	-84	44	40	0	6	-6
γ_3	-51	32	19	-70	60	10	1	1	-2
γ_4	-40	30	10	-30	33	-3	1	-4	3
γ_5	-6	16	-10	-15	17	-2	-2	4	-2
δ_1	-1	155	-154	-1	154	-153	-12	11	1
δ_2	1	198	-199	-1	183	-182	-2	5	-3
δ_3	3	157	-160	0	152	-152	-3	2	1
δ_4	1	160	-161	0	128	-128	-2	0	2
δ_5	2	109	-111	1	88	-89	4	-5	1
A	-175	175	0	-241	240	1	-36	79	-43
B	-1	61	-60	-1	88	-87	-76	176	-100
Γ	-105	-33	138	-110	-26	136	15	-11	-4
Δ	0	60	-60	-1	69	-68	-4	11	-7

Table C.6: Difference in Frequency of Changes in Model Parameters for Deletion Events (OLS vs. WLS)

The table reports the difference in frequency of changes in model parameters between OLS and WLS for the deletion events. The 1% is defined as the significance level. Bold format denotes the significant difference including increase and decrease in the magnitude of parameters which is greater than or equal to 10% of the sample size, 728.

	Before Deletion			After Deletion			Difference (After minus Before)		
	-ive	ntnl	sig. dec.	sig. dec.	sig. dec.	+ve	sig. dec.	no sig. change	sig. inc.
α_0	-17	1	16	-10	-4	14	4	-7	3
α_1	-122	122	0	-141	141	0	-19	60	-41
α_2	-186	184	2	-206	206	0	-6	34	-28
α_3	-195	193	2	-201	203	-2	-8	24	-16
α_4	-101	98	3	-142	139	3	-5	16	-11
α_5	-50	47	3	-75	75	0	-1	4	-3
β_0	0	124	-124	1	123	-124	-83	118	-35
β_1	1	193	-194	1	196	-197	-37	58	-21
β_2	2	145	-147	-1	161	-160	-19	32	-13
β_3	0	80	-80	0	100	-100	-5	8	-3
β_4	2	37	-39	1	36	-37	-3	4	-1
β_5	0	14	-14	1	13	-14	0	1	-1
γ_0	-35	-8	43	-23	-27	50	6	-11	5
γ_1	-110	26	84	-134	55	79	-3	21	-18
γ_2	-52	29	23	-65	45	20	-2	10	-8
γ_3	-40	34	6	-43	33	10	0	6	-6
γ_4	-17	11	6	-18	14	4	0	-1	1
γ_5	-5	3	2	-4	0	4	0	-2	2
δ_1	1	98	-99	0	123	-123	3	-8	5
δ_2	3	105	-108	3	101	-104	3	2	-5
δ_3	-1	73	-72	0	94	-94	0	1	-1
δ_4	0	52	-52	-1	50	-49	0	-3	3
δ_5	-1	40	-39	0	20	-20	-2	3	-1
A	-214	214	0	-209	209	0	-13	47	-34
B	0	120	-120	2	76	-78	-61	90	-29
Γ	-85	23	62	-96	41	55	-5	4	1
Δ	0	73	-73	0	79	-79	11	-12	1

**Table C.7: Difference in Frequency of Changes in
Model Parameters for Addition Events (VAR vs. WVAR)**

The table reports the difference in frequency of changes in model parameters between VAR and WVAR for the addition events. The 1% is defined as the significance level. Bold format denotes the significant difference including increase and decrease in the magnitude of parameters which is greater than or equal to 10% of the sample size, 1,127.

	Before Addition			After Addition			Difference (After minus Before)		
	-ive	ntrl	sig. dec.	sig. dec.	sig. dec.	+ve	sig. dec.	no sig. change	sig. inc.
α_0	-50	29	21	-46	19	27	-1	2	-1
α_1	-145	144	1	-160	158	2	-58	88	-30
α_2	-224	218	6	-230	227	3	-33	68	-35
α_3	-267	258	9	-313	295	18	-29	53	-24
α_4	-257	241	16	-230	214	16	-15	31	-16
α_5	-200	182	18	-179	169	10	-5	9	-4
β_0	0	97	-97	0	140	-140	-83	174	-91
β_1	1	117	-118	4	136	-140	-13	41	-28
β_2	6	156	-162	8	123	-131	3	-11	8
β_3	5	94	-99	4	95	-99	-6	9	-3
β_4	8	49	-57	9	58	-67	5	-2	-3
β_5	5	-7	2	5	14	-19	8	-29	21
γ_0	-93	-41	134	-75	-25	100	41	-34	-7
γ_1	-76	-174	250	-67	-177	244	72	-103	31
γ_2	-37	-85	122	-48	-65	113	44	-64	20
γ_3	-29	-61	90	-51	-17	68	39	-50	11
γ_4	-22	-48	70	-17	-21	38	32	-47	15
γ_5	2	-53	51	1	-40	39	31	-34	3
δ_1	-1	-41	42	4	-45	41	123	-220	97
δ_2	3	84	-87	14	44	-58	40	-115	75
δ_3	6	65	-71	9	40	-49	33	-82	49
δ_4	-5	50	-45	4	27	-31	23	-66	43
δ_5	6	12	-18	7	21	-28	34	-71	37
A	-138	136	2	-146	141	5	-27	86	-59
B	0	48	-48	0	88	-88	-78	163	-85
Γ	-76	-129	205	-72	-131	203	77	-104	27
Δ	0	29	-29	2	41	-43	70	-124	54

Table C.8: Difference in Frequency of Changes in Model Parameters for Deletion Events (VAR vs. WVAR)

The table reports the difference in frequency of changes in model parameters between VAR and WVAR for the deletion events. The 1% is defined as the significance level. Bold format denotes the significant difference including increase and decrease in the magnitude of parameters which is greater than or equal to 10% of the sample size, 728.

	Before Deletion			After Deletion			Difference (After minus Before)		
	-ive	ntrl	sig. dec.	sig. dec.	sig. dec.	+ve	sig. dec.	no sig. change	sig. inc.
α_0	-11	-6	17	-16	-6	22	5	-11	6
α_1	-101	101	0	-88	88	0	-11	36	-25
α_2	-149	147	2	-145	142	3	-17	50	-33
α_3	-163	157	6	-168	162	6	-15	29	-14
α_4	-98	87	11	-135	125	10	-7	25	-18
α_5	-76	71	5	-85	74	11	1	2	-3
β_0	0	114	-114	1	100	-101	-82	117	-35
β_1	4	95	-99	4	88	-92	1	-5	4
β_2	-2	63	-61	8	78	-86	0	-9	9
β_3	11	11	-22	11	29	-40	11	-19	8
β_4	8	4	-12	7	-5	-2	15	-21	6
β_5	8	-20	12	5	-18	13	10	-13	3
γ_0	-46	9	37	-37	-4	41	3	1	-4
γ_1	-100	-44	144	-96	-25	121	12	-19	7
γ_2	-42	-20	62	-44	-8	52	11	-13	2
γ_3	-21	-9	30	-18	-19	37	2	-16	14
γ_4	-9	-9	18	-3	-20	23	3	-14	11
γ_5	2	-19	17	-1	-19	20	3	-11	8
δ_1	4	-141	137	2	-116	114	84	-170	86
δ_2	12	36	-48	5	38	-43	31	-59	28
δ_3	4	28	-32	4	33	-37	12	-22	10
δ_4	10	4	-14	9	-4	-5	10	-32	22
δ_5	5	-15	10	10	-8	-2	13	-24	11
A	-166	161	5	-166	162	4	-21	73	-52
B	0	113	-113	1	88	-89	-66	98	-32
Γ	-72	-14	86	-68	-18	86	13	-34	21
Δ	2	26	-28	1	35	-36	24	-63	39

**Table C.9: Difference in Frequency of Changes in
Model Parameters for Addition Events (OLS vs. VAR)**

The table reports the difference in frequency of changes in model parameters between OLS and VAR for the addition events. The 1% is defined as the significance level. Bold format denotes the significant difference including increase and decrease in the magnitude of parameters which is greater than or equal to 10% of the sample size, 1,127.

	Before Addition			After Addition			Difference (After minus Before)		
	-ive	ntrl	sig. dec.	sig. dec.	sig. dec.	+ve	sig. dec.	no sig. change	sig. inc.
α_0	-8	17	-9	-8	22	-14	-1	-2	3
α_1	-7	6	1	-3	4	-1	86	-134	48
α_2	20	-20	0	58	-59	1	62	-114	52
α_3	73	-68	-5	79	-88	9	48	-82	34
α_4	71	-77	6	66	-76	10	31	-59	28
α_5	69	-75	6	53	-63	10	11	-29	18
β_0	-1	-11	12	1	-18	17	3	-18	15
β_1	1	-41	40	0	-52	52	54	-135	81
β_2	0	-5	5	0	-13	13	15	-62	47
β_3	0	4	-4	2	-26	24	15	-41	26
β_4	1	-24	23	2	-15	13	4	-20	16
β_5	0	-13	13	11	-33	22	1	-16	15
γ_0	12	-6	-6	22	-19	-3	-7	2	5
γ_1	9	-14	5	-21	8	13	7	-13	6
γ_2	-15	4	11	-9	7	2	13	-16	3
γ_3	-8	7	1	-17	17	0	4	-4	0
γ_4	-8	4	4	-2	-12	14	3	-6	3
γ_5	-4	2	2	-11	11	0	4	-6	2
δ_1	0	-4	4	-1	-14	15	71	-126	55
δ_2	3	-133	130	0	-136	136	78	-123	45
δ_3	1	-156	155	0	-160	160	49	-83	34
δ_4	10	-137	127	3	-128	125	48	-88	40
δ_5	5	-126	121	5	-130	125	29	-65	36
A	46	-45	-1	54	-55	1	85	-172	87
B	-2	8	-6	-1	0	1	3	8	-11
Γ	-6	4	2	-15	16	-1	13	-16	3
Δ	0	-34	34	-1	-35	36	117	-196	79

Table C.10: Difference in Frequency of Changes in Model Parameters for Deletion Events (OLS vs. VAR)

The table reports the difference in frequency of changes in model parameters between OLS and VAR for the deletion events. The 1% is defined as the significance level. Bold format denotes the significant difference including increase and decrease in the magnitude of parameters which is greater than or equal to 10% of the sample size, 728.

	Before Deletion			After Deletion			Difference (After minus Before)		
	-ive	ntrl	sig. dec.	sig. dec.	sig. dec.	+ve	sig. dec.	no sig. change	sig. inc.
α_0	-7	12	-5	3	4	-7	-2	4	-2
α_1	-46	45	1	-59	58	1	28	-62	34
α_2	-9	9	0	-18	18	0	30	-58	28
α_3	20	-23	3	0	1	-1	23	-36	13
α_4	25	-30	5	12	-13	1	9	-26	17
α_5	44	-49	5	20	-21	1	4	-10	6
β_0	0	-10	10	0	-8	8	9	-16	7
β_1	0	-43	43	1	-22	21	16	-35	19
β_2	4	-9	5	-1	-5	6	14	-15	1
β_3	-1	6	-5	0	6	-6	0	-4	4
β_4	6	-11	5	2	5	-7	1	-2	1
β_5	0	5	-5	2	5	-7	0	2	-2
γ_0	17	-13	-4	17	-16	-1	0	-3	3
γ_1	5	-3	-2	-15	15	0	8	-17	9
γ_2	1	-3	2	-15	12	3	0	-6	6
γ_3	-13	9	4	-14	19	-5	1	1	-2
γ_4	-3	-2	5	-7	6	1	3	-4	1
γ_5	-1	-1	2	3	-4	1	4	-4	0
δ_1	0	87	-87	-1	79	-78	10	-12	2
δ_2	3	-107	104	0	-108	108	24	-45	21
δ_3	3	-86	83	1	-68	67	26	-43	17
δ_4	2	-72	70	0	-67	67	16	-29	13
δ_5	4	-59	55	3	-65	62	8	-18	10
A	64	-65	1	51	-51	0	49	-119	70
B	0	1	-1	0	-15	15	1	-2	1
Γ	-14	11	3	-16	13	3	0	2	-2
Δ	0	-49	49	-1	-47	48	48	-104	56

**Table C.11: Difference in Frequency of Changes in
Model Parameters for Addition Events (WLS vs. WVAR)**

The table reports the difference in frequency of changes in model parameters between WLS and WVAR for the addition events. The 1% is defined as the significance level. Bold format denotes the significant difference including increase and decrease in the magnitude of parameters which is greater than or equal to 10% of the sample size, 1,127.

	Before Addition			After Addition			Difference (After minus Before)		
	-ive	ntrl	sig. dec.	sig. dec.	sig. dec.	+ve	sig. dec.	no sig. change	sig. inc.
α_0	-4	22	-18	3	-6	3	-1	1	0
α_1	-66	64	2	-54	52	2	68	-120	52
α_2	20	-26	6	74	-78	4	52	-89	37
α_3	53	-60	7	47	-66	19	32	-55	23
α_4	12	-33	21	41	-63	22	24	-49	25
α_5	13	-41	28	8	-30	22	10	-25	15
β_0	-2	-27	29	0	-21	21	-2	-9	11
β_1	1	-209	208	3	-196	193	93	-239	146
β_2	7	-183	176	5	-187	182	39	-134	95
β_3	5	-153	148	11	-135	124	23	-88	65
β_4	8	-122	114	9	-114	105	16	-47	31
β_5	7	-101	94	21	-110	89	14	-59	45
γ_0	-1	1	0	1	11	-12	-2	5	-3
γ_1	47	-122	75	37	-124	87	82	-128	46
γ_2	36	-104	68	27	-102	75	57	-86	29
γ_3	14	-86	72	2	-60	58	42	-55	13
γ_4	10	-74	64	11	-66	55	34	-49	15
γ_5	4	-67	63	5	-46	41	37	-44	7
δ_1	0	-200	200	4	-213	209	206	-357	151
δ_2	5	-247	242	15	-275	260	120	-243	123
δ_3	4	-248	244	9	-272	263	85	-167	82
δ_4	4	-247	243	7	-229	222	73	-154	81
δ_5	9	-223	214	11	-197	186	59	-131	72
A	83	-84	1	149	-154	5	94	-165	71
B	-1	-5	6	0	0	0	1	-5	4
Γ	23	-92	69	23	-89	66	75	-109	34
Δ	0	-65	65	2	-63	61	191	-331	140

**Table C.12: Difference in Frequency of Changes in
Model Parameters for Deletion Events (WLS vs. WVAR)**

The table reports the difference in frequency of changes in model parameters between WLS and WVAR for the deletion events. The 1% is defined as the significance level. Bold format denotes the significant difference including increase and decrease in the magnitude of parameters which is greater than or equal to 10% of the sample size, 728.

	Before Deletion			After Deletion			Difference (After minus Before)		
	-ive	ntrl	sig. dec.	sig. dec.	sig. dec.	+ve	sig. dec.	no sig. change	sig. inc.
α_0	-1	5	-4	-3	2	1	-1	0	1
α_1	-25	24	1	-6	5	1	36	-86	50
α_2	28	-28	0	43	-46	3	19	-42	23
α_3	52	-59	7	33	-40	7	16	-31	15
α_4	28	-41	13	19	-27	8	7	-17	10
α_5	18	-25	7	10	-22	12	6	-12	6
β_0	0	-20	20	0	-31	31	10	-17	7
β_1	3	-141	138	4	-130	126	54	-98	44
β_2	0	-91	91	8	-88	80	33	-56	23
β_3	10	-63	53	11	-65	54	16	-31	15
β_4	12	-44	32	8	-36	28	19	-27	8
β_5	8	-29	21	6	-26	20	10	-12	2
γ_0	6	4	-10	3	7	-10	-3	9	-6
γ_1	15	-73	58	23	-65	42	23	-57	34
γ_2	11	-52	41	6	-41	35	13	-29	16
γ_3	6	-34	28	11	-33	22	3	-21	18
γ_4	5	-22	17	8	-28	20	6	-17	11
γ_5	6	-23	17	6	-23	17	7	-13	6
δ_1	3	-152	149	1	-160	159	91	-174	83
δ_2	12	-176	164	2	-171	169	52	-106	54
δ_3	8	-131	123	5	-129	124	38	-66	28
δ_4	12	-120	108	10	-121	111	26	-58	32
δ_5	10	-114	104	13	-93	80	23	-45	22
A	112	-118	6	94	-98	4	41	-93	52
B	0	-6	6	-1	-3	4	-4	6	-2
Γ	-1	-26	27	12	-46	34	18	-36	18
Δ	2	-96	94	0	-91	91	61	-155	94

Table C.13: Parameter Dynamics (WLS)

Based on the model defined at (2.1) using 60 days of trade data by WLS. Each panel shows the results of estimates of the sums of the α s, β s, γ s and δ s respectively. Rows of each panel are before the event and columns are after it. “N” and “P” denotes negative and positive respectively. All other changes at a significance level of which is greater than 5% are classified as neutral. Bold entries in the table are percentages which are substantial. More detailed tables which cover individual parameters are available on request.

	Addition					Deletion				
	N (1%)	N (5%)	Neutral	P (5%)	P (1%)	N (1%)	N (5%)	Neutral	P (5%)	P (1%)
<i>A</i>										
N (1%)	59.54	5.94	9.76	0.00	0.00	37.23	3.57	11.26	0.14	0.00
N (5%)	3.37	1.24	3.46	0.00	0.00	6.04	2.47	6.87	0.00	0.00
Neutral	4.97	3.11	8.34	0.09	0.09	11.26	4.40	16.35	0.14	0.00
P (5%)	0.00	0.00	0.00	0.00	0.00	0.14	0.00	0.14	0.00	0.00
P (1%)	0.00	0.00	0.09	0.00	0.00	0.00	0.00	0.00	0.00	0.00
<i>B</i>										
N (1%)	0.00	0.00	0.09	0.00	0.09	0.00	0.00	0.00	0.00	0.00
N (5%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.14	0.14
Neutral	0.00	0.09	2.57	0.62	2.31	0.14	0.00	4.81	3.16	7.28
P (5%)	0.00	0.00	1.06	0.98	2.40	0.00	0.00	2.47	1.37	5.77
P (1%)	0.00	0.09	4.79	3.28	81.63	0.14	0.00	9.20	4.53	60.85
Γ										
N (1%)	3.11	0.35	3.64	0.00	0.09	1.79	0.41	3.98	0.00	0.00
N (5%)	0.35	0.71	3.82	0.09	0.00	0.69	1.24	4.12	0.00	0.00
Neutral	3.02	2.66	44.54	3.11	4.88	3.57	4.81	55.91	3.30	4.81
P (5%)	0.09	0.00	4.08	0.53	1.60	0.14	0.00	3.16	0.27	0.69
P (1%)	0.18	0.27	7.99	2.31	12.60	0.00	0.14	5.49	1.37	4.12
Δ										
N (1%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
N (5%)	0.00	0.00	0.00	0.00	0.09	0.00	0.00	0.14	0.00	0.00
Neutral	0.00	0.00	1.69	0.62	4.61	0.00	0.14	4.81	2.47	8.79
P (5%)	0.00	0.00	1.06	0.18	3.46	0.00	0.00	2.88	0.96	5.63
P (1%)	0.00	0.09	6.12	4.08	77.99	0.14	0.00	10.85	6.32	56.87

Table C.14: Parameter Dynamics (VAR)

Based on the model defined at (2.1) using 60 days of trade data by VAR. Each panel shows the results of estimates of the sums of the α s, β s, γ s and δ s respectively. Rows of each panel are before the event and columns are after it. “N” and “P” denotes negative and positive respectively. All other changes at a significance level of which is greater than 5% are classified as neutral. Bold entries in the table are percentages which are substantial. More detailed tables which cover individual parameters are available on request.

	Addition					Deletion				
	N (1%)	N (5%)	Neutral	P (5%)	P (1%)	N (1%)	N (5%)	Neutral	P (5%)	P (1%)
<i>A</i>										
N (1%)	91.13	2.22	1.42	0.00	0.09	83.52	2.88	3.98	0.00	0.00
N (5%)	1.24	0.09	0.80	0.00	0.00	3.30	0.41	0.55	0.00	0.00
Neutral	1.69	0.09	1.24	0.00	0.00	3.43	0.41	1.37	0.00	0.00
P (5%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
P (1%)	0.00	0.00	0.00	0.00	0.00	0.14	0.00	0.00	0.00	0.00
<i>B</i>										
N (1%)	0.00	0.00	0.00	0.00	0.09	0.00	0.00	0.00	0.00	0.00
N (5%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
Neutral	0.00	0.09	1.24	0.62	1.51	0.00	0.00	2.61	0.41	3.16
P (5%)	0.00	0.00	0.71	0.09	1.06	0.00	0.00	0.82	0.55	1.37
P (1%)	0.00	0.00	1.77	1.24	91.57	0.00	0.00	5.91	2.88	82.28
<i>Γ</i>										
N (1%)	8.16	1.60	5.77	0.18	0.27	7.42	2.20	6.32	0.00	0.00
N (5%)	1.06	0.62	3.99	0.09	0.18	2.61	1.10	4.81	0.14	0.27
Neutral	5.32	4.53	48.09	1.86	2.48	6.87	6.59	53.16	1.79	1.37
P (5%)	0.27	0.09	3.46	0.44	0.27	0.27	0.14	1.37	0.14	0.41
P (1%)	0.35	0.18	5.68	1.24	3.82	0.00	0.14	2.20	0.27	0.41
<i>Δ</i>										
N (1%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
N (5%)	0.00	0.00	0.09	0.00	0.00	0.00	0.00	0.00	0.00	0.00
Neutral	0.00	0.00	0.27	0.35	1.15	0.00	0.00	1.24	0.27	4.40
P (5%)	0.00	0.00	0.35	0.09	1.06	0.00	0.00	0.27	0.14	2.75
P (1%)	0.00	0.00	2.22	1.24	93.17	0.00	0.00	6.32	3.02	81.59

Table C.15: Parameter Dynamics (WVAR)

Based on the model defined at (2.1) using 60 days of trade data by WVAR. Each panel shows the results of estimates of the sums of the α s, β s, γ s and δ s respectively. Rows of each panel are before the event and columns are after it. “N” and “P” denotes negative and positive respectively. All other changes at a significance level of which is greater than 5% are classified as neutral. Bold entries in the table are percentages which are substantial. More detailed tables which cover individual parameters are available on request.

	Addition					Deletion				
	N (1%)	N (5%)	Neutral	P (5%)	P (1%)	N (1%)	N (5%)	Neutral	P (5%)	P (1%)
<i>A</i>										
N (1%)	71.78	3.64	6.74	0.09	0.35	53.02	4.40	9.75	0.00	0.41
N (5%)	3.28	0.53	1.33	0.00	0.00	4.67	0.41	2.47	0.00	0.00
Neutral	5.86	0.89	4.88	0.18	0.18	9.20	2.75	11.26	0.14	0.14
P (5%)	0.00	0.00	0.09	0.00	0.00	0.27	0.00	0.27	0.00	0.00
P (1%)	0.18	0.00	0.00	0.00	0.00	0.41	0.14	0.27	0.00	0.00
<i>B</i>										
N (1%)	0.00	0.00	0.00	0.00	0.09	0.00	0.00	0.00	0.00	0.00
N (5%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
Neutral	0.00	0.00	2.75	0.62	3.02	0.00	0.00	5.77	2.88	7.83
P (5%)	0.00	0.00	0.89	0.44	1.86	0.00	0.00	1.51	1.92	4.53
P (1%)	0.00	0.00	5.24	3.64	81.46	0.14	0.00	7.69	5.49	62.23
Γ										
N (1%)	3.90	0.89	4.26	0.09	0.09	2.20	0.69	3.02	0.14	0.00
N (5%)	0.71	0.35	3.37	0.09	0.18	0.41	0.55	3.02	0.00	0.00
Neutral	3.28	1.95	37.53	2.57	5.50	4.81	4.81	52.61	2.61	5.63
P (5%)	0.18	0.35	3.19	0.27	1.77	0.14	0.41	3.43	0.14	0.55
P (1%)	0.71	0.00	8.87	2.40	17.48	0.27	0.00	4.95	1.51	8.10
Δ										
N (1%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.14	0.00	0.14
N (5%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.14	0.00	0.27
Neutral	0.00	0.00	1.06	0.18	2.66	0.00	0.00	2.20	0.41	5.63
P (5%)	0.00	0.00	0.27	0.44	1.33	0.14	0.00	1.10	0.14	2.61
P (1%)	0.18	0.18	3.64	2.48	87.58	0.00	0.00	8.52	3.43	75.14

Table C.16: P-Values for Tests of Model Parameters (OLS)

Based on the model defined at (2.1) using 60 days of trade data by OLS. The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters and probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. Bold format denotes significance at the 1% level.

	Before		Addition		Difference		Before		Deletion		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
α_0	-0.026	0.001	0.025	0.009	-0.001	0.909	0.068	0.001	0.131	0.002	0.063	0.185
α_1	-0.204	0.000	-0.203	0.000	-0.001	0.497	-0.249	0.000	-0.266	0.000	0.017	0.000
α_2	-0.123	0.000	-0.118	0.000	-0.004	0.016	-0.143	0.000	-0.151	0.000	0.008	0.018
α_3	-0.066	0.000	-0.064	0.000	-0.002	0.190	-0.082	0.000	-0.088	0.000	0.006	0.037
α_4	-0.042	0.000	-0.043	0.000	0.000	0.883	-0.047	0.000	-0.058	0.000	0.011	0.000
α_5	-0.022	0.000	-0.024	0.000	0.002	0.211	-0.025	0.000	-0.032	0.000	0.007	0.009
β_0	0.180	0.000	0.217	0.000	0.037	0.013	0.379	0.000	0.452	0.000	0.073	0.182
β_1	0.143	0.000	0.079	0.000	-0.064	0.001	0.097	0.000	0.130	0.007	0.032	0.509
β_2	0.001	0.975	0.019	0.059	0.018	0.486	0.048	0.000	0.139	0.018	0.091	0.126
β_3	-0.005	0.613	0.015	0.107	0.011	0.416	0.030	0.001	0.051	0.085	0.021	0.504
β_4	0.025	0.176	0.021	0.008	-0.004	0.847	0.020	0.016	-0.081	0.087	0.062	0.201
β_5	-0.053	0.013	-0.009	0.545	-0.044	0.085	0.003	0.789	0.017	0.659	0.015	0.716
γ_0	0.100	0.000	0.074	0.000	-0.026	0.214	-0.256	0.000	-0.107	0.040	-0.148	0.023
γ_1	-0.064	0.000	-0.094	0.000	0.030	0.000	-0.088	0.000	-0.119	0.000	0.031	0.000
γ_2	-0.053	0.000	-0.066	0.000	0.014	0.001	-0.052	0.000	-0.076	0.000	0.025	0.000
γ_3	-0.027	0.000	-0.043	0.000	0.016	0.000	-0.027	0.000	-0.049	0.000	0.022	0.000
γ_4	-0.010	0.000	-0.021	0.000	0.011	0.008	-0.017	0.000	-0.031	0.000	0.014	0.010
γ_5	0.007	0.010	-0.003	0.140	-0.004	0.307	-0.007	0.026	-0.014	0.000	0.007	0.121
δ_1	0.132	0.000	0.131	0.000	-0.001	0.443	0.145	0.000	0.148	0.000	0.003	0.352
δ_2	0.075	0.000	0.069	0.000	-0.006	0.000	0.073	0.000	0.074	0.000	0.001	0.714
δ_3	0.053	0.000	0.048	0.000	-0.004	0.011	0.050	0.000	0.054	0.000	0.005	0.075
δ_4	0.038	0.000	0.040	0.000	0.001	0.545	0.036	0.000	0.033	0.000	-0.003	0.205
δ_5	0.034	0.000	0.033	0.000	-0.001	0.548	0.032	0.000	0.031	0.000	-0.001	0.592
A	-0.458	0.000	-0.451	0.000	-0.006	0.232	-0.546	0.000	-0.595	0.000	0.049	0.000
B	0.291	0.000	0.343	0.000	0.052	0.166	0.577	0.000	0.707	0.000	0.131	0.316
Γ	-0.146	0.000	-0.227	0.000	0.081	0.000	-0.191	0.000	-0.289	0.000	0.098	0.000
Δ	0.333	0.000	0.321	0.000	-0.012	0.000	0.336	0.000	0.339	0.000	0.004	0.451

Table C.17: P-Values for Tests of Model Parameters (WLS)

Based on the model defined at (2.1) using 60 days of trade data by WLS. The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters and probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. Bold format denotes significance at the 1% level.

	Before		Addition		Difference		Before		Deletion		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
α_0	0.005	0.000	0.015	0.001	0.010	0.028	0.022	0.000	0.015	0.068	-0.007	0.449
α_1	-0.123	0.000	-0.125	0.000	0.003	0.212	-0.147	0.000	-0.150	0.000	0.003	0.401
α_2	-0.065	0.000	-0.061	0.000	-0.003	0.222	-0.068	0.000	-0.069	0.000	0.001	0.861
α_3	-0.027	0.000	-0.019	0.000	-0.007	0.001	-0.025	0.000	-0.029	0.000	0.003	0.258
α_4	-0.017	0.000	-0.016	0.000	-0.001	0.760	-0.013	0.000	-0.014	0.000	0.001	0.837
α_5	-0.009	0.000	-0.013	0.000	0.004	0.039	-0.011	0.000	-0.009	0.000	-0.002	0.365
β_0	0.144	0.000	0.239	0.000	0.094	0.001	0.359	0.000	0.535	0.000	0.176	0.017
β_1	0.052	0.000	0.067	0.000	0.015	0.399	0.069	0.000	0.143	0.000	0.073	0.060
β_2	0.046	0.000	0.009	0.261	-0.037	0.010	0.024	0.047	0.003	0.938	-0.022	0.551
β_3	-0.011	0.190	0.028	0.012	0.017	0.206	0.028	0.026	0.067	0.129	0.039	0.394
β_4	0.015	0.176	0.051	0.003	0.036	0.080	0.005	0.694	-0.032	0.192	0.027	0.313
β_5	0.018	0.094	0.028	0.065	0.010	0.578	-0.003	0.809	0.142	0.082	0.139	0.092
γ_0	0.119	0.000	0.115	0.000	-0.004	0.609	0.027	0.103	0.083	0.000	0.056	0.007
γ_1	0.118	0.000	0.065	0.000	-0.053	0.000	0.075	0.000	0.036	0.000	-0.039	0.001
γ_2	0.047	0.000	0.023	0.000	-0.024	0.001	0.028	0.000	0.005	0.509	-0.023	0.013
γ_3	0.057	0.000	0.019	0.014	-0.038	0.001	0.017	0.000	0.009	0.202	-0.008	0.350
γ_4	0.041	0.000	0.023	0.000	-0.019	0.021	0.022	0.000	0.009	0.146	-0.012	0.129
γ_5	0.053	0.000	0.030	0.000	-0.023	0.007	0.014	0.000	0.017	0.009	0.002	0.783
δ_1	0.131	0.000	0.126	0.000	-0.005	0.016	0.143	0.000	0.142	0.000	0.000	0.963
δ_2	0.069	0.000	0.065	0.000	-0.004	0.040	0.061	0.000	0.066	0.000	0.004	0.202
δ_3	0.047	0.000	0.046	0.000	-0.002	0.303	0.043	0.000	0.042	0.000	-0.001	0.644
δ_4	0.037	0.000	0.034	0.000	-0.003	0.101	0.029	0.000	0.028	0.000	-0.001	0.762
δ_5	0.032	0.000	0.031	0.000	-0.002	0.347	0.029	0.000	0.025	0.000	-0.004	0.131
A	-0.240	0.000	-0.236	0.000	-0.004	0.535	-0.265	0.000	-0.270	0.000	0.005	0.537
B	0.264	0.000	0.422	0.000	0.158	0.000	0.482	0.000	0.858	0.000	0.376	0.018
Γ	0.316	0.000	0.160	0.000	-0.156	0.000	0.156	0.000	0.076	0.000	-0.080	0.001
Δ	0.316	0.000	0.301	0.000	-0.015	0.000	0.304	0.000	0.302	0.000	-0.002	0.723

Table C.18: P-Values for Tests of Model Parameters (VAR)

Based on the model defined at (2.1) using 60 days of trade data by VAR. The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters and probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. Bold format denotes significance at the 1% level.

	Before		Addition				Before		Deletion			
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
α_0	-0.004	0.600	-0.015	0.072	0.011	0.310	0.047	0.051	0.005	0.880	-0.042	0.305
α_1	-0.190	0.000	-0.189	0.098	-0.001	0.995	-0.227	0.707	-0.235	0.854	0.008	0.995
α_2	-0.106	0.010	-0.105	0.366	-0.001	0.994	-0.119	0.844	-0.125	0.918	0.006	0.997
α_3	-0.059	0.154	-0.057	0.628	-0.003	0.984	-0.068	0.912	-0.067	0.954	0.000	1.000
α_4	-0.036	0.371	-0.034	0.769	-0.002	0.987	-0.038	0.950	-0.046	0.969	0.008	0.995
α_5	-0.020	0.616	-0.020	0.863	0.001	0.996	-0.021	0.973	-0.025	0.982	0.004	0.998
β_0	0.205	0.000	0.233	0.000	0.029	0.051	0.444	0.000	0.469	0.000	0.025	0.645
β_1	0.079	0.000	0.083	0.000	0.004	0.738	0.071	0.000	0.200	0.000	0.129	0.017
β_2	0.024	0.004	0.059	0.000	0.035	0.006	0.044	0.021	0.080	0.097	0.036	0.489
β_3	0.004	0.658	0.027	0.007	0.023	0.071	0.011	0.562	-0.013	0.777	0.002	0.969
β_4	0.004	0.661	0.001	0.926	-0.003	0.842	0.019	0.314	-0.033	0.485	0.013	0.794
β_5	0.004	0.600	-0.005	0.591	0.001	0.914	-0.010	0.611	0.021	0.627	0.011	0.814
γ_0	0.095	0.000	0.067	0.000	-0.028	0.164	-0.279	0.000	-0.176	0.000	-0.103	0.082
γ_1	-0.060	0.000	-0.089	0.000	0.029	0.000	-0.082	0.000	-0.110	0.000	0.028	0.000
γ_2	-0.048	0.000	-0.061	0.000	0.013	0.000	-0.045	0.000	-0.066	0.000	0.022	0.000
γ_3	-0.024	0.000	-0.039	0.000	0.015	0.000	-0.023	0.000	-0.043	0.000	0.020	0.000
γ_4	-0.009	0.000	-0.017	0.000	0.009	0.014	-0.013	0.000	-0.025	0.000	0.012	0.020
γ_5	0.008	0.001	-0.002	0.404	-0.006	0.105	-0.006	0.066	-0.013	0.002	0.007	0.194
δ_1	0.134	0.000	0.134	0.000	0.000	0.890	0.144	0.000	0.148	0.000	0.003	0.704
δ_2	0.072	0.000	0.069	0.000	-0.003	0.001	0.071	0.000	0.071	0.000	0.000	0.935
δ_3	0.051	0.000	0.047	0.000	-0.005	0.000	0.046	0.000	0.051	0.000	0.005	0.007
δ_4	0.036	0.000	0.035	0.000	-0.001	0.452	0.032	0.000	0.030	0.000	-0.002	0.299
δ_5	0.032	0.000	0.031	0.000	-0.001	0.196	0.026	0.000	0.028	0.000	0.002	0.278
A	-0.411	0.000	-0.406	0.000	-0.006	0.029	-0.473	0.000	-0.498	0.000	0.025	0.000
B	0.318	0.000	0.398	0.000	0.080	0.033	0.580	0.000	0.724	0.000	0.144	0.269
Γ	-0.133	0.000	-0.208	0.000	0.075	0.000	-0.168	0.000	-0.257	0.000	0.089	0.000
Δ	0.326	0.000	0.315	0.000	-0.010	0.000	0.320	0.000	0.328	0.000	0.008	0.010

Table C.19: P-Values for Tests of Model Parameters (WVAR)

Based on the model defined at (2.1) using 60 days of trade data by WVAR. The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters and probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. Bold format denotes significance at the 1% level.

	Before		Addition		Difference		Before		Deletion		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
α_0	0.006	0.000	0.005	0.001	-0.001	0.542	0.022	0.000	0.019	0.002	-0.002	0.766
α_1	-0.114	0.572	-0.115	0.894	0.001	0.999	-0.130	0.837	-0.136	0.988	0.006	0.999
α_2	-0.056	0.789	-0.056	0.949	0.000	1.000	-0.059	0.929	-0.061	0.995	0.002	1.000
α_3	-0.028	0.893	-0.022	0.978	-0.005	0.995	-0.025	0.970	-0.025	0.998	0.000	1.000
α_4	-0.015	0.940	-0.012	0.988	-0.003	0.997	-0.011	0.986	-0.015	0.999	0.004	1.000
α_5	-0.009	0.965	-0.008	0.993	-0.001	0.999	-0.005	0.994	-0.007	0.999	0.002	1.000
β_0	0.211	0.000	0.284	0.000	0.073	0.013	0.451	0.000	0.583	0.000	0.132	0.074
β_1	0.050	0.000	0.040	0.005	-0.010	0.584	0.025	0.148	0.137	0.031	0.113	0.086
β_2	0.012	0.239	0.040	0.006	0.028	0.114	0.020	0.249	0.037	0.560	0.017	0.799
β_3	0.014	0.168	0.037	0.006	0.023	0.173	0.033	0.064	-0.007	0.918	-0.026	0.696
β_4	0.011	0.249	0.004	0.744	-0.007	0.682	-0.004	0.823	-0.006	0.930	0.002	0.979
β_5	0.002	0.814	0.020	0.148	0.017	0.300	0.004	0.826	0.017	0.792	0.013	0.844
γ_0	0.121	0.000	0.118	0.000	-0.004	0.679	0.025	0.088	0.085	0.000	0.060	0.004
γ_1	0.115	0.000	0.062	0.000	-0.053	0.000	0.069	0.000	0.028	0.000	-0.042	0.000
γ_2	0.046	0.000	0.023	0.000	-0.023	0.000	0.030	0.000	0.003	0.648	-0.027	0.001
γ_3	0.057	0.000	0.022	0.000	-0.035	0.000	0.019	0.001	0.009	0.149	-0.010	0.217
γ_4	0.042	0.000	0.024	0.000	-0.018	0.000	0.021	0.000	0.008	0.163	-0.013	0.109
γ_5	0.054	0.000	0.030	0.000	-0.024	0.000	0.016	0.002	0.014	0.013	-0.001	0.893
δ_1	0.132	0.000	0.126	0.000	-0.006	0.000	0.143	0.000	0.142	0.000	-0.001	0.860
δ_2	0.069	0.000	0.064	0.000	-0.005	0.000	0.059	0.000	0.064	0.000	0.005	0.003
δ_3	0.046	0.000	0.043	0.000	-0.003	0.003	0.041	0.000	0.041	0.000	0.000	0.969
δ_4	0.035	0.000	0.032	0.000	-0.003	0.004	0.025	0.000	0.026	0.000	0.001	0.661
δ_5	0.030	0.000	0.029	0.000	-0.002	0.101	0.026	0.000	0.024	0.000	-0.002	0.229
A	-0.222	0.000	-0.214	0.000	-0.008	0.001	-0.230	0.000	-0.245	0.000	0.014	0.001
B	0.300	0.000	0.425	0.000	0.124	0.006	0.528	0.000	0.763	0.000	0.234	0.141
Γ	0.314	0.000	0.162	0.000	-0.152	0.000	0.154	0.000	0.062	0.000	-0.092	0.000
Δ	0.312	0.000	0.294	0.000	-0.018	0.000	0.294	0.000	0.298	0.000	0.003	0.302

Appendix D

The Example of Data Format

Table D.1: The Example of Data Format

The table reports the example of data format cleaned from Trade Record and Bid and Ask Record datasets. Columns are as follows: 1. Stock ID; 2. Effective date; 3. Estimated date; 4. Traded time (HH:MM:SS); 5. Traded price; 6. Traded volume; 7. Prevailing mid-quotes; 8. Subsequent mid-quotes.

1	2	3	4	5	6	7	8
1000060	20020225	20011205	145657	0.56	500000	0.565	0.565
1000060	20020225	20011205	150607	0.56	50000	0.565	0.565
1000060	20020225	20011205	150934	0.56	252000	0.565	0.565
1000060	20020225	20011205	151009	0.56	130000	0.565	0.565
1000060	20020225	20011205	151708	0.56	40000	0.565	0.565
1000060	20020225	20011205	151922	0.56	80000	0.565	0.565
1000060	20020225	20011205	152417	0.57	100000	0.565	0.565
1000060	20020225	20011205	152713	0.56	100000	0.565	0.565
1000060	20020225	20011205	152822	0.57	20000	0.565	0.565
1000060	20020225	20011205	153248	0.56	900000	0.565	0.565
1000060	20020225	20011205	153330	0.56	50000	0.565	0.555
1000060	20020225	20011205	153331	0.56	32000	0.565	0.555
1000060	20020225	20011205	153332	0.56	50000	0.565	0.555
1000060	20020225	20011205	153851	0.56	100000	0.555	0.555
1000060	20020225	20011205	153931	0.56	160000	0.555	0.555
1000060	20020225	20011205	154033	0.56	40000	0.555	0.555
1000060	20020225	20011205	154046	0.56	30000	0.555	0.555
1000060	20020225	20011205	154324	0.56	40000	0.555	0.555

Chapter 3

Short Sales and Price Discovery using GARCH and BEKK Models

3.1 Introduction

Volatility as a measure of risk varies and has a tendency to cluster over times and this phenomenon corresponds to the fluctuations in volatility. Due to the dependence of volatility upon past realization of the asset price and the related volatility process, large swings tend to be followed by large swings while small changes tend to be not far behind small changes [Mandelbrot \(1963\)](#). Modelling dynamic volatility has attracted much attention of academics ever since the introduction of the Autoregressive Conditional Heteroscedasticity [ARCH henceforth] model by [Engle \(1982\)](#). Since then numerous variants and extensions of ARCH models have been proposed. Among these, the GARCH model developed by [Bollerslev \(1986\)](#) is the most widely used GARCH form to describe the phenomenon of volatility clustering and capture the characteristic of heteroscedasticity in time series analysis.

While volatility modelling has been the main centre of attention, understanding the dynamic co-movements of volatility is also of great practical importance since this

will identify volatility and correlation transmission and spillover effects from one to another. This issue can be solved by extending the considerations to multivariate GARCH [MGARCH henceforth] models. However, the curse of dimensionality becomes an issue for a MGARCH model as the dimension of volatility matrix increases rapidly when the number of variable increases. On the one hand, the specification of a MGARCH model should be flexible enough to be able to represent the dynamics of the conditional variances and covariances. On the other hand, as the number of parameters in a MGARCH model often increases rapidly with the dimension of the model, the specification should be parsimonious enough to allow for relatively easy estimation of the model and also allow for easy interpretation of the model parameters. However, parsimony may reduce the number of parameters, in which situation the relevant dynamics in the covariance matrix cannot be captured. Therefore, it is important to achieve a balance between parsimony and flexibility when designing the MGARCH model specifications. Another feature a MGARCH model must satisfy is that the covariance matrix should be positive definite. To overcome these difficulties, a BEKK model as proposed by [Engle & Kroner \(1995\)](#) which is the one of the most popular MGARCH models of conditional covariances and correlations is used to investigate the volatility spillover among financial markets. The BEKK model is considered as a truncated and low-dimensional application and it ensures the positive semi-definiteness of covariance matrix.

In Chapter 2, the high frequency based quote returns and signed volumes in the VAR model of [Hasbrouck \(1991\)](#) invariably exhibit heterogeneity of variance and thus they could affect the model parameters which measure the price efficiency. To further investigate the impact of short sales on the speed of price adjustment under dynamic volatility and co-volatility, GARCH-based analysis is conducted in this chapter on the basis of formal econometric diagnostic tests on high frequency data. This chapter makes the following contributions to the literature. To our best knowledge, this is the first study which investigates the relationship between short sales and speed of price adjustment based on the VAR model with univariate GARCH and BEKK extensions. The univariate GARCH models help capture the volatility dynamics of individual quote returns or signed volumes. The BEKK model takes account of possible volatility spillover between quote returns

and signed volumes. This chapter also carries out different diagnostic checking on the ARCH effect including univariate tests applied independently to each series and multivariate tests applied to the vector series as a whole. By comparing the GARCH and BEKK models with the least squares based models in Chapter 2, this chapter further concludes the preferred models to capture the effect of short sales on price efficiency are the WVAR and BEKK models.

The results are summarised as follows. First, under GARCH and BEKK frameworks, there are more stocks experiencing significant changes in parameters, including increases and decreases in the magnitude, than those under the models in Chapter 2. It suggests that more stocks' price efficiency is affected by short sales when heterogeneity of variance is taken into account for both addition and deletion events. Autocorrelations in quote revision and trades are weakened by short sales during the addition events. Quote autocorrelations become stronger during the deletion events. These findings suggest that short sales improve the price process that quote revision and trades follow a random walk with smaller absolute value of autocorrelations which are consistent with those from [Boehmer & Wu \(2013\)](#). Thus, it is concluded that speed of price adjustment becomes faster by short sales when the models are estimated under heteroscedasticity. Secondly, trade autocorrelations are the most affected parameters which decrease significantly during the addition events and this is robust under all models including OLS, WLS, VAR, WVAR, GARCH and BEKK. For the deletion event, it is observed that quote autocorrelations are the most affected parameters with significant increases under the majority of models while no evidence is shown for significant changes in trade continuity. Furthermore, it is noted that WVAR and BEKK models that estimate simultaneously with consideration of heterogeneous variances are more powerful to capture the effect of short sales than other models especially for the addition events. Thirdly, a study of parameter dynamics suggests that parameters under GARCH and BEKK models remain consistent as far as sign is concerned; that is, an estimated model parameter which is significantly negative (positive) before the addition/deletion event will remain negative (positive) after the event. It is also found that, under the GARCH model, there are more stocks with consistency in the sign of parameters especially for those measuring the speed of price adjustment

compared with OLS model.

The remainder of this chapter is organised as follows. Section 3.2 presents GARCH and MGARCH models with model diagnostics. Section 3.3 provides data description. Section 3.4 discusses the empirical results. Section 3.5 concludes the chapter.

3.2 The Models and Diagnostic Checking

This section summarises the univariate GARCH and BEKK models that are used in this chapter in conjunction with the VAR model of [Hasbrouck \(1991\)](#). There are three parts in the section. Section 3.2.1 describes the univariate GARCH model. Section 3.2.2 describes the BEKK model. Section 3.2.3 summarises the diagnostic tests that are employed.

3.2.1 The Univariate GARCH Model

A limitation of the ARCH model of [Engle \(1982\)](#) is that it often requires many lags. [Bollerslev \(1986\)](#) proposed the GARCH model, in which the current conditional variance is represented in terms of past squared residuals and past conditional variances. It is thus in effect an ARMA model in which the past conditional variances are the auto-regressive terms and the past squared residuals the moving averages. In common with standard ARMA models this often results in a more parsimonious parametrisation. In general a GARCH (p, q) model has the following form

$$\sigma_t^2 = \varphi_0 + \varphi_1 \varepsilon_{t-1}^2 + \dots + \varphi_p \varepsilon_{t-p}^2 + \phi_1 \sigma_{t-1}^2 + \dots + \phi_q \sigma_{t-q}^2,$$

where $\varphi_i > 0$, $\phi_j > 0$, $i = 1, \dots, p$, and $j = 1, \dots, q$, respectively and

$$1 - \sum_{i=1}^p \varphi_i - \sum_{j=1}^q \phi_j > 0.$$

The most frequently used model in finance is the GARCH (1, 1) model, which is adopted in this chapter. In this model, the conditional variance matrix is calculated from a long-run average variance rate $\omega > 0$ and from the lag terms ε_{t-1} and σ_{t-1} . The univariate GARCH (1, 1) model can be written as

$$\sigma_t^2 = \omega + \varphi \varepsilon_{t-1}^2 + \phi \sigma_{t-1}^2,$$

where $\omega > 0$, $\varphi \geq 0$, $\phi \geq 0$ and $1 - \varphi - \phi > 0$. Applying the GARCH (1, 1) model to the VAR model of [Hasbrouck \(1991\)](#), the univariate form of the model is defined as

$$\begin{aligned} R_t &= \alpha_0 + \sum_{i=1}^5 \alpha_i R_{t-i} + \sum_{i=0}^5 \beta_i X_{t-i} + \varepsilon_{R,t}, \\ \sigma_{R,t}^2 &= \omega_R + \varphi_R \varepsilon_{R,t-1}^2 + \phi_R \sigma_{R,t-1}^2, \\ X_t &= \gamma_0 + \sum_{i=1}^5 \gamma_i R_{t-i} + \sum_{i=1}^5 \delta_i X_{t-i} + \varepsilon_{X,t}, \\ \sigma_{X,t}^2 &= \omega_X + \varphi_X \varepsilon_{X,t-1}^2 + \phi_X \sigma_{X,t-1}^2. \end{aligned} \tag{3.1}$$

where $\omega_R > 0$, $\omega_X > 0$, $\varphi_R \geq 0$, $\varphi_X \geq 0$, $\phi_R \geq 0$ and $\phi_X \geq 0$.

The rest of the variables in the model (3.1) above keep the same definition as Chapter 2. The variable denoted as R is the difference in the natural logarithm of the mid-quotes for two successive transactions,

$$R_t = \ln M_t - \ln M_{t-1}, M_t = (Q_t^b + Q_t^a)/2,$$

where $Q_t^{b,a}$ are the bid and ask quotes at time t . The variable denoted X is the signed volume of the trade and the trade sign is defined by the classification of [Lee & Ready \(1991\)](#). The quote return variable denoted R is rescaled by multiplying 10^3 and the trade variable denoted X is rescaled by dividing by 10^4 . The coefficients α_i and δ_i are autocorrelations in quote revision and signed volume respectively. The coefficients β_i indicate the impact on quote revision subsequent to each trade. The coefficients γ_i capture Granger causality of lagged quote revision on trades.

Moreover, the chapter includes models from the family of first-order GARCH processes in the model selection exercise in Section 3.4.2. Based on the information criteria selection, they are compared with each other to examine their performance in capturing the time-varying volatility. These models are the GARCH-in-Mean model of [Engle *et al.* \(1987\)](#) [GARCH-M henceforth], the Integrated GARCH of [Engle & Bollerslev \(1986\)](#) [IGARCH henceforth], the Exponential GARCH of [Nelson \(1991\)](#) [EGARCH henceforth] and the Threshold GARCH of [Zakoian \(1994\)](#) [TGARCH henceforth]. All models are estimated based on the same sample stocks for both addition and deletion events and each criteria information is the average across all the stock/events.

3.2.2 The BEKK Model

The rationale to extend the univariate GARCH model to a multivariate framework in financial applications is that it is considered important to be able to predict dependence in the co-movements of volatility and covariance. To ensure positive definiteness of the conditional covariance matrix, [Engle & Kroner \(1995\)](#) introduced a parameterisation of the conditional variance matrix which has become known as the BEKK model. This has advantages when compared to other MGARCH specifications such as the VEC-GARCH model of [Bollerslev *et al.* \(1988\)](#). It achieves the positive definiteness of the conditional covariance matrix by formulating the

model in such a way that this property is implied by the model structure. Unlike the dynamic conditional correlation model of Engle (2002), which estimates the time-varying correlations directly, the BEKK specification allows for time-varying correlations and also for interactions between the variances in a lead-lag framework. As an MGARCH model, the BEKK model is selected in this chapter based on a number of important considerations. The BEKK model overcomes difficulties commonly associated with the VEC-GARCH model. The VEC-GARCH model has two main problems. First, the number of parameters to be estimated under the VEC-GARCH model is large. According to Bauwens *et al.* (2006), the number of parameters is $N(N + 1)(N(N + 1) + 1)/2$ (for $N = 2$ gives 24) where N denotes the number of variables in the model. For the BEKK model, the number of parameters to be estimated is reduced to $N(5N + 1)/2$ (for $N = 2$ gives 11). The BEKK model thus reduces computational demands and improves the efficiency of parameter estimation. Secondly, the BEKK model guarantees that the conditional covariance matrix H_t is positive definite. Furthermore, the BEKK model is superior theoretically to its diagonal model counterparts where each element of the matrix depends only on its own lagged values of shocks and volatility. The conditional covariance matrix in the BEKK model is estimated using a quasi maximum likelihood approach (Bollerslev & Wooldridge, 1992).

The most general specification for the BEKK model is as follows

$$H_t = CC' + \sum_{j=1}^Q \sum_{k=1}^K A'_{kj} u_{t-j} u'_{t-j} A_{kj} + \sum_{j=1}^P \sum_{k=1}^K B'_{kj} H_{t-j} B_{kj},$$

where A_{kj} , B_{kj} and C are $N \times N$ parameter matrices, and C is a lower triangular matrix and the notation $'$ denotes transpose. This model specification ensures that the conditional covariance matrix H_t is at least positive semi-definite. It is generally assumed that $K = 1$ to ensure identifiability. It is further assumed in financial applications that $P = Q = 1$. In this case the first-order BEKK model is

$$H_t = CC' + A'u_{t-1}u'_{t-1}A + B'H_{t-1}B.$$

In matrix form, it may also be written as

$$\begin{bmatrix} h_{11,t} & h_{12,t} \\ h_{21,t} & h_{22,t} \end{bmatrix} = CC' + A' \begin{bmatrix} u_{1,t-1}^2 & u_{1,t-1}u_{2,t-1} \\ u_{2,t-1}u_{1,t-1} & u_{2,t-1}^2 \end{bmatrix} A + B' \begin{bmatrix} h_{11,t-1} & h_{12,t-1} \\ h_{21,t-1} & h_{22,t-1} \end{bmatrix} B,$$

with

$$C = \begin{bmatrix} c_{11} & 0 \\ c_{21} & c_{22} \end{bmatrix}, A = \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{bmatrix}, B = \begin{bmatrix} b_{11} & b_{12} \\ b_{21} & b_{22} \end{bmatrix}.$$

The matrix C is lower triangular, A and B are ARCH and GARCH parameter matrices respectively. In the BEKK specification, each conditional variance and covariance in H_t is modelled as a function of lagged conditional variances and covariances, lagged squared innovations and the cross-product of the innovations. Volatility is transmitted between quote returns and signed volume through two channels represented by the off-diagonal parameters in the ARCH and GARCH matrices: a symmetric shock $u_{ii,t-1}$ and the conditional variance $H_{ii,t-1}$.

The BEKK model applied in this chapter is based on the reduced form of the VAR model with the assumption of conditional variance and covariance. The bivariate VAR model with a first-order BEKK representation in matrix form shows is as follows

$$\begin{aligned} Y_t &= \Psi_0 + \sum_{i=1}^5 \Psi_i Y_{t-i} + u_t, \\ H_t &= CC' + A'u_{t-1}u'_{t-1}A + B'H_{t-1}B. \end{aligned} \tag{3.2}$$

where

$$Y_t = \begin{bmatrix} R_t \\ X_t \end{bmatrix}, \Psi_j = \Lambda^{-1}\Phi_j, u_t = \Lambda^{-1}\varepsilon_t,$$

with

$$\Lambda^{-1} = \begin{bmatrix} 1 & -\beta_0 \\ 0 & 1 \end{bmatrix}^{-1}, \Phi_0 = \begin{bmatrix} \alpha_0 \\ \gamma_0 \end{bmatrix}, \Phi_i = \begin{bmatrix} \alpha_i & \beta_i \\ \gamma_i & \delta_i \end{bmatrix}, \varepsilon_t = \begin{bmatrix} \varepsilon_{R,t} \\ \varepsilon_{X,t} \end{bmatrix}.$$

The rest of the variables in the BEKK model defined at (3.2) above keep the same definition as in Chapter 2. The conditional covariance matrix in the BEKK model is estimated by quasi maximum likelihood method.

In summary, this chapter applies univariate GARCH models and a bivariate BEKK model to examine the effect of short sales on the speed of price adjustment with heteroscedasticity.

3.2.3 Diagnostic Checking

A time series exhibiting conditional heteroscedasticity is said to exhibit autoregressive conditional heteroscedastic (ARCH) effects. Since estimating univariate GARCH and MGARCH models is time-consuming, both in terms of computations and programming, it is desirable to check whether the data present evidence of ARCH effects. There are two kinds of specification tests; namely univariate tests applied independently on each series and multivariate tests applied to the vector series as a whole. These diagnostics are summarised in this section.

Normality Test

The normality test is used to detect misspecification of the ARCH models. Based on skewness and kurtosis, [Bera & Jarque \(1982\)](#) proposed the test statistic

$$JB_T = \left[\frac{T}{6} S^2 + \frac{T}{24} K^2 \right],$$

where T is the number of observations and

$$S = \frac{\sqrt{T} \sum_{t=1}^T \hat{e}_t^3}{\left(\sum_{t=1}^T \hat{e}_t^2 \right)^{\frac{3}{2}}}, K = \frac{T \sum_{t=1}^T \hat{e}_t^4}{\left(\sum_{t=1}^T \hat{e}_t^2 \right)^2}.$$

The Jarque-Bera test is distributed as $\chi_{(2)}^2$. When a GARCH model is estimated, the Jarque-Bera test uses the standardised estimated residuals, $\hat{e}_t = \hat{e}_t / \hat{\sigma}_t$. Skewness (S) and kurtosis (K) may be tested separately using the two components of JB_T , each of which is distributed as Chi-squared with one degree of freedom.

Portmanteau Q Test

For nonlinear time series models, the portmanteau Q test statistic based on squared residuals is used to test for independence for the series ([McLeod & Li, 1983](#))

$$Q(q) = T(T+2) \sum_{i=1}^q \frac{r(i; \hat{e}_t^2)}{(T-i)},$$

where T is the number of observations and

$$r(i; \hat{\varepsilon}_t^2) = \frac{\sum_{t=i+1}^T (\hat{\varepsilon}_t^2 - \hat{\sigma}^2)(\hat{\varepsilon}_{t-i}^2 - \hat{\sigma}^2)}{\sum_{t=1}^T (\hat{\varepsilon}_t^2 - \hat{\sigma}^2)^2}, \hat{\sigma}^2 = \frac{1}{T} \sum_{t=1}^T \hat{\varepsilon}_t^2.$$

This Q statistic is used to detect the nonlinear effects present in the residuals. The GARCH (p, q) process can be considered as an ARMA $(\max(p, q), p)$ process. The Q statistic calculated from the squared residuals can be used to identify the order of the GARCH process.

Lagrange Multiplier Test for ARCH Disturbances

Engle (1982) proposed a Lagrange multiplier test for ARCH disturbances. The test statistic is asymptotically equivalent to the test used by Breusch & Pagan (1979). Engle's Lagrange multiplier test for the q^{th} order ARCH process is written as

$$LM(q) = \frac{TW'Z(Z'Z)^{-1}Z'W}{W'W},$$

where T is the number of observations, W is a vector of standardised squared residuals

$$W = \left(\frac{\hat{\varepsilon}_1^2}{\hat{\sigma}^2}, \dots, \frac{\hat{\varepsilon}_T^2}{\hat{\sigma}^2} \right)',$$

and

$$Z = \begin{bmatrix} 1 & \hat{\varepsilon}_0^2 & \cdots & \hat{\varepsilon}_{-q+1}^2 \\ \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots \\ 1 & \hat{\varepsilon}_{T-1}^2 & \cdots & \hat{\varepsilon}_{T-q}^2 \end{bmatrix}.$$

The presample values $(\varepsilon_0^2, \dots, \varepsilon_{-q+1}^2)$ have been set to 0. The $LM(q)$ and Q statistics are computed from the OLS residuals assuming that disturbances are white noise. The $LM(q)$ and Q statistics have an approximate $\chi_{(q)}^2$ distribution under the white-noise null hypothesis.

Tests for Multivariate GARCH Models

Compared to the diagnostic tests devoted to univariate models, there are few tests on ARCH effects specific to multivariate models. The most widely used to detect whether the model residuals remain correlated are probably the Box-Pierce/Ljung-Box portmanteau Q tests. A multivariate version of the Ljung-Box test statistic following [Hosking \(1980\)](#) is given by

$$LB(L) = T(T + 2) \sum_{j=1}^L (T - j)^{-1} \text{trace} \left\{ \hat{P}_{0j} \hat{P}_{00}^{-1} \hat{P}'_{0j} \hat{P}_{00}^{-1} \right\},$$

where $\hat{P}_{0j} = T^{-1} \sum_{t=j+1}^T \hat{u}_t \hat{u}'_{t-j}$, \hat{u}_t is the estimated vector of residuals at time t , L is the order of autocorrelation, and T is the number of observations.

Under the null hypothesis that the residuals are uncorrelated, the Ljung-Box test statistic has approximately a Chi-squared distribution with $n^2(L - p)$ degrees of freedom where n is the number of equations and p is the lag length in the model. It is important to note that the test can be implemented only when the order of autocorrelation is higher than the lag length in the model, i.e. $L > p$.

3.3 Data Description

This chapter uses the same intraday data of Chapter 2, which covers a period of 10 years from May 2001 to May 2011. An addition (deletion) event is defined when an individual stock is added to (removed from) the D-list for short sales from the effective date. Re-entry (re-quit) of an individual stock is considered as a separate addition (deletion) event. The 10-year period intraday data excludes unit trusts, exchange traded funds, mutual funds, investment companies and stocks traded on the Growth Enterprise Market (GEM). First time IPO firms in the addition events are excluded as there are inadequate observations in the pre-event period. Stocks with options and futures written are excluded as option and future trading can be considered as an alternative to short selling but at a lower cost ([Diamond & Verrecchia, 1987](#)).

The estimation in this chapter contains two models under heteroscedasticity including GARCH (1, 1) and BEKK (1, 1) models, using all transactions in the 60-day period before and after the event date in the 10-year period. It should be noted that there is no requirement to specify weighted GARCH or BEKK models. This is because the variance dynamics equation implicitly allows for the effect of the time gap between trades and it can therefore be argued that weighted GARCH or BEKK models would be over specified as weighted methodology is considered as an alternative solution to heteroscedasticity. Stocks for which estimation errors occurred using GARCH and BEKK models are excluded, which results in the reduction of the sample of this chapter to 1,082 addition events and 690 deletion events¹².

¹²For a small number of stock/events, the software fails to achieve convergence for BEKK or GARCH estimation.

3.4 Results

This section presents the results using the 60-day trade data based on 1,082 addition events and 690 deletion events with two model estimations including GARCH and BEKK. The SS test used in this chapter is defined the same as Chapter 2 and the significance level is 1%. Section 3.4.1 reports the diagnostic tests for ARCH effects for the univariate model (OLS) and the multivariate model (VAR) respectively using the diagnostic checking methods introduced in Section 2.3 above. Section 3.4.2 reports the selection criteria for different GARCH models to approve the model selections. Section 3.4.3 reports the model validation tests using the same procedure in Chapter 2. Section 3.4.4 presents the results about changes in model parameters as a result of addition/deletion event. Sections 3.4.5 demonstrates the results on dynamics and p-values of model parameters for both events. Sections 3.4.6 reports the Z scores for tests on difference in model parameters for both events.

3.4.1 Diagnostic Tests for ARCH Effects

For addition events, Table 3.1 reports the diagnostic tests results based on OLS, showing the percentage of event stocks at the corresponding significance levels which are 0.1%, 1% and 5%. Any results with a significance level greater than 5% are classified as neutral. The test is used to test for normality of residuals. The Portmanteau Q test and the Lagrange Multiplier Test (LM test) are used to test the significance of ARCH effects which is concerned with a relationship within the heteroscedasticity. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for R and X equations in the model before and after the event. Bold format in the table denotes the greatest percentage value.

[Insert Table 3.1 about here]

Table 3.1 shows that the residuals of nearly 99% of the sample stocks do not follow a normal distribution. At least 67% and 43% of the sample stocks experience ARCH

effects at a significance level of 1% as shown by the Portmanteau Q test and the LM test for R and X equations, respectively. Furthermore, there are more stocks with corresponding significant ARCH effects in the equation before and after the addition events compared with the X equation. The results in Table 3.2 report the similar results for deletion events. Although the percentage of event stocks with significant ARCH effects reduces in the X equation compared to the addition events, there are still more than 32% of the sample stocks in the X equation and 68% in the R equation which present evidence of ARCH effects at the significance level of 1%.

[Insert Table 3.2 about here]

To perform diagnostic checking for MGARCH models, we use the residuals from the VAR model of Chapter 2 and apply the multivariate Portmanteau Q test to test whether correlation remains. The multivariate Portmanteau Q test is valid only when lags are larger than the model's lag order, which is 5 in this thesis. In addition, the univariate Normality Test (JB test) is also used to test for normality of residuals in each equation. The F tests for AR disturbance (AR test) are used to test whether the residuals of the univariate $AR(1)$, $AR(1, 2)$, $AR(1, 2, 3)$ and $AR(1, 2, 3, 4)$ models are uncorrelated. The F test for autoregressive conditional heteroscedastic disturbances ($ARCH$ test) is used to test whether the residuals have equal covariances. Table 3.3 reports the diagnostic tests for ARCH effects under the VAR model for addition events.

[Insert Table 3.3 about here]

The JB test results in Table 3.3 show that nearly all the sample stocks' residuals in R and X equations do not follow a normal distribution. All residuals of the univariate $AR(1)$, $AR(1, 2)$, $AR(1, 2, 3)$ and $AR(1, 2, 3, 4)$ models are uncorrelated. The univariate $ARCH$ test shows that more stocks' residuals exhibit significantly unequal covariances in R equations than that in X equations. The multivariate

Portmanteau Q test indicates that there are cross correlations remaining on the model residuals for more than approximately half of the addition stocks at a significance level of 1%. Table 3.4 reports the similar results for ARCH effects under the VAR model for deletion events. At least 71% (35%) of the sample stocks have unequal covariances at the significance level of 1% in the R (X) equation during the event according to the univariate *ARCH* test. The percentage from the multivariate Portmanteau Q test for cross correlations of residuals at significance level of 1% remains between 25% and 37% during deletion events.

[Insert Table 3.4 about here]

In a sum, the results from the diagnostic tests under univariate and multivariate frameworks suggests that in general more than half of sample stocks experience significant ARCH effect during both events under OLS and VAR models. The only exception is those stocks under VAR for deletion events with percentage of 31% for ARCH effect at the significance level of 1%. Therefore, GARCH and MGARCH frameworks are applicable with OLS and VAR to capture dynamic volatility and comovements of volatility in the analysis on the effect of short sales on price efficiency.

3.4.2 Selection Criteria for Different GARCH Models

In addition to the diagnostic tests for ARCH effects, this section provides the model selection criteria for different GARCH models. Model selection covers GARCH (1, 1), IGARCH (1, 1), EGARCH (1, 1), GARCH-M (1, 1) and TGARCH (1, 1) models. The selection criteria used are Akaike's information criterion (*AIC*), corrected Akaike's criterion (*AICC*), Schwartz Bayesian criterion (*SBC*), Hannan-Quinn criterion (*HQC*) and R^2 . Results are reported for both R and X equations and for addition and deletion events. Each model selection criterion is calculated for each stock for each event using the 60-day estimation window. The results shown in the tables are averages computed across all stock/events.

[Insert Tables 3.5 and 3.6 about here]

Tables 3.5 & 3.6 present the selection criteria for different GARCH models for addition and deletion events. Among these competing GARCH models, the model deemed to be the best model has the smallest value of the information criterion in question. As both Tables 3.5 and 3.6 shown, the average differences in the value for each selection criterion are small. Therefore, no single GARCH model appears to be systematically superior and so the standard GARCH (1, 1) model is used along with the standard BEKK (1, 1) model.

3.4.3 Model Validation

A consistency check of the parameters is conducted using the same procedure as in Chapter 2. The Z statistic defined in Chapter 2 has a standard normal distribution under the null hypothesis $H_0 : \Omega = 0$ and rejection of the null hypothesis could indicate inconsistency of the model. Table 3.7 reports the percentage of occasions for which the model consistency test based on equation (2.2) yields results at the specified level of significance under GARCH (1, 1) and BEKK (1, 1) models as well as four models in Chapter 2 including OLS, WLS, VAR and WVAR for both addition and deletion events. Compared with the results under models from Chapter 2, the table shows that the model is not validated at the 0.5% level of significance for between 2% and 5% of all stock/events. The overall Z scores in the last column of the table are all less than one in magnitude, indicating that overall the models with GARCH and BEKK extensions may be considered to be validated. Tables for the detailed number of model validation under GARCH and BEKK are available in Appendix E.

[Insert Table 3.7 about here]

3.4.4 Changes in Model Parameters

Changes in the estimated parameters using the GARCH and BEKK models are reported in this section. There is also a comparison with the corresponding tables from Chapter 2 at the end of the section. Table 3.8 is related to addition events using the GARCH (1, 1) model. The table has four vertical sections and two horizontal panels. The first panel shows results for individual parameters. The second panel shows results for the sums of the α s (A), β s (B), γ s (Γ) and δ s (Δ) respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the addition event. The second vertical section shows the corresponding results after the addition event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in the before and after period for each variable. The last vertical section shows the results for the differences in parameters during addition events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during addition events. Differences are calculated by absolute value of model parameters between after and before addition events. The significance level is defined at the 1%. Cells in which there is a substantial number (108, 10% of the 1,082 addition events or 69, 10% of the 690 deletion events) of statistically significant changes are shown in bold.

[Insert Table 3.8 about here]

Figures in bold in Table 3.8 show that there are a great number of significant changes in the first two lags of α_i (measuring the quote autocorrelations), the first lag of β_i with the current β_0 (measuring the price impact of trades) and the first lag of δ_i (measuring the trade autocorrelations). For instance, 764 stocks (70.6% of addition stocks) experience significant changes in β_0 including 344 decreases and 420 increases. The individual changes result in many corresponding significant changes in the sums (A , B and Δ) except the sum Γ which measures the Granger causality of aggregated lagged quote revision on trades. Overall, the coefficients α_i , β_i , δ_i

and their sums exhibit a substantial number of changes during the addition events using the GARCH (1, 1) framework. The Chi-squared scores reports that none of the sets of parameters before and after the addition events experience significant differences. This means that the signs of model parameters for the majority stocks are consistent during the event although many significant changes including both increase and decrease occurred.

The corresponding results for deletion events by GARCH (1, 1) are presented in Table 3.9 using 69 stocks (10% of the 690 deletion events) as a threshold. The table shows the similar results with smaller percentage of significant changes for fewer parameters compared with the addition events.

[Insert Table 3.9 about here]

For the BEKK model, figures in bold in Table 3.10 show that a number of stocks are significantly affected during the addition events and these significant increases and decreases occurred among parameters measuring quote returns autocorrelations, price impact of trades and trade autocorrelations expect for those measuring Granger causality. In addition, more lagged parameters measuring price impact of trades and trade continuity experience significant decline rather than significant increase and it suggests that the magnitude of more parameters gets weakened by short sales under BEKK framework compared with GARCH. The insignificant Chi-squared scores approve the consistency in the sign of parameters under BEKK framework during the addition events. Table 3.11 shows the corresponding results for deletion events under the BEKK. The table provides similar results that there are fewer significant changes and fewer parameters affected during deletion events rather than addition events.

[Insert Tables 3.10 and 3.11 about here]

Based on the results from the GARCH and BEKK models, it is concluded that there is a number of stocks with significant parameter changes when time-varying volatility and co-movements of volatility are incorporated in the underlying VAR model. For those stocks that are significantly affected, the parameters exhibiting substantial changes are those measuring quote revision autocorrelations, price impact of trades and trade autocorrelations. However, the overall effect of short sales on model parameters is not clear as significant changes contain positive and negative ones during both events.

Table 3.12 presents a comparison of the summary results for the sums denoted A , B , Γ and Δ from Tables 3.8 and 3.10 with the corresponding tables under OLS and VAR in Chapter 2 for addition events. The table entries are shown as percentages as the size of the data sets for GARCH and BEKK models are not the same as those for OLS and VAR models. The table shows that the percentage changes are broadly similar for all four models for addition events. It can be noted however that the summary results under OLS are somewhat different from the other three models especially in the sums of parameters A and Δ measuring the speed of price adjustment. It is also reported that GARCH and BEKK models capture more significant decreases and increases during the event in all aggregated parameters compared with OLS and VAR and this finding suggests that models with GARCH and BEKK extensions are more powerful in capturing the changes in the autocorrelations in quote returns and trades on the high frequency basis. Table 3.13 reports the similar comparison results for deletion events that there are more significant changes in parameters are observed under models with consideration of heteroscedasticity during the event.

[Insert Tables 3.12 and 3.13 about here]

3.4.5 Parameter Dynamics

In this section, we carry out a study of parameter dynamics to examine the consistency in the sign of model parameters before and after addition and deletion events. Table 3.14 shows parameter dynamics using the GARCH model for both addition and deletion events. Each panel shows the results of estimates of the sums of the α s, β s, γ s and δ s respectively. Rows of each panel are before the event and columns are after it. “N” and “P” denotes negative and positive respectively. All other changes at a significance level of which is greater than 5% are classified as neutral. The horizontal and vertical direction of each significance category denote before and after the event, respectively. The table is divided vertically into two sections for addition and deletion events. The predominant cells are shown in bold¹³.

In the first row of the addition section, 92.98% of stock/events have an estimated value of parameter A (quote revision autocorrelations) which is significantly less than zero at 1% level before and after the addition event. In the deletion section, the last row of the table shows that 2.61% of sample stocks have an estimated value of parameter Δ (trade autocorrelations) which is positively significant at the 5% level before the event and is positively significant at the 1% level after the event. The results show that the majority of the sums of estimated model parameters, with the exception of the sum Γ (measures Granger causality), have consistent signs during both addition and deletion events for both events. Furthermore, for the sums of A , B and Δ , the table points that the percentages of the stocks with consistent signs at the significance level of 1% are higher for the addition events than those for the deletion events. It implies that more stocks exhibit consistency in the sign of these sums of parameters during the addition events rather than deletion events. The results for the BEKK model in Table 3.15 are consistent with those from the GARCH model. It is concluded that the majority of stocks have significant consistency in the signs of the sums of A , B and Δ for both events and both models.

¹³More detailed tables which cover individual parameters are available on request.

The differences in sign consistency between addition and deletion events can be linked to prospect theory in behavioural finance introduced by [Kahneman & Tversky \(1979\)](#), namely that the reaction to good news (addition events) should be different from the reaction to bad news (deletion events). [Kahneman & Tversky \(1979\)](#) explain particular forms of irrational behaviours of investors by stating that “*The value function is normally concave for gains, commonly convex for losses, and is generally steeper for losses than for gains*”. A consequence of behavioural finance is that investors will react irrationally to deletion events which are perceived as bad news and which make investors nervous and lack confidence. It is conjectured that these behavioural effects are the cause of less parameter consistency for deletion events.

[Insert Tables [3.14](#) and [3.15](#) about here]

Table [3.16](#) reports a summary of key cells of parameter dynamics for the sums denoted A , B , Γ and Δ under OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) for addition and deletion events. The table entries are shown as percentages of stocks with consistency in the sign of model parameters before and after both events. The table shows that there are more stocks with consistency in the sign of parameters during addition events than deletion events. For each event, less stocks exhibit consistency when models are estimated on a time-weighted basis especially for the sums A , B and Δ . Furthermore, it is noted that the consistency in the sign of autocorrelations in quote returns (A) and trades (Δ) remains stronger under GARCH model compared with OLS model.

[Insert Table [3.16](#) about here]

3.4.6 Z Scores for Tests of Difference in Model Parameters

The results from Section 3.4.4 reveal that more stocks are observed to experience significant changes in the model parameters under GARCH and BEKK frameworks during both events however it is not clear to determine the effect of short sales on the speed of price adjustment as these changes contain both increases and decreases in the corresponding parameters.

To investigate the overall effect of short sales on price efficiency, Table 3.17 summarises the Z scores for tests of the differences in estimated model parameters using 60 days of trade data under GARCH (1, 1) and BEKK (1, 1) frameworks as well as four models from Chapter 2 during addition events. Differences are calculated by absolute value of model parameters between after and before both events (After minus Before). The Z scores for the differences are computed using the SS test statistic at (2.10) defined in Chapter 2. There are six sub-sections corresponding to the test results for the six models. Z scores shown in bold in the table are significantly different from zero at the 1% level for a two-tailed test.

[Insert Table 3.17 about here]

To compare GARCH with BEKK, it is observed that more individual and aggregated model parameters especially for α_i and δ_i are affected significantly under BEKK model compared with GARCH model during the addition events. For quote revision autocorrelations, it is found that more individual lagged parameters α_i experience changes at the 1% significance level under BEKK model. Only the aggregated quote autocorrelation parameter A under BEKK model experiences a significant decrease during the addition events and such reduction in the magnitude of quote autocorrelations indicates a faster speed of price adjustment enhanced by short sales. For trade continuity, the majority of significant decreases in the first three lagged trade autocorrelations, δ_i are observed in BEKK estimations. The aggregated trade autocorrelation parameters Δ decrease at the significance level

of 1% under both GARCH and BEKK models and this confirms that short sales fasten the speed of price adjustment by weakening autocorrelations in trades. The majority of the lagged parameters β_1, \dots, β_5 do not have significant changes under BEKK and GARCH models and these results in the insignificant differences in the aggregated price impact parameter B during the event. The finding shows that the effect of short sales on price impact of trades disappears when heteroscedasticity is considered. Significant increases are found in the individual and aggregated values of Granger causality under GARCH and BEKK models.

By comparing results above with those under models from Chapter 2 for the addition events, the aggregated trade continuity decreases throughout six models and it suggests that trades become less correlated after being added to the D-list with robustness under all models. The aggregated autocorrelations in quote returns decrease at the significance level of 1% under WVAR and BEKK and at the significance level of 5% under VAR. It implies that short sales improve price efficiency by weakening the autocorrelations in quote returns while the effect is only captured when the model is estimated simultaneously. The aggregated price impact of trades increase significantly under WLS and WVAR and it suggests that the price impact of trades is enhanced by short sales when time-duration is considered. The Granger causality experience significant changes at the significance level of 1% under all models during the addition events. However, it is not clear to conclude the effect of short sales on Granger causality as these changes contain both increases under unweighted models and decreases under weighted models.

Table 3.18 summarises the corresponding Z scores for test of the differences in model parameters under six models for the deletion events.

[Insert Table 3.18 about here]

For results under GARCH and BEKK models, there are less model parameters with significant changes compared with the addition events. For quote revision, a decline in the value of the first lagged parameter α_1 at the 1% significance level

is observed and it is also the main attribution to the significant decline in the aggregated autocorrelations in the sum A under BEKK model. It shows that quote returns are more correlated if stocks are removed from the short-selling list and it is consistent with the hypothesis that short sales constraints delay the speed of price adjustment. The individual and aggregated trade autocorrelations do not have significant changes under GARCH and BEKK and it indicates that trade continuity is not affected by short sales significantly during the deletion events. Similar results are also obtained for price impact of trades for deletion stocks. There is a significant increase in the aggregated Granger causality attributed by the increase in values of lagged parameter γ_i under GARCH and BEKK models. However, combined with the findings on Granger causality parameters for the addition events, the effect of short sales on Granger causality remains unclear.

Compared with results under models from Chapter 2, the aggregated quote autocorrelations increase under the majority of models. It indicates that the quote autocorrelations are the most affected parameters and the speed of price adjustment gets slower during the deletion events. However, there is no significant evidence that trade continuity is affected when stocks lose their eligibility of short sales and the findings are robust under all models. Similarly, price impact of trades do not change significantly under these six models. Consistent with the results from the addition events, the effect of short sales on Granger causality is inconclusive.

To conclude, it is found that autocorrelations in quote returns and trades decrease significantly during the addition events while only quote autocorrelations have significant increase during the deletion events under both GARCH and BEKK models. The results suggest that the faster speed of price adjustment after the addition events is attributed to the decreased autocorrelations in both quote returns and trades however the slower speed of price adjustment during the deletion events is mainly due to the increased quote autocorrelations. However, there is no remarkable evidence to support the effect of short sales on price impact of trades and Granger causality when time-varying variance is considered.

The overall results under six models indicate that more parameters are affected during the addition events than the deletion events. Among these changes, trade continuity is the most affected parameter weakened by short sales than quote autocorrelation during the addition events while quote autocorrelation is the most affected one to be strengthened during the deletion events. Tables for p-values for tests of individual and aggregated model parameters before and after both events under GARCH and BEKK models are available in Appendix E.

3.5 Conclusions

This chapter carries out an investigation on the effect of short sales on the speed of price adjustment by extending the VAR model to univariate GARCH and MGARCH models with consideration of heteroscedasticity which may be caused by high frequency trade-by-trade data in the Hong Kong stock market. In this chapter, it applies a bivariate VAR model with the extensions of GARCH (1, 1) and BEKK (1, 1) models. The univariate GARCH model captures dynamic volatility while the BEKK model captures co-movements of volatility between quote returns and trades. Furthermore, the BEKK model has advantages compared with other MGARCH models as its quadratic forms ensure the positive definiteness of the conditional covariance matrices.

Based on the intraday data over 10 years, the main conclusions of this chapter are as follows. First, there are a number of stocks for which certain parameters in the model do exhibit significant changes including increases and decreases in both directions during both events under GARCH and BEKK models. To compare with OLS and VAR models assuming that the residuals have homogeneous variance, more significant changes are captured under models with consideration of heterogeneous variances on a high frequency basis. Second, the results from parameter dynamics suggests that the majority of model parameters stay consistent during both events. For example a positive parameter that is statistically significant will generally remain in the same category, even if the estimated value changes as a

result of the addition/deletion event. It is reported that there is a stronger consistency especially for the parameters A and Δ as proxies for price efficiency under GARCH model rather than OLS model.

For the overall effect of short sales on the speed of price adjustment, the results show that short sales play an important role in recovering price efficiency during both events when volatility and co-volatility are considered and the findings are consistent with those under the models from Chapter 2. Combined with the results for Chapters 2 & 3, it is noted that the parameters which measure trade continuity are affected most during the addition events and the trade stickiness is weakened by short sales with robustness under all models. The significant decrease in the value of quote autocorrelations under WVAR and BEKK implies that quote returns are less correlated after being added to the D-list which is consistent with those from [Boehmer & Wu \(2013\)](#). For the deletion events, quote autocorrelations are the most affected parameters under the majority of model estimations and the significant increase in the value of parameters indicates that quote returns are more likely to follow the historical quote returns when stocks are not allowed for short sales. The results also show that trade continuity is not affected significantly during the deletion events.

Overall, results from six models show that short sales enhance price efficiency by reducing (increasing) autocorrelations in quote returns and trades during the addition (deletion) events. Moreover, parameters as proxies for price efficiency including autocorrelations in quote returns and trades are affected to different extent under a variety of models during different events. The combined results reveal that WVAR and BEKK models, that is models estimated simultaneously and with consideration of heterogeneous variances, have more power in ability to capture the changes in parameters measuring price efficiency than other models especially when stocks are added to the D-list for short sales. However, particularly at the aggregate level, the differences between all six models reported in Chapters 2 & 3 are not great and, as the chapter will show, the univariate or equation by equation approach may be used for some of the more detailed investigations.

The next chapter plans to extend the empirical analyses on the determinants of the effect of short sales on price efficiency by investigating market returns, market capitalisation, trading volume and shorting activities.

Table 3.1: Diagnostic Tests for ARCH effects (OLS Addition Events)

Based on the model defined at (2.1) using 60 days of trade data by OLS for the addition events, the table represents the diagnostic tests for ARCH effects with the percentage at the corresponding significance level of p-value of the total sample size. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The Jarque-Bera test (JB test) is used to test for normality of residuals. The Ljung-Box test (Q test) and the Engles Lagrange Multiplier ARCH test (LM test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for R and X equations in the model before and after the event. Bold format denotes the greatest percentage value.

	Before Addition (R equation)			Before Addition (X equation)			After Addition (R equation)			After Addition (X equation)						
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%				
JB	98.90%	0.10%	0.10%	0.90%	99.50%	0.00%	0.10%	0.40%	98.80%	0.10%	0.10%	1.10%	98.90%	0.10%	0.10%	0.90%
$Q(1)$	60.40%	6.50%	4.60%	28.50%	37.10%	9.80%	8.60%	44.50%	61.90%	4.60%	5.70%	27.80%	33.20%	9.70%	9.20%	47.90%
$Q(2)$	71.60%	4.30%	4.20%	20.00%	46.20%	9.00%	6.80%	38.00%	71.90%	4.30%	3.60%	20.10%	40.80%	9.60%	9.00%	40.60%
$Q(3)$	76.40%	2.50%	3.00%	18.10%	50.30%	8.30%	6.80%	34.60%	75.20%	4.30%	2.70%	17.90%	44.30%	8.30%	7.50%	40.00%
$Q(4)$	78.00%	2.40%	2.00%	17.60%	52.70%	7.50%	7.50%	32.30%	77.40%	2.90%	3.30%	16.40%	47.30%	7.80%	5.90%	39.00%
$Q(5)$	79.40%	2.10%	2.00%	16.50%	53.00%	7.50%	6.60%	32.90%	78.50%	2.90%	3.00%	15.50%	49.80%	5.80%	6.30%	38.20%
$Q(6)$	80.90%	2.10%	1.80%	15.20%	54.60%	6.70%	7.10%	31.60%	79.50%	3.00%	2.40%	15.10%	50.80%	5.60%	6.60%	37.10%
$Q(7)$	81.50%	2.00%	1.90%	14.60%	55.80%	6.90%	5.80%	31.50%	79.90%	3.20%	2.60%	14.40%	50.80%	5.60%	6.10%	37.50%
$Q(8)$	82.30%	1.80%	2.20%	13.80%	56.70%	6.60%	5.90%	30.80%	81.30%	3.00%	2.20%	13.50%	50.90%	5.20%	6.00%	37.80%
$Q(9)$	83.10%	2.00%	1.30%	13.70%	57.00%	6.30%	6.60%	30.20%	81.70%	2.80%	2.20%	13.20%	51.60%	5.40%	5.60%	37.40%
$Q(10)$	83.20%	2.00%	1.60%	13.20%	57.60%	6.30%	5.30%	30.80%	82.20%	2.40%	2.60%	12.90%	51.50%	6.10%	6.40%	36.00%
$Q(11)$	83.40%	1.90%	1.90%	12.90%	58.40%	6.00%	5.20%	30.30%	82.50%	2.70%	2.40%	12.40%	52.40%	5.80%	5.80%	36.10%
$Q(12)$	83.50%	2.30%	1.20%	13.00%	58.80%	5.90%	5.20%	30.10%	82.90%	2.30%	2.60%	12.20%	52.70%	5.80%	6.30%	35.20%

Table continued on the following page.

Table 3.1: Continued from previous page

	Before Addition (R equation)			Before Addition (X equation)			After Addition (R equation)			After Addition (X equation)				
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	ntrl	
$LM(1)$	60.40%	6.50%	4.60%	28.50%	37.10%	8.60%	44.50%	61.90%	4.50%	5.80%	27.80%	33.20%	9.70%	48.00%
$LM(2)$	71.30%	4.30%	4.40%	20.10%	45.80%	7.10%	38.20%	71.30%	4.80%	3.60%	20.30%	40.50%	9.20%	41.20%
$LM(3)$	75.90%	2.80%	2.90%	18.40%	49.50%	7.10%	35.10%	74.70%	4.20%	3.00%	18.10%	43.10%	8.80%	40.30%
$LM(4)$	77.80%	2.30%	1.80%	18.10%	51.60%	8.30%	32.80%	76.30%	3.50%	3.00%	17.10%	46.10%	7.80%	39.50%
$LM(5)$	79.10%	2.30%	2.00%	16.60%	51.60%	7.50%	33.40%	77.20%	3.70%	3.30%	15.80%	48.60%	5.90%	39.00%
$LM(6)$	80.20%	2.50%	2.00%	15.40%	52.60%	8.00%	32.20%	78.40%	3.40%	2.60%	15.60%	48.30%	6.70%	38.40%
$LM(7)$	81.00%	2.10%	2.10%	14.70%	53.90%	7.40%	32.60%	79.10%	2.90%	3.10%	14.90%	49.50%	5.90%	39.00%
$LM(8)$	81.50%	2.40%	2.30%	13.80%	54.70%	7.10%	31.80%	79.90%	3.00%	3.00%	14.00%	49.70%	5.30%	39.10%
$LM(9)$	82.40%	2.00%	1.90%	13.80%	54.70%	6.70%	31.70%	80.40%	2.70%	3.10%	13.80%	49.50%	6.10%	38.60%
$LM(10)$	82.30%	2.30%	2.00%	13.40%	56.20%	6.10%	31.40%	80.80%	2.60%	3.00%	13.60%	50.40%	5.70%	38.10%
$LM(11)$	82.80%	2.00%	2.00%	13.20%	56.50%	6.20%	31.30%	80.80%	2.80%	3.10%	13.20%	50.30%	5.90%	37.50%
$LM(12)$	82.80%	2.70%	1.50%	13.00%	57.10%	6.10%	31.40%	81.30%	2.90%	2.70%	13.10%	50.80%	6.00%	36.80%

Table 3.2: Diagnostic Tests for ARCH effects (OLS Deletion Events)

Based on the model defined at (2.1) using 60 days of trade data by OLS for the deletion events, the table represents the diagnostic tests for ARCH effects with the percentage at the corresponding significance level of p-value of the total sample size. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The Jarque-Bera test (JB test) is used to test for normality of residuals. The Ljung-Box test (Q test) and the Engles Lagrange Multiplier ARCH test (LM test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for R and X equations in the model before and after the event. Bold format denotes the greatest percentage value.

	Before Deletion (R equation)			Before Deletion (X equation)			After Deletion (R equation)			After Deletion (X equation)						
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%				
JB	98.10%	0.50%	0.50%	0.80%	99.20%	0.10%	0.10%	0.50%	95.60%	1.20%	0.80%	2.30%	98.90%	0.10%	0.10%	0.80%
$Q(1)$	65.70%	5.90%	5.90%	22.50%	27.30%	8.50%	8.10%	56.00%	60.20%	8.40%	5.50%	26.00%	23.50%	8.40%	9.20%	58.90%
$Q(2)$	74.60%	4.30%	4.00%	17.20%	33.80%	8.00%	8.10%	50.10%	70.20%	5.10%	4.50%	20.20%	30.10%	8.00%	7.40%	54.50%
$Q(3)$	76.40%	5.10%	2.50%	16.10%	36.70%	7.60%	7.60%	48.20%	73.10%	4.70%	4.30%	18.00%	33.10%	7.70%	6.90%	52.30%
$Q(4)$	78.40%	3.20%	3.60%	14.80%	38.60%	6.90%	5.90%	48.60%	75.00%	4.00%	4.00%	17.00%	35.00%	8.10%	4.70%	52.20%
$Q(5)$	79.00%	3.60%	3.70%	13.70%	39.80%	6.30%	5.90%	47.90%	75.80%	3.80%	3.40%	16.90%	34.80%	8.40%	5.40%	51.50%
$Q(6)$	79.10%	3.30%	3.40%	14.10%	41.80%	6.60%	3.60%	48.10%	75.80%	4.30%	4.30%	15.70%	35.70%	7.40%	7.10%	49.70%
$Q(7)$	78.80%	3.80%	2.90%	14.40%	42.00%	5.80%	5.40%	46.80%	75.70%	4.40%	4.10%	15.80%	36.70%	7.30%	7.00%	49.00%
$Q(8)$	78.80%	4.00%	3.20%	14.00%	42.90%	5.20%	5.40%	46.60%	76.40%	4.70%	2.90%	16.10%	37.50%	6.90%	6.90%	48.80%
$Q(9)$	78.80%	3.70%	3.60%	13.90%	44.00%	4.30%	5.40%	46.40%	76.50%	4.00%	3.40%	16.10%	38.00%	6.50%	7.40%	48.10%
$Q(10)$	79.00%	4.00%	3.20%	13.90%	44.20%	5.10%	4.50%	46.20%	77.50%	2.30%	3.30%	16.90%	38.00%	7.00%	6.60%	48.40%
$Q(11)$	79.40%	3.40%	3.00%	14.10%	44.50%	5.10%	5.10%	45.30%	77.20%	3.00%	3.30%	16.50%	38.90%	5.20%	7.70%	48.20%
$Q(12)$	79.50%	2.90%	3.60%	14.00%	45.20%	4.90%	4.30%	45.60%	76.80%	3.00%	2.70%	17.40%	38.90%	5.40%	6.60%	49.20%

Table continued on the following page.

Table 3.2: Continued from previous page

	Before Deletion (<i>R</i> equation)			Before Deletion (<i>X</i> equation)			After Deletion (<i>R</i> equation)			After Deletion (<i>X</i> equation)						
	0.1%	1%	5%	ntrl	0.1%	1%	5%	0.1%	1%	5%	ntrl	0.1%	1%	5%	ntrl	
<i>LM</i> (1)	65.70%	6.00%	5.80%	22.50%	27.30%	8.40%	8.20%	56.00%	60.20%	8.00%	5.90%	26.00%	23.40%	8.50%	9.20%	58.90%
<i>LM</i> (2)	72.90%	5.50%	4.30%	17.30%	33.20%	8.20%	7.70%	50.80%	68.70%	6.00%	4.50%	20.70%	29.40%	8.10%	7.60%	54.90%
<i>LM</i> (3)	75.80%	5.10%	2.60%	16.50%	36.00%	7.30%	7.40%	49.30%	72.00%	5.50%	4.10%	18.40%	31.90%	7.80%	7.30%	53.00%
<i>LM</i> (4)	77.20%	3.60%	4.10%	15.10%	37.40%	6.70%	6.60%	49.30%	73.40%	4.80%	3.60%	18.30%	33.00%	8.20%	6.60%	52.20%
<i>LM</i> (5)	76.80%	4.10%	4.80%	14.30%	38.30%	6.60%	5.80%	49.30%	73.80%	4.70%	4.40%	17.20%	32.60%	9.30%	6.30%	51.80%
<i>LM</i> (6)	77.50%	3.70%	4.00%	14.80%	39.60%	7.10%	4.30%	49.00%	73.90%	4.80%	4.80%	16.50%	33.90%	8.10%	7.30%	50.70%
<i>LM</i> (7)	77.10%	4.10%	3.60%	15.20%	40.40%	5.80%	5.80%	48.10%	74.00%	4.80%	3.60%	17.60%	34.80%	7.80%	6.90%	50.50%
<i>LM</i> (8)	76.60%	4.90%	3.30%	15.10%	41.10%	5.10%	6.00%	47.80%	74.30%	4.70%	4.00%	17.00%	35.20%	7.70%	6.50%	50.70%
<i>LM</i> (9)	76.60%	4.50%	4.00%	14.80%	42.30%	4.40%	5.90%	47.40%	73.50%	5.80%	3.40%	17.30%	36.00%	6.60%	7.30%	50.10%
<i>LM</i> (10)	76.90%	4.70%	3.60%	14.80%	42.60%	5.50%	4.70%	47.30%	73.50%	5.50%	3.40%	17.60%	36.00%	7.00%	7.30%	49.70%
<i>LM</i> (11)	76.90%	4.10%	4.10%	14.80%	42.70%	5.50%	4.70%	47.10%	73.50%	5.60%	3.20%	17.70%	36.80%	5.20%	7.80%	50.10%
<i>LM</i> (12)	76.60%	4.30%	4.40%	14.70%	43.50%	5.10%	3.80%	47.50%	73.10%	5.90%	3.30%	17.70%	37.20%	4.90%	7.00%	50.80%

Table 3.3: Diagnostic Tests for ARCH effects (VAR Addition Events)

Based on the model defined at (2.1) using 60 days of trade data by VAR for the addition events, the table represents the diagnostic tests for ARCH effects with the percentage at the corresponding significance level of p-value of the total sample size. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The univariate Jarque-Bera test (JB test) is used to test for normality of residuals. The F tests for AR disturbance (AR test) are used to test whether the residuals of the univariate $AR(1)$, $AR(1, 2)$, $AR(1, 2, 3)$ and $AR(1, 2, 3, 4)$ models are uncorrelated. The F test for autoregressive conditional heteroscedastic disturbances ($ARCH$ test) is used to test whether the residuals have equal covariances. The multivariate Portmanteau (Q) test is used to test whether correlation remains on the model residuals. The test is valid only for lags larger than the model's lag order. Bold format denotes the greatest percentage value.

	Before Addition (R equation)			Before Addition (X equation)			After Addition (R equation)			After Addition (X equation)		
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%
JB	99.70%	0.10%	0.10%	99.90%	0.00%	0.00%	100.00%	0.00%	0.00%	99.70%	0.10%	0.20%
$AR(1)$	0.00%	0.00%	0.10%	99.90%	0.00%	0.10%	0.10%	0.00%	0.20%	0.00%	0.00%	100.00%
$AR(2)$	0.00%	0.00%	0.10%	99.90%	0.00%	0.10%	0.10%	0.00%	0.20%	0.00%	0.00%	100.00%
$AR(3)$	0.00%	0.00%	0.10%	99.90%	0.00%	0.10%	0.10%	0.00%	0.20%	0.00%	0.00%	100.00%
$AR(4)$	0.00%	0.00%	0.10%	99.90%	0.00%	0.10%	0.10%	0.00%	0.20%	0.00%	0.00%	100.00%
$ARCH$	64.60%	4.80%	4.80%	38.80%	10.00%	8.60%	63.00%	5.20%	5.70%	35.60%	9.40%	44.30%
	Before Addition											
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%
$Q(6)$	35.40%	11.00%	11.20%	42.40%	33.30%	12.00%	10.60%	44.10%				
$Q(7)$	40.60%	11.30%	12.20%	35.80%	37.80%	11.30%	11.80%	39.10%				
$Q(8)$	43.70%	12.70%	11.40%	32.20%	41.80%	11.90%	10.30%	36.00%				
$Q(9)$	47.00%	12.10%	10.60%	30.30%	43.90%	11.40%	8.90%	35.80%				
$Q(10)$	49.20%	11.50%	10.50%	28.70%	45.10%	11.30%	9.80%	33.80%				
$Q(11)$	50.80%	11.30%	9.80%	28.10%	46.10%	12.50%	8.90%	32.60%				

Table 3.4: Diagnostic Tests for ARCH effects (VAR Deletion Events)

Based on the model defined at (2.1) using 60 days of trade data by VAR for the deletion events, the table represents the diagnostic tests for ARCH effects with the percentage at the corresponding significance level of p-value of the total sample size. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The univariate Jarque-Bera test (JB test) is used to test for normality of residuals. The F tests for AR disturbance (AR test) are used to test whether the residuals of the univariate $AR(1)$, $AR(1, 2)$, $AR(1, 2, 3)$ and $AR(1, 2, 3, 4)$ models are uncorrelated. The F test for autoregressive conditional heteroscedastic disturbances ($ARCH$ test) is used to test whether the residuals have equal covariances. The multivariate Portmanteau (Q) test is used to test whether correlation remains on the model residuals. The test is valid only for lags larger than the model's lag order. Bold format denotes the greatest percentage value.

	Before Deletion (R equation)			Before Deletion (X equation)			After Deletion (R equation)			After Deletion (X equation)		
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%
JB	99.90%	0.00%	0.10%	0.00%	99.70%	0.10%	0.00%	0.00%	0.10%	99.70%	0.10%	0.00%
$AR(1)$	0.00%	0.00%	0.00%	100.00%	0.00%	0.00%	100.00%	0.00%	0.10%	99.90%	0.00%	0.00%
$AR(2)$	0.00%	0.00%	0.00%	100.00%	0.00%	0.00%	100.00%	0.00%	0.10%	99.90%	0.00%	0.00%
$AR(3)$	0.00%	0.00%	0.00%	100.00%	0.00%	0.00%	100.00%	0.00%	0.10%	99.90%	0.00%	0.00%
$AR(4)$	0.00%	0.00%	0.00%	100.00%	0.00%	0.00%	100.00%	0.00%	0.10%	99.90%	0.00%	0.00%
$ARCH$	66.60%	8.40%	5.10%	19.90%	28.70%	10.30%	8.80%	62.80%	8.40%	7.00%	27.70%	55.80%
	Before Deletion			Before Deletion			After Deletion			After Deletion		
	0.1%	1%	5%	ntrl	1%	5%	ntrl	0.1%	1%	5%	ntrl	0.1%
$Q(6)$	16.90%	9.30%	11.50%	62.20%	15.90%	9.60%	10.70%	9.60%	10.70%	10.70%	10.70%	63.70%
$Q(7)$	19.60%	8.80%	13.30%	58.20%	20.20%	9.60%	11.00%	9.60%	11.00%	11.00%	11.00%	59.20%
$Q(8)$	21.40%	8.70%	14.60%	55.40%	22.40%	10.60%	12.80%	10.60%	12.80%	12.80%	12.80%	54.30%
$Q(9)$	23.10%	9.30%	13.00%	54.50%	23.90%	10.70%	13.30%	10.70%	13.30%	13.30%	13.30%	52.10%
$Q(10)$	24.20%	9.80%	11.50%	54.50%	24.90%	10.90%	13.60%	10.90%	13.60%	13.60%	13.60%	50.70%
$Q(11)$	26.40%	9.50%	12.40%	51.80%	25.70%	11.30%	11.70%	11.30%	11.70%	11.70%	11.70%	51.40%

Table 3.5: Selection Criteria for Different GARCH-type Models (Addition Events)

Based on the model defined at (2.1) using 60 days of trade data, the table represents the selection criteria for different GARCH-type models before and after the addition events. The criteria information contains Akaike's criterion (AIC), the corrected Akaike's criterion ($AICC$), Schwartz Bayesian criterion (SBC), Hannan-Quinn criterion (HQC) and R^2 for both R and X equations in the model. Each criterion is calculated for each stock for each event in a 60-day period and averaged across the stocks.

	Before Addition									
	$AIC - R$	$AIC - X$	$AICC - R$	$AICC - X$	$SBC - R$	$SBC - X$	$HQC - R$	$HQC - X$	$R^2 - R$	$R^2 - X$
GARCH (1, 1)	44077.54	76053.78	44077.87	76054.01	44175.13	76145.84	44111.28	76085.59	0.11	0.05
IGARCH (1, 1)	44452.3	76320.78	44452.61	76320.99	44544.22	76406.38	44484.09	76350.37	0.11	0.05
EGARCH (1, 1)	43965.4	76632.21	43965.68	76632.47	44070.31	76730.91	44001.67	76666.33	0.11	0.05
GARCH-M (1, 1)	44152.23	76192.25	44152.51	76192.5	44255.55	76290.83	44187.99	76226.32	0.11	0.05
TGARCH (1, 1)	43881.3	75435.62	43881.56	75435.83	43984.11	75532.96	43916.84	75469.23	0.11	0.05
After Addition										
GARCH (1, 1)	35585.45	57000.63	35585.94	57000.95	35679.93	57089.61	35618.46	57031.7	0.11	0.05
IGARCH (1, 1)	35807.47	57245.23	35807.93	57245.5	35896.3	57328.01	35838.52	57274.12	0.11	0.04
EGARCH (1, 1)	35535.01	57003.79	35535.37	57004.12	35636.35	57099.21	35570.4	57037.11	0.11	0.04
GARCH-M (1, 1)	35582.35	56998.43	35582.83	56998.77	35682.33	57093.52	35617.27	57031.61	0.11	0.05
TGARCH (1, 1)	35599.08	57069.35	35599.38	57069.59	35699.25	57163.29	35634.06	57102.11	0.11	0.05

Table 3.6: Selection Criteria for Different GARCH-type Models (Deletion Events)

Based on the model defined at (2.1) using 60 days of trade data, the table represents the selection criteria for different GARCH-type models before and after the deletion events. The criteria information contains Akaike criterion (AIC), the corrected Akaike criterion ($AICC$), Schwartz Bayesian criterion (SBC), Hannan-Quinn criterion (HQC) and R^2 for both R and X equations in the model. Each criterion is calculated for each stock for each event in a 60-day period and averaged across the stocks.

	Before Deletion									
	$AIC - R$	$AIC - X$	$AICC - R$	$AICC - X$	$SBC - R$	$SBC - X$	$HQC - R$	$HQC - X$	$R^2 - R$	$R^2 - X$
GARCH (1, 1)	19200.48	24903.88	19201.16	24904.35	19283.22	24981.5	19230.59	24932.1	0.14	0.06
IGARCH (1, 1)	19260.08	25021.29	19260.71	25021.68	19337.51	25093.7	19288.26	25047.61	0.13	0.05
EGARCH (1, 1)	19273.67	24903.17	19274.3	24903.69	19361.89	24986.24	19305.75	24933.35	0.14	0.05
GARCH-M (1, 1)	19200.3	24900.54	19201.07	24901.04	19288.19	24982.93	19232.3	24930.47	0.14	0.06
TGARCH (1, 1)	19304.1	25061.64	19304.64	25062	19391.17	25141.56	19335.75	25090.59	0.14	0.06
						After Deletion				
GARCH (1, 1)	19285.75	27151.86	19287.58	27152.64	19366.92	27228.05	19315.27	27179.53	0.15	0.06
IGARCH (1, 1)	19362.15	27229.89	19363.75	27230.6	19438.07	27300.99	19389.76	27255.75	0.14	0.06
EGARCH (1, 1)	19326.26	27155.59	19327.36	27156.41	19412.3	27237.05	19357.51	27185.17	0.15	0.06
GARCH-M (1, 1)	19346.92	27231.89	19348.42	27232.67	19433.27	27313.08	19378.3	27261.37	0.15	0.07
TGARCH (1, 1)	19378.17	27222.99	19378.91	27223.54	19463.28	27301.43	19409.06	27251.38	0.15	0.06

Table 3.7: Model Validation

The table shows the percentage of occasions for which the model consistency test based on equation (2.2) yields results at the specified level of significance under OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) models. The Z scores shown in the last column are computed for all sample stocks using the Scott-Smith test.

		$\leq 0.5\%$	$\leq 1\%$	$\leq 2.5\%$	$\leq 5\%$	$\leq 10\%$	$\geq 90\%$	$\geq 95\%$	$\geq 97.5\%$	$\geq 99\%$	$\geq 99.5\%$	Z	
OLS	Addition	<i>Before</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	-0.512	
		<i>After</i> 0.000	0.000	0.000	0.000	0.000	0.090	0.090	0.090	0.000	0.000	0.136	
	Deletion	<i>Before</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	
		<i>After</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	-0.162	
	Addition	<i>Before</i> 0.090	0.090	0.090	0.090	0.090	0.090	0.000	0.000	0.000	0.000	0.000	-0.295
		<i>After</i> 0.090	0.090	0.090	0.090	0.090	0.180	0.000	0.000	0.000	0.000	0.000	-0.116
WLS	Addition	<i>Before</i> 0.140	0.140	0.140	0.140	0.140	0.000	0.000	0.000	0.000	0.000	0.000	-0.432
		<i>After</i> 0.140	0.140	0.270	0.410	0.410	0.140	0.140	0.140	0.140	0.140	0.000	-0.134
	Deletion	<i>Before</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
		<i>After</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	-0.157
	Addition	<i>Before</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.055
		<i>After</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.065
VAR	Deletion	<i>Before</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.010
		<i>After</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	-0.108
	Addition	<i>Before</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.042
		<i>After</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.185
	Deletion	<i>Before</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.062
		<i>After</i> 0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	-0.002
GARCH (1, 1)	Addition	<i>Before</i> 1.570	2.500	3.230	5.730	1.940	1.480	1.200	1.020	0.740	0.550	0.280	-0.173
		<i>After</i> 1.880	2.030	2.170	3.040	2.750	2.320	1.590	1.300	1.160	1.010	0.430	-1.102
	Deletion	<i>Before</i> 1.450	1.880	2.030	2.320	1.590	1.300	0.870	0.870	0.870	0.720	0.720	-0.048
		<i>After</i> 2.870	3.700	3.970	5.270	3.330	2.770	2.130	1.480	1.390	1.390	0.830	-0.764
	Addition	<i>Before</i> 2.870	3.970	4.900	6.560	4.900	3.880	3.050	2.590	2.310	2.310	1.760	0.062
		<i>After</i> 1.450	1.880	2.610	3.910	3.480	2.610	1.590	1.300	1.300	1.010	0.720	-0.085
BEKK (1, 1)	Deletion	<i>Before</i> 1.450	1.880	2.610	3.910	3.480	2.610	1.590	1.300	1.300	1.010	0.720	-0.085
		<i>After</i> 1.590	1.880	2.610	3.330	2.610	1.740	1.300	1.300	1.160	0.870	0.593	

**Table 3.8: Frequency of Changes in
Model Parameters for Addition Events (GARCH)**

Based on the model defined at (3.1) using 60 days of trade data by GARCH (1, 1). The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the addition event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the addition event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the addition events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the addition events. Differences are calculated by absolute value of model parameters between after and before the addition events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 1,082) are shown in bold.

	Before Addition			After Addition			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	149	869	64	146	870	66	0.970	40	999	43
α_1	1032	50	0	1031	51	0	0.995	188	716	178
α_2	994	86	2	982	97	3	0.627	162	795	125
α_3	849	222	11	811	253	18	0.101	107	868	107
α_4	666	389	27	601	446	35	0.016	85	912	85
α_5	468	573	41	463	578	41	0.976	63	943	76
β_0	4	59	1019	2	65	1015	0.617	344	318	420
β_1	3	107	972	5	136	941	0.107	205	550	327
β_2	10	265	807	1	334	747	0.000	106	784	192
β_3	10	474	598	9	516	557	0.193	58	899	125
β_4	16	654	412	14	705	363	0.076	42	956	84
β_5	22	772	288	20	794	268	0.570	31	986	65
γ_0	225	745	112	207	758	117	0.615	58	952	72
γ_1	267	674	141	281	692	109	0.096	98	914	70
γ_2	187	817	78	169	853	60	0.133	57	993	32
γ_3	88	916	78	113	908	61	0.073	34	1018	30
γ_4	65	958	59	75	950	57	0.676	29	1028	25
γ_5	52	948	82	61	968	53	0.028	28	1032	22
δ_1	7	64	1011	8	95	979	0.036	147	810	125
δ_2	11	179	892	18	242	822	0.001	90	905	87
δ_3	13	338	731	29	388	665	0.002	87	919	76
δ_4	15	467	600	26	477	579	0.180	58	951	73
δ_5	19	529	534	29	558	495	0.114	70	939	73
A	1038	42	2	1040	41	1	0.841	210	687	185
B	1	39	1042	0	59	1023	0.072	290	385	407
Γ	189	767	126	217	774	91	0.022	77	955	50
Δ	7	39	1036	11	44	1027	0.541	182	762	138

**Table 3.9: Frequency of Changes in
Model Parameters for Deletion Events (GARCH)**

Based on the model defined at (3.1) using 60 days of trade data by GARCH (1, 1). The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the deletion event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the deletion event. The section headed “Chsq” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the deletion events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the deletion events. Differences are calculated by absolute value of model parameters between after and before the deletion events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 690) are shown in bold.

	Before Deletion			After Deletion			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	63	573	54	61	572	57	0.944	25	649	16
α_1	653	37	0	641	49	0	0.409	85	484	121
α_2	597	90	3	581	107	2	0.390	76	539	75
α_3	488	196	6	460	220	10	0.201	61	571	58
α_4	323	349	18	339	335	16	0.673	35	619	36
α_5	238	427	25	240	428	22	0.904	36	613	41
β_0	2	41	647	1	54	635	0.329	275	254	161
β_1	5	161	524	5	188	497	0.246	130	491	69
β_2	7	354	329	9	351	330	0.876	65	589	36
β_3	11	471	208	8	459	223	0.563	32	634	24
β_4	16	571	103	17	562	111	0.818	23	652	15
β_5	22	616	52	23	602	65	0.443	18	657	15
γ_0	118	514	58	87	543	60	0.063	34	625	31
γ_1	186	467	37	204	464	22	0.098	33	580	77
γ_2	95	575	20	123	549	18	0.116	10	642	38
γ_3	65	603	22	80	588	22	0.419	15	640	35
γ_4	47	618	25	47	622	21	0.835	11	658	21
γ_5	36	626	28	43	622	25	0.669	11	658	21
δ_1	11	130	549	5	141	544	0.257	66	556	68
δ_2	16	265	409	13	269	408	0.843	27	618	45
δ_3	13	412	265	13	376	301	0.140	23	632	35
δ_4	23	445	222	10	450	230	0.071	30	626	34
δ_5	22	476	192	16	475	199	0.585	28	633	29
A	642	46	2	620	69	1	0.070	94	485	111
B	1	54	635	1	77	612	0.107	208	353	129
Γ	121	547	22	155	516	19	0.070	19	610	61
Δ	7	76	607	2	91	597	0.122	69	528	93

**Table 3.10: Frequency of Changes in
Model Parameters for Addition Events (BEKK)**

Based on the model defined at (3.2) using 60 days of trade data by BEKK (1, 1). The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the addition event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the addition event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the addition events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the addition events. Differences are calculated by absolute value of model parameters between after and before the addition events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 1,082) are shown in bold.

	Before Addition			After Addition			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	133	887	62	140	873	69	0.717	29	999	54
α_1	1044	36	2	1030	49	3	0.319	195	716	171
α_2	940	137	5	914	163	5	0.270	151	834	97
α_3	752	321	9	700	365	17	0.028	100	903	79
α_4	554	518	10	487	576	19	0.006	81	938	63
α_5	381	681	20	324	716	42	0.001	61	972	49
β_0	16	48	1018	13	57	1012	0.577	305	374	403
β_1	168	155	759	149	194	739	0.056	271	594	217
β_2	172	330	580	138	390	554	0.009	179	775	128
β_3	171	520	391	157	551	374	0.392	154	830	98
β_4	181	670	231	158	688	236	0.396	142	886	54
β_5	184	731	167	166	747	169	0.574	140	892	50
γ_0	212	695	175	213	701	168	0.918	77	918	87
γ_1	311	625	146	321	652	109	0.047	108	890	84
γ_2	223	787	72	221	790	71	0.989	70	954	58
γ_3	139	872	71	165	850	67	0.270	50	991	41
γ_4	114	897	71	115	893	74	0.963	55	993	34
γ_5	69	930	83	82	933	67	0.243	44	997	41
δ_1	45	84	953	38	94	950	0.561	216	735	131
δ_2	37	208	837	47	256	779	0.016	140	852	90
δ_3	54	372	656	53	414	615	0.167	139	862	81
δ_4	55	493	534	54	532	496	0.235	123	886	73
δ_5	61	559	462	59	605	418	0.132	103	901	78
A	1027	54	1	1005	73	4	0.087	213	732	137
B	13	42	1027	11	38	1033	0.825	265	423	394
Γ	255	696	131	264	718	100	0.097	99	912	71
Δ	29	30	1023	25	47	1010	0.127	227	684	171

**Table 3.11: Frequency of Changes in
Model Parameters for Deletion Events (BEKK)**

Based on the model defined at (3.2) using 60 days of trade data by BEKK (1, 1). The first panel shows results for individual parameters. The second panel shows results for the sums of the α s, β s, γ s and δ s respectively. The first vertical section show the numbers of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the deletion event. The significance level is defined at the 1%. The second vertical section shows the corresponding results after the deletion event. The section headed “*Chsq*” shows the value of a Chi-squared statistic that is used to test the difference between the observed frequencies in each row. The last section shows the results for the differences in parameters during the deletion events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the deletion events. Differences are calculated by absolute value of model parameters between after and before the deletion events. Cells in which there is a substantial number of statistically significant changes (more than 10% of the sample size, 690) are shown in bold.

	Before Deletion			After Deletion			Chsq 2 DoF	Difference (After minus Before)		
	-ive	ntrl	+ve	-ive	ntrl	+ve		sig. dec.	no sig. change	sig. inc.
α_0	75	571	44	66	587	37	0.496	25	654	11
α_1	644	43	3	643	46	1	0.576	88	507	95
α_2	544	139	7	526	162	2	0.089	70	563	57
α_3	395	288	7	398	288	4	0.660	56	593	41
α_4	240	438	12	277	404	9	0.108	36	622	32
α_5	170	506	14	192	489	9	0.257	37	630	23
β_0	3	40	647	10	42	638	0.144	262	266	162
β_1	83	213	394	90	210	390	0.850	135	504	51
β_2	80	381	229	95	368	227	0.468	93	573	24
β_3	88	470	132	93	468	129	0.915	85	590	15
β_4	97	539	54	101	523	66	0.467	77	600	13
β_5	94	561	35	104	560	26	0.400	77	603	10
γ_0	156	453	81	119	503	68	0.013	40	605	45
γ_1	235	424	31	248	418	24	0.526	39	589	62
γ_2	137	529	24	159	511	20	0.315	24	627	39
γ_3	102	570	18	113	563	14	0.575	24	633	33
γ_4	95	579	16	71	597	22	0.096	24	648	18
γ_5	75	585	30	70	597	23	0.544	28	642	20
δ_1	33	106	551	41	119	530	0.364	81	536	73
δ_2	37	262	391	43	267	380	0.721	56	585	49
δ_3	44	379	267	47	373	270	0.921	50	606	34
δ_4	48	446	196	46	446	198	0.974	60	604	26
δ_5	52	485	153	55	482	153	0.954	50	616	24
A	616	72	2	600	86	4	0.347	87	515	88
B	3	50	637	8	60	622	0.186	207	349	134
Γ	190	472	28	191	471	28	0.998	39	597	54
Δ	29	65	596	31	71	588	0.825	88	507	95

Table 3.12: Comparison on Frequency of Changes in Model Parameters for Addition Events

The table reports a comparison of the summary results for the sums of the α s, β s, γ s and δ s under GARCH (1, 1) and BEKK (1, 1) with the corresponding tables under OLS and VAR for the addition events. The first vertical section show the percentage of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the addition events. The significance level is defined at the 1%. The second vertical section shows the corresponding results after addition event. The third section shows the results for the differences in parameters during the addition events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the addition events. Differences are calculated by absolute value of model parameters between after and before the addition events. The last column shows the size of data sets used under each model.

	Addition (%)									Sample Size
	Before			After			Difference (After minus Before)			
	-ive	ntrl	+ve	-ive	ntrl	+ve	sig. dec.	no sig. change	sig. inc.	
OLS										
<i>A</i>	90.77	9.14	0.09	89.26	10.74	0.00	7.90	84.29	7.81	1,127
<i>B</i>	0.27	4.61	95.12	0.09	5.77	94.14	18.19	52.26	29.55	1,127
Γ	16.50	72.40	11.09	16.50	76.40	7.10	5.06	92.10	2.84	1,127
Δ	0.00	6.39	93.61	0.09	7.72	92.19	5.68	90.42	3.90	1,127
VAR										
<i>A</i>	94.85	5.15	0.00	94.06	5.86	0.09	15.44	69.03	15.53	1,127
<i>B</i>	0.09	5.32	94.59	0.00	5.77	94.23	18.46	52.97	28.57	1,127
Γ	15.97	72.76	11.27	15.17	77.82	7.01	6.21	90.68	3.11	1,127
Δ	0.00	3.37	96.63	0.00	4.61	95.39	16.06	73.03	10.91	1,127
GARCH										
<i>A</i>	95.93	3.88	0.18	96.12	3.79	0.09	19.41	63.49	17.10	1,082
<i>B</i>	0.09	3.60	96.30	0.00	5.45	94.55	26.80	35.58	37.62	1,082
Γ	17.47	70.89	11.65	20.06	71.53	8.41	7.12	88.26	4.62	1,082
Δ	0.65	3.60	95.75	1.02	4.07	94.92	16.82	70.43	12.75	1,082
BEKK										
<i>A</i>	94.92	4.99	0.09	92.88	6.75	0.37	19.69	67.65	12.66	1,082
<i>B</i>	1.20	3.88	94.92	1.02	3.51	95.47	24.49	39.09	36.41	1,082
Γ	23.57	64.33	12.11	24.40	66.36	9.24	9.15	84.29	6.56	1,082
Δ	2.68	2.77	94.55	2.31	4.34	93.35	20.98	63.22	15.80	1,082

Table 3.13: Comparison on Frequency of Changes in Model Parameters for Deletion Events

The table reports a comparison of the summary results for the sums of the α s, β s, γ s and δ s under GARCH (1, 1) and BEKK (1, 1) with the corresponding tables under OLS and VAR for the deletion events. The first vertical section show the percentage of stocks for which the corresponding parameter is significantly less than zero, neutral or significantly greater than zero before the deletion events. The significance level is defined at the 1%. The second vertical section shows the corresponding results after deletion event. The third section shows the results for the differences in parameters during the deletion events and it shows the numbers of stocks for which the corresponding parameter exhibits significant decrease, neutral change or significant increase in its magnitude during the deletion events. Differences are calculated by absolute value of model parameters between after and before the deletion events. The last column shows the size of data sets used under each model.

	Deletion (%)									Sample Size
	Before			After			Difference (After minus Before)			
	-ive	ntrl	+ve	-ive	ntrl	+ve	sig. dec.	no sig. change	sig. inc.	
OLS										
<i>A</i>	81.59	18.41	0.00	83.38	16.62	0.00	3.57	88.46	7.97	728
<i>B</i>	0.00	8.79	91.21	0.00	15.25	84.75	21.70	65.80	12.50	728
Γ	17.86	79.53	2.61	19.37	78.57	2.06	2.20	93.54	4.26	728
Δ	0.00	15.80	84.20	0.14	17.72	82.14	2.47	93.13	4.40	728
VAR										
<i>A</i>	90.38	9.48	0.14	90.38	9.62	0.00	10.30	72.12	17.58	728
<i>B</i>	0.00	8.93	91.07	0.00	13.19	86.81	21.84	65.52	12.64	728
Γ	15.93	81.04	3.02	17.17	80.36	2.47	2.20	93.82	3.98	728
Δ	0.00	9.07	90.93	0.00	11.26	88.74	9.07	78.85	12.09	728
GARCH										
<i>A</i>	93.04	6.67	0.29	89.86	10.00	0.14	13.62	70.29	16.09	690
<i>B</i>	0.14	7.83	92.03	0.14	11.16	88.70	30.14	51.16	18.70	690
Γ	17.54	79.28	3.19	22.46	74.78	2.75	2.75	88.41	8.84	690
Δ	1.01	11.01	87.97	0.29	13.19	86.52	10.00	76.52	13.48	690
BEKK										
<i>A</i>	89.28	10.43	0.29	86.96	12.46	0.58	12.61	74.64	12.75	690
<i>B</i>	0.43	7.25	92.32	1.16	8.70	90.14	30.00	50.58	19.42	690
Γ	27.54	68.41	4.06	27.68	68.26	4.06	5.65	86.52	7.83	690
Δ	4.20	9.42	86.38	4.49	10.29	85.22	12.75	73.48	13.77	690

Table 3.14: Parameter Dynamics (GARCH)

Based on the model defined at (3.1) using 60 days of trade data by GARCH (1, 1). Each panel shows the results of estimates of the sums of the α s, β s, γ s and δ s respectively. Rows of each panel are before the event and columns are after it. “N” and “P” denotes negative and positive respectively. All other changes at a significance level of which is greater than 5% are classified as neutral. Bold entries in the table are percentages which are substantial. More detailed tables which cover individual parameters are available on request.

	Addition					Deletion				
	N (1%)	N (5%)	Neutral	P (5%)	P (1%)	N (1%)	N (5%)	Neutral	P (5%)	P (1%)
<i>A</i>										
N (1%)	92.98	1.02	1.85	0.00	0.09	85.80	2.90	4.20	0.00	0.14
N (5%)	0.74	0.09	0.37	0.00	0.00	0.72	0.00	0.58	0.00	0.00
Neutral	2.22	0.09	0.28	0.09	0.00	3.04	0.87	1.45	0.00	0.00
P (5%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
P (1%)	0.18	0.00	0.00	0.00	0.00	0.29	0.00	0.00	0.00	0.00
<i>B</i>										
N (1%)	0.00	0.00	0.09	0.00	0.00	0.00	0.00	0.00	0.14	0.00
N (5%)	0.00	0.00	0.00	0.00	0.09	0.00	0.00	0.29	0.00	0.00
Neutral	0.00	0.00	1.20	0.28	1.11	0.14	0.00	3.19	0.14	2.61
P (5%)	0.00	0.00	0.09	0.09	0.74	0.00	0.00	0.29	0.29	0.87
P (1%)	0.00	0.00	2.59	1.11	92.61	0.00	0.00	5.07	1.74	85.22
Γ										
N (1%)	8.69	1.29	6.56	0.37	0.55	7.54	1.45	7.83	0.14	0.58
N (5%)	1.94	0.74	3.51	0.00	0.37	3.33	0.58	5.36	0.00	0.29
Neutral	8.13	3.51	44.64	1.85	3.14	10.87	4.93	47.10	2.03	1.59
P (5%)	0.18	0.37	1.94	0.18	0.37	0.43	0.00	2.17	0.29	0.29
P (1%)	1.11	0.55	5.45	0.55	3.97	0.29	0.14	2.46	0.29	0.00
Δ										
N (1%)	0.00	0.00	0.09	0.00	0.55	0.00	0.00	0.29	0.14	0.58
N (5%)	0.00	0.00	0.00	0.00	0.18	0.00	0.00	0.00	0.00	0.00
Neutral	0.18	0.00	0.18	0.28	1.48	0.00	0.00	1.88	0.43	4.64
P (5%)	0.18	0.00	0.00	0.00	1.11	0.00	0.00	0.87	0.72	2.46
P (1%)	0.65	0.00	2.59	0.92	91.59	0.29	0.00	6.23	2.61	78.84

Table 3.15: Parameter Dynamics (BEKK)

Based on the model defined at (3.2) using 60 days of trade data by BEKK (1, 1). Each panel shows the results of estimates of the sums of the α_s , β_s , γ_s and δ_s respectively. Rows of each panel are before the event and columns are after it. “N” and “P” denotes negative and positive respectively. All other changes at a significance level of which is greater than 5% are classified as neutral. Bold entries in the table are percentages which are substantial. More detailed tables which cover individual parameters are available on request.

	Addition					Deletion				
	N (1%)	N (5%)	Neutral	P (5%)	P (1%)	N (1%)	N (5%)	Neutral	P (5%)	P (1%)
<i>A</i>										
N (1%)	89.28	1.48	3.70	0.09	0.37	80.72	3.62	4.35	0.00	0.58
N (5%)	1.20	0.00	0.37	0.00	0.00	1.45	0.29	0.58	0.00	0.00
Neutral	2.31	0.37	0.74	0.00	0.00	4.35	0.87	2.75	0.00	0.00
P (5%)	0.00	0.00	0.00	0.00	0.00	0.14	0.00	0.00	0.00	0.00
P (1%)	0.09	0.00	0.00	0.00	0.00	0.29	0.00	0.00	0.00	0.00
<i>B</i>										
N (1%)	0.28	0.00	0.00	0.00	0.92	0.00	0.00	0.00	0.00	0.43
N (5%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
Neutral	0.00	0.00	0.28	0.37	1.94	0.00	0.00	1.45	0.58	2.46
P (5%)	0.09	0.00	0.00	0.00	1.20	0.14	0.14	0.58	0.29	1.59
P (1%)	0.65	0.00	1.39	1.48	91.40	1.01	0.00	3.77	1.88	85.65
<i>Γ</i>										
N (1%)	11.74	2.31	7.67	0.37	1.48	12.03	3.33	11.16	0.29	0.72
N (5%)	1.48	1.11	4.44	0.28	0.09	2.61	0.29	4.20	0.29	0.43
Neutral	9.15	4.07	33.09	1.57	3.97	11.88	5.51	36.96	1.74	2.32
P (5%)	0.55	0.37	3.42	0.18	0.55	0.72	0.14	1.01	0.14	0.14
P (1%)	1.48	0.28	6.56	0.65	3.14	0.43	0.14	2.75	0.29	0.43
<i>Δ</i>										
N (1%)	1.29	0.00	0.09	0.00	1.29	2.32	0.00	0.14	0.14	1.59
N (5%)	0.00	0.00	0.09	0.00	0.09	0.00	0.00	0.00	0.00	0.14
Neutral	0.00	0.00	0.18	0.18	1.76	0.29	0.14	1.59	0.14	4.20
P (5%)	0.00	0.00	0.00	0.00	0.46	0.00	0.00	0.43	0.14	2.32
P (1%)	1.02	0.09	2.50	1.20	89.74	1.88	0.14	4.78	2.61	76.96

Table 3.16: Summary of Parameter Dynamics

The table reports a summary of key cells of parameter dynamics for the sums of the α s, β s, γ s and δ s under OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) for addition and deletion events respectively. The table entries are shown as percentages of stocks with consistency in the sign of model parameters before and after both events. “N” and “P” denotes negative and positive respectively. All other changes at a significance level of which is greater than 5% are classified as neutral. More detailed tables which cover individual parameters are available on request.

Parameter	Cell	OLS	WLS	VAR	WVAR	GARCH	BEKK
Addition							
<i>A</i>	N (1%) - N (1%)	83.85	59.54	91.13	71.78	92.98	89.28
<i>B</i>	P (1%) - P (1%)	92.10	81.63	91.57	81.46	92.61	91.40
Γ	Neutral - Neutral	45.87	44.54	48.09	37.53	44.64	33.09
Δ	P (1%) - P (1%)	87.49	77.99	93.17	87.58	91.59	89.74
Deletion							
<i>A</i>	N (1%) - N (1%)	70.88	37.23	83.52	53.02	85.80	80.72
<i>B</i>	P (1%) - P (1%)	81.04	60.85	82.28	62.23	85.22	85.65
Γ	Neutral - Neutral	47.94	55.91	53.16	52.61	47.10	36.96
Δ	P (1%) - P (1%)	72.80	56.87	81.59	75.14	78.84	76.96

Table 3.17: Z Scores for Tests of Difference in Model Parameters (Addition)

The table is based on 60-day trade data under OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) models for the addition events. Differences are calculated by absolute value of model parameters between after and before the event (After minus Before). Z scores for difference in individual parameters and sums are computed as described in Section 2.3.3. Bold format denotes significance at least at the 1% level.

	Addition					
	OLS	WLS	VAR	WVAR	GARCH	BEKK
α_0	-0.113	2.202	1.014	-0.609	0.172	1.090
α_1	-0.678	1.246	-0.006	0.001	-0.339	-0.889
α_2	-2.404	-1.220	-0.008	0.000	-1.438	-2.965
α_3	-1.311	-3.348	-0.020	-0.006	0.002	-1.661
α_4	0.147	-0.306	-0.015	-0.003	0.234	-3.735
α_5	1.249	2.060	0.004	-0.001	4.631	-3.496
β_0	2.488	3.219	1.947	2.492	-0.194	1.654
β_1	-3.281	0.844	0.334	-0.547	-1.815	-0.634
β_2	0.695	-2.560	2.743	1.578	-0.063	2.572
β_3	0.813	1.263	1.805	1.364	0.272	2.748
β_4	-0.192	1.750	-0.199	-0.410	-0.246	-0.089
β_5	-1.720	0.556	0.107	1.035	-1.002	-0.327
γ_0	-1.241	-0.510	-1.392	-0.413	0.373	-0.104
γ_1	6.854	-5.951	8.273	-10.457	9.896	7.799
γ_2	3.212	-3.422	3.549	-4.582	6.554	5.347
γ_3	3.793	-3.480	4.136	-6.859	6.517	3.086
γ_4	2.665	-2.313	2.469	-3.550	5.154	2.900
γ_5	-1.021	-2.702	-1.619	-4.685	1.566	4.208
δ_1	-0.766	-2.409	-0.138	-5.163	-1.348	-4.006
δ_2	-3.541	-2.058	-3.292	-4.937	-1.653	-6.007
δ_3	-2.535	-1.030	-4.872	-3.010	-3.443	-5.167
δ_4	0.605	-1.640	-0.752	-2.919	0.105	-0.664
δ_5	-0.600	-0.940	-1.291	-1.638	-1.338	-1.903
A	-1.195	-0.620	-2.177	-3.211	0.491	-4.552
B	1.386	3.487	2.132	2.753	-0.229	1.488
Γ	7.068	-7.355	7.788	-12.257	10.999	8.928
Δ	-3.484	-4.489	-5.772	-9.804	-4.701	-9.625

Table 3.18: Z Scores for Tests of Difference in Model Parameters (Deletion)

The table is based on 60-day trade data under OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) models for the deletion events. Differences are calculated by absolute value of model parameters between after and before the event (After minus Before). Z scores for difference in individual parameters and sums are computed as described in Section 2.3.3. Bold format denotes significance at least at the 1% level.

	Deletion					
	OLS	WLS	VAR	WVAR	GARCH	BEKK
α_0	1.325	-0.756	-1.025	-0.297	1.014	0.930
α_1	4.866	0.839	0.005	0.000	0.731	3.325
α_2	2.374	0.174	0.004	0.000	0.455	0.580
α_3	2.088	1.132	0.000	0.000	0.559	0.821
α_4	3.810	0.205	0.005	0.000	0.259	2.223
α_5	2.616	-0.906	0.002	0.000	1.047	0.478
β_0	1.335	2.380	0.460	1.787	1.976	0.423
β_1	0.660	1.882	2.395	1.715	-0.365	2.427
β_2	1.531	-0.595	0.691	0.254	0.012	0.519
β_3	0.668	0.852	0.039	-0.390	2.288	0.055
β_4	1.277	1.009	0.261	0.025	1.425	0.851
β_5	0.363	1.686	0.234	0.197	0.159	0.221
γ_0	-2.276	2.681	-1.739	2.867	-2.711	-0.079
γ_1	5.395	-3.174	5.470	-5.177	6.227	2.075
γ_2	4.251	-2.488	4.093	-3.316	2.449	3.200
γ_3	4.302	-0.934	3.719	-1.233	1.846	10.165
γ_4	2.584	-1.520	2.319	-1.603	1.951	0.713
γ_5	1.552	0.275	1.299	-0.134	3.028	1.260
δ_1	0.931	-0.045	0.380	-0.176	1.406	-0.548
δ_2	0.367	1.275	-0.081	2.976	0.917	-0.499
δ_3	1.780	-0.461	2.702	0.039	0.898	1.433
δ_4	-1.267	-0.302	-1.039	0.438	0.834	-0.995
δ_5	-0.536	-1.509	1.084	-1.202	-0.068	1.271
A	4.923	0.617	5.394	3.347	1.483	2.780
B	1.001	2.355	1.105	1.471	0.857	0.565
Γ	6.188	-3.374	5.806	-4.377	5.111	6.218
Δ	0.754	-0.354	2.565	1.032	2.259	0.211

Appendix E

Other Tables

Table E.1: Model Validation

The table shows the number of occasions for which the model consistency test based on equation (2.2) yields results at the specified level of significance under GARCH (1, 1) and BEKK (1, 1) models. The Z scores shown in the last column are computed for all sample stocks using the Scott-Smith test.

		$\leq 0.5\%$	$\leq 1\%$	$\leq 2.5\%$	$\leq 5\%$	$\leq 10\%$	$\geq 90\%$	$\geq 95\%$	$\geq 97.5\%$	$\geq 99\%$	$\geq 99.5\%$	Z	
GARCH (1, 1)	Addition												
		<i>Before</i>	22	23	29	40	62	21	16	13	8	6	-0.002
		<i>After</i>	14	17	27	35	57	18	13	11	8	8	-0.173
	Deletion	<i>Before</i>	11	13	14	15	21	19	16	11	8	7	-1.102
	<i>After</i>	9	10	13	14	16	11	9	6	6	5	-0.048	
BEKK (1, 1)	Addition												
		<i>Before</i>	27	31	40	43	57	36	30	23	16	15	-0.764
		<i>After</i>	29	31	43	53	71	53	42	33	28	25	0.062
	Deletion	<i>Before</i>	7	10	13	18	27	24	18	11	9	7	-0.085
	<i>After</i>	10	11	13	18	23	18	12	9	9	8	0.593	

Table E.2: P-Values for Tests of Model Parameters (GARCH)

Based on the model defined at (3.1) using 60 days of trade data by GARCH (1, 1). The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters and probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. Bold format denotes significance at the 1% level.

	Before		Addition		Difference		Before		Deletion		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
α_0	-0.003	0.807	0.006	0.700	0.003	0.863	-0.002	0.907	0.129	0.294	0.126	0.311
α_1	-0.208	0.000	-0.208	0.000	-0.001	0.734	-0.255	0.000	-0.264	0.000	0.010	0.464
α_2	-0.128	0.000	-0.125	0.000	-0.003	0.150	-0.149	0.000	-0.153	0.000	0.004	0.649
α_3	-0.071	0.000	-0.071	0.000	0.000	0.998	-0.088	0.000	-0.091	0.000	0.003	0.576
α_4	-0.045	0.000	-0.045	0.000	0.001	0.815	-0.051	0.000	-0.053	0.000	0.002	0.795
α_5	-0.024	0.000	-0.030	0.000	0.006	0.000	-0.028	0.000	-0.031	0.000	0.003	0.295
β_0	0.194	0.000	0.182	0.003	-0.012	0.845	0.304	0.000	0.379	0.000	0.075	0.048
β_1	0.174	0.000	0.075	0.145	-0.099	0.069	0.104	0.000	0.092	0.004	-0.012	0.715
β_2	0.026	0.062	0.023	0.711	-0.004	0.949	0.060	0.000	0.061	0.110	0.000	0.990
β_3	0.002	0.915	0.020	0.760	0.018	0.785	0.026	0.004	0.107	0.002	0.081	0.022
β_4	0.038	0.066	0.024	0.646	-0.014	0.806	0.000	0.982	-0.067	0.143	0.066	0.154
β_5	-0.060	0.014	0.001	0.982	-0.058	0.316	0.007	0.602	-0.012	0.680	0.005	0.873
γ_0	-0.054	0.000	-0.062	0.001	0.008	0.709	-0.276	0.000	-0.059	0.375	-0.217	0.007
γ_1	-0.051	0.000	-0.092	0.000	0.041	0.000	-0.075	0.000	-0.123	0.000	0.048	0.000
γ_2	-0.040	0.000	-0.069	0.000	0.028	0.000	-0.055	0.000	-0.077	0.000	0.023	0.014
γ_3	-0.016	0.000	-0.046	0.000	0.030	0.000	-0.035	0.000	-0.051	0.000	0.016	0.065
γ_4	-0.006	0.017	-0.030	0.000	0.024	0.000	-0.024	0.000	-0.041	0.000	0.017	0.051
γ_5	0.002	0.540	-0.008	0.017	0.007	0.117	-0.009	0.039	-0.030	0.000	0.021	0.002
δ_1	0.157	0.000	0.154	0.000	-0.002	0.177	0.162	0.000	0.168	0.000	0.005	0.160
δ_2	0.087	0.000	0.084	0.000	-0.003	0.098	0.081	0.000	0.084	0.000	0.003	0.359
δ_3	0.059	0.000	0.053	0.000	-0.005	0.001	0.053	0.000	0.056	0.000	0.003	0.369
δ_4	0.044	0.000	0.044	0.000	0.000	0.916	0.037	0.000	0.039	0.000	0.003	0.404
δ_5	0.038	0.000	0.036	0.000	-0.002	0.181	0.035	0.000	0.035	0.000	0.000	0.945
A	-0.476	0.000	-0.478	0.000	0.003	0.623	-0.571	0.000	-0.592	0.000	0.021	0.138
B	0.375	0.000	0.325	0.137	-0.050	0.819	0.500	0.000	0.559	0.000	0.059	0.391
Γ	-0.111	0.000	-0.245	0.000	0.134	0.000	-0.198	0.000	-0.322	0.000	0.123	0.000
Δ	0.384	0.000	0.371	0.000	-0.012	0.000	0.368	0.000	0.382	0.000	0.013	0.024

Table E.3: P-Values for Tests of Model Parameters (BEKK)

Based on the model defined at (3.2) using 60 days of trade data by BEKK (1, 1). The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters and probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. Bold format denotes significance at the 1% level.

	Before		Addition				Before		Deletion		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
α_0	-0.010	0.059	-0.019	0.002	0.009	0.276	-0.003	0.907	-0.036	0.210	0.033	0.352
α_1	-0.193	0.000	-0.192	0.000	-0.001	0.374	-0.228	0.000	-0.235	0.000	0.007	0.001
α_2	-0.114	0.000	-0.111	0.000	-0.003	0.003	-0.128	0.000	-0.129	0.000	0.001	0.562
α_3	-0.066	0.000	-0.065	0.000	-0.002	0.097	-0.072	0.000	-0.074	0.000	0.001	0.412
α_4	-0.041	0.000	-0.038	0.000	-0.004	0.000	-0.041	0.000	-0.045	0.000	0.004	0.026
α_5	-0.025	0.000	-0.022	0.000	-0.003	0.000	-0.025	0.000	-0.026	0.000	0.001	0.632
β_0	0.203	0.000	0.225	0.000	0.022	0.098	0.384	0.000	0.423	0.000	0.039	0.672
β_1	0.084	0.000	0.078	0.000	-0.006	0.526	0.101	0.000	0.226	0.000	0.126	0.015
β_2	0.038	0.000	0.059	0.000	0.021	0.010	0.057	0.000	0.079	0.059	0.022	0.604
β_3	0.005	0.448	0.034	0.000	0.029	0.006	0.014	0.480	-0.017	0.685	0.003	0.956
β_4	0.006	0.253	0.005	0.577	-0.001	0.928	0.011	0.259	-0.047	0.249	0.035	0.394
β_5	0.006	0.343	0.002	0.845	-0.004	0.744	-0.008	0.317	0.016	0.668	0.009	0.824
γ_0	0.006	0.603	-0.004	0.796	-0.002	0.916	-0.144	0.000	-0.139	0.000	-0.004	0.937
γ_1	-0.056	0.000	-0.087	0.000	0.030	0.000	-0.074	0.000	-0.084	0.000	0.010	0.038
γ_2	-0.060	0.000	-0.082	0.000	0.022	0.000	-0.055	0.000	-0.071	0.000	0.016	0.001
γ_3	-0.040	0.000	-0.054	0.000	0.014	0.002	-0.027	0.000	-0.076	0.000	0.049	0.000
γ_4	-0.019	0.000	-0.030	0.000	0.011	0.004	-0.015	0.000	-0.019	0.000	0.003	0.475
γ_5	-0.006	0.016	-0.021	0.000	0.015	0.000	-0.006	0.032	-0.012	0.001	0.006	0.208
δ_1	0.155	0.000	0.150	0.000	-0.005	0.000	0.157	0.000	0.156	0.000	-0.001	0.583
δ_2	0.085	0.000	0.078	0.000	-0.006	0.000	0.080	0.000	0.079	0.000	-0.001	0.617
δ_3	0.058	0.000	0.053	0.000	-0.005	0.000	0.049	0.000	0.052	0.000	0.002	0.152
δ_4	0.041	0.000	0.040	0.000	-0.001	0.507	0.035	0.000	0.034	0.000	-0.002	0.319
δ_5	0.035	0.000	0.033	0.000	-0.002	0.057	0.028	0.000	0.030	0.000	0.002	0.204
A	-0.440	0.000	-0.428	0.000	-0.013	0.000	-0.494	0.000	-0.508	0.000	0.014	0.005
B	0.341	0.000	0.403	0.000	0.062	0.137	0.559	0.000	0.682	0.002	0.123	0.572
Γ	-0.182	0.000	-0.273	0.000	0.091	0.000	-0.177	0.000	-0.262	0.000	0.084	0.000
Δ	0.374	0.000	0.355	0.000	-0.018	0.000	0.350	0.000	0.350	0.000	0.001	0.833

Chapter 4

The Effect of Market Returns, Crises, Market Capitalisation, Trading Volume and Short Sales.

4.1 Introduction

The impact of short sales on the speed of price adjustment is crucial to price efficiency in the capital market, given that an efficient price discovery process is defined as the one when stock prices respond immediately to the arrival of new information (Fama, 1991). Diamond & Verrecchia (1987) predict that short sales constraints impede the speed of price adjustment, especially for negative information by preventing short selling from investors with negative information about fundamentals. This asymmetric speed of price adjustment caused by short sale constraints can become worse in bear than in bull markets. However, regulators believe that short sale constraints or prohibition in bear markets can reduce the speed of price adjustment to bad news which reflects a negative bubble or herding behaviour rather than fundamental information (Beber & Pagano, 2013).

Most empirical studies on the speed of price adjustment based on daily or weekly data suggest that short sales improve informational efficiency of prices and that short sales constraints or bans reduce market efficiency and impede the process of price discovery (Beber & Pagano, 2013; Bris *et al.*, 2007; Saffi & Sigurdsson, 2010). Based on trade-by-trade high frequency data, researchers including Boehmer & Wu (2013) and Chen & Rhee (2010) use the VAR model to characterise more precisely how price efficiency is affected by short sales and document that short selling is an important driver of price discovery.

In line with previous work, Chapter 2 significantly extends the VAR model of Hasbrouck (1991) and investigates the impact of short sales on price efficiency based on higher-frequency data over a 10-year period in the Hong Kong stock market. The Hong Kong market provides a rare opportunity for investigation on the impact of short sales as it only allows stocks which meet certain requirements to be available for short selling and the D-list for short sales is revised on a quarterly basis. In Chapter 2, the results suggest that a number of the stocks in the data set experience significant decreases in the autocorrelations in quote returns and trades implying that price efficiency is improved by short sales, which is consistent with findings from Boehmer & Wu (2013). These results are robust under the OLS, WLS, VAR and WVAR models. To address the issue of heteroscedasticity caused by high frequency data, Chapter 3 investigates further the impact of short sales on price discovery with consideration of dynamic volatility and co-volatility with application of GARCH and BEKK models. The results of Chapter 3 provide similar evidence supporting that short selling is important to the process of price discovery by capturing the characteristics of heteroscedasticity in the high frequency time series.

This chapter extends the empirical analyses by investigating how the speed of price adjustment is affected by market conditions, market capitalisation, trading volume and short sellers activities. This chapter contributes to the literature in several ways. First, it introduces an extension to the VAR model including a dummy variable of current or one-day lagged market return to explore the impact of market returns on the speed of price adjustment. Moreover, it also includes interaction

variables in the VAR model using the GARCH framework to deal with the issue of heterogeneity of variance. Secondly, the estimations are conducted in crisis and non-crisis periods in order to examine how speed of price adjustment reacts to the extreme market condition. Thirdly, analyses are conducted based on transaction days with positive market returns and those with non-positive market returns. This method provides evidence on the difference in speed of price adjustment under different market conditions. Fourthly, this chapter provides a direct comparison of the effect of short sales on price efficiency by dividing the sample stocks into four quartiles based on firm size, trading volume and measures of short sales.

The findings of this chapter are summarised as follows. First, the three different approaches to market conditions show that short sales improve the process of price discovery in both good and bad market conditions. Stocks react more quickly to new information in bad markets. By comparing the speed of price adjustment during the recent global financial crisis with non-crisis period, the results show that short sales accelerate informational efficiency in the non-crisis period. However mixed changes in parameters during the crisis period indicate that the effect of short sales is unclear in extreme market conditions.

Secondly, large-sized firms are found to adjust more quickly to new information than small-sized firms and the results are robust under all models. It is also noted that medium-sized addition stocks and small-sized deletion stocks are the most affected by the changing eligibility for short sales. Thirdly, the improvement in price efficiency is more significant when stocks have greater trading volume as they adjust their prices to new information more quickly. Fourthly, the speed of price adjustment to new information is faster for stocks with more shorting activity measured by both short interest and shorting flows. It supports the prediction that short sellers are important participants in the process of price discovery.

The remainder of the chapter is organised as follows. Section 4.2 reviews the relevant literature. Section 4.3 presents the model with interaction variables to capture market conditions. Section 4.4 provides some data description. Section 4.5 reports

the empirical findings and Section 4.6 concludes.

4.2 Literature Review

The determinants of the speed of price adjustment have been explored in the literature and are related to variables including market conditions, firm size, trading volume and shorting activities. This section considers the effects predicted by theory and presents a brief account for the evidences so far.

To measure the loss in price efficiency caused by short sales constraints under different market conditions, [Bris *et al.* \(2007\)](#) compare among international markets with different institutional settings on short sales constraints. Based on two price efficiency measurements, the downside-minus-upside market R^2 based on the separate market model estimation conditional on negative and positive market returns and the relative co-movement of an individual stock return with the signed market return lagged one week, they find that prices adjust more quickly to new information especially to bad news in countries where short sales are allowed and practised. Following [Bris *et al.* \(2007\)](#), [Beber & Pagano \(2013\)](#) examine the effect of short sales bans on price efficiency around the world during the 2007-09 crisis. They estimate the autocorrelation between stock returns and market returns and compare them between countries with short sales bans and those without bans during the crisis. A significant higher autocorrelation is observed when stocks are subject to a short selling ban and this finding is consistent with a lower speed of price adjustment with shorting bans during the crisis period, especially for negative market returns. [Chang *et al.* \(2014\)](#) regress weekly stock returns on contemporaneous market returns, conditional on negative and positive market returns. They compare the R^2 of the market model between down and up markets during the implementation of the pilot scheme for short sales in China in 2010-11. They find that there is a significant lower R^2 in both down and up markets with short sales activity. It provides supporting evidences that short sales enhance price efficiency not only in the bad market but also in the good market. By adopting the VAR

model (Hasbrouck, 1991) based on high frequency transaction data, Chen & Rhee (2010) directly measure the speed of price adjustment in both up and down market conditions and compare the speed difference of stocks before and after the effective date of short selling in the Hong Kong stock market. They divide the daily regression results in the study period into up and down market categories based on the sign of market returns. The trading day is defined as an up (down) market day if it has a positive (non-positive) daily open-to-close market return. The results show that there are significant differences in the speed of price adjustment between shortable and non-shortable stocks in both market conditions while the short sales effect is more significant in a bear market.

Since Lo & MacKinlay (1990) find the lead-lag cross-autocorrelations in the returns of larger-sized firms generally leading those of smaller ones, the sources of this lead-lag cross-autocorrelations as a delayed speed of new information has been debated by academics. Lo & MacKinlay (1990) find that the correlation between current small firm's returns and lagged large firm's returns is higher than that between current large firm's returns and lagged small firm's returns. They believe that the lead-lag cross-autocorrelations is caused by the stronger price stickiness or the tendency to adjust more slowly to new information for small firms. Lo & MacKinlay (1990) believe that firm size is a significant determinant of the cross-autocorrelation patterns in stock returns.

By controlling for firm's size effect, Chordia & Swaminathan (2000) examine the cross-autocorrelation patterns between high volume and low volume stocks. They use the ratio of the number of shares traded in a day to the number of shares outstanding to measure of the trading volume. This specific definition of trading volume can offset the size effect on the raw trading volume as they find a high correlation between firm size and raw trading volume. By conducting the VAR model involving pairs of high and low volume portfolios, they find that daily and weekly returns with high volume lead those with low volume even in the largest size quartile, indicating that trading volume plays an independent role in explaining the cross-autocorrelation patterns and is not driven by firm size. Chordia & Swaminathan (2000) suggest that infrequently traded stocks respond more slowly

to information in market returns than actively traded stocks.

Nevertheless, the impact of firm size and trading volume on the speed of price adjustment found in recent empirical studies is controversial. Consistent with [Lo & MacKinlay \(1990\)](#), [Saffi & Sigurdsson \(2010\)](#) find that firms with larger market capitalisation are associated with less price delay based on a research of 26 countries. However, [Saffi & Sigurdsson \(2010\)](#) do not find a significant impact of turnover on price efficiency. [Chen & Rhee \(2010\)](#) report that the speed of price adjustment for smaller firms is even faster than the larger firms in the Hong Kong stock market. Their finding on trading volume is consistent with [Chordia & Swaminathan \(2000\)](#) that stocks traded more frequently react more quickly than those traded less frequently.

In theoretical models of [Diamond & Verrecchia \(1987\)](#), short sellers are assumed to be rational informed traders. They point out that short sellers improve price efficiency by pulling overpriced stocks back to their fundamental levels. They suggest that informed traders are less likely to short for liquidity reasons as short sales are costly, and therefore high short interest from informed traders conveys adverse (negative) information about the stock fundamentals. Most empirical studies show that short sellers are well informed. By using monthly short interest data, [Desai *et al.* \(2002\)](#) reveal that short sellers have private information and their shorting targets are the overpriced firms. Their findings show that heavily shorted firms experience significantly negative abnormal returns and the negative abnormal returns are increasing in the level of short interest. By using shorting flow data, [Boehmer *et al.* \(2008\)](#) find that stocks, which are heavily shorted by institutions, underperform significantly compared to stocks that are lightly shorted by institutions. Following [Boehmer *et al.* \(2008\)](#), [Boehmer & Wu \(2013\)](#) investigate the impact of short sellers' daily activities on price discovery. They believe that short sellers are important economic drivers of price discovery as they are informed traders with private information. The more active short sellers are, the closer stock prices are to their fundamental levels and the faster stock prices incorporate the public information. Using the daily shorting flow standardised by the stock's daily share trading volume, [Boehmer & Wu \(2013\)](#) find that more shorting flow only not makes

stock prices more accurate but also accelerates the speed of price adjustment to the arrival of new information by reducing price delays.

4.3 Model for the Effect of Market Conditions

One of the aims of this chapter is to examine the effect of market returns on the process of price discovery when stocks become shortable/non-shortable. [Chen & Rhee \(2010\)](#) compare the speed of price adjustment between up and down markets based on their analysis of intraday data on a single day using the VAR model. As [Chen & Rhee \(2010\)](#) estimate the VAR model on a daily basis, it is straightforward to divide the estimation results into two groups for comparison: those with positive daily open-to-close market returns and those with daily open-to-close negative market returns. As described in Chapter 2, a close examination of the data used in the study reveals that during the 60-day period before and after addition and deletion events over a 10-year period, there are 17.9% and 39.0% of the total transaction days with less than 20 trades respectively. Estimating a small number of transactions on a daily basis could cause multi-collinearity and other statistical problems. The model estimations in this study therefore are based on the 60 days before and after addition and deletion events as a whole period. To explore this issue, interaction variables with the dummy variable of market conditions and the corresponding parameters are introduced in the model as follows

$$\begin{aligned} R_t &= \alpha_0 + \sum_{i=1}^5 \alpha_i R_{t-i} + \sum_{i=0}^5 \beta_i X_{t-i} + \beta_{m,R} D_{m,t} + \sum_{i=1}^5 \alpha_{i,D} (R_{t-i} \times D_{m,t}) + \sum_{i=0}^5 \beta_{i,D} (X_{t-i} \times D_{m,t}) + \varepsilon_{Rt}, \\ X_t &= \gamma_0 + \sum_{i=1}^5 \gamma_i R_{t-i} + \sum_{i=1}^5 \delta_i X_{t-i} + \beta_{m,X} D_{m,t} + \sum_{i=1}^5 \gamma_{i,D} (R_{t-i} \times D_{m,t}) + \sum_{i=1}^5 \delta_{i,D} (X_{t-i} \times D_{m,t}) + \varepsilon_{Xt}, \end{aligned} \quad (4.1)$$

where $D_{m,t}$ is a dummy variable denoted as 1 if the market return on that day is positive or 0 if it is non-positive. The rest of the variables in the model at (4.1) above keep the same definition as in Chapters 2 & 3.

Interpretations of the coefficients on interaction variables are as follows. The positive (negative) $\sum_{i=1}^5 \alpha_{i,D}$ indicates that the magnitude of quote autocorrelations becomes smaller (greater) with positive (negative) market returns as the quote autocorrelations are initially negative. The positive (negative) $\sum_{i=0}^5 \beta_{i,D}$ indicates that price impact of trades is enhanced (reduced) with positive (negative) market returns as the price impact of trades are initially positive. The impact of market returns on Granger causality depends on the sign of $\sum_{i=1}^5 \gamma_{i,D}$ as well as the initial sign of Granger causality. The positive (negative) $\sum_{i=1}^5 \delta_{i,D}$ indicates that trade continuity is increased (decreased) with positive (negative) market returns as the trade continuity is initially positive. As for the measurements of speed of price adjustment, the negative (positive) $\sum_{i=1}^5 \alpha_{i,D}$ and the positive (negative) $\sum_{i=1}^5 \delta_{i,D}$ indicate that the speed of price adjustment is slower when market condition is good (bad).

The dummy variable of market returns used in the model at (4.1), however, embodies foresight. That is, a quote-based return measured at an arbitrary time of the day would be associated with the dummy variable computing using the price at the end of the day. This means that the model can in principle offer an explanation of price discovery under different market conditions, but cannot be used for forecasting purpose. To overcome the foresight issue, an extension of the model is to use market return on the previous day. These results are also included in this chapter.

The VAR model with interaction variables defined at (4.1) is estimated by OLS and WLS to compare the speed difference in price adjustment under different market conditions before and after the event. Furthermore, this chapter also employs the GARCH models of Chapter 3, which are defined as follows

$$\begin{aligned}
R_t &= \alpha_0 + \sum_{i=1}^5 \alpha_i R_{t-i} + \sum_{i=0}^5 \beta_i X_{t-i} + \beta_{m,R} D_{m,t} + \sum_{i=1}^5 \alpha_{i,D} (R_{t-i} \times D_{m,t}) + \sum_{i=0}^5 \beta_{i,D} (X_{t-i} \times D_{m,t}) + \varepsilon_{Rt}, \\
\sigma_{R,t}^2 &= \omega_R + \varphi_R \varepsilon_{R,t-1}^2 + \phi_R \sigma_{R,t-1}^2, \\
X_t &= \gamma_0 + \sum_{i=1}^5 \gamma_i R_{t-i} + \sum_{i=1}^5 \delta_i X_{t-i} + \beta_{m,X} D_{m,t} + \sum_{i=1}^5 \gamma_{i,D} (R_{t-i} \times D_{m,t}) + \sum_{i=1}^5 \delta_{i,D} (X_{t-i} \times D_{m,t}) + \varepsilon_{Xt}, \\
\sigma_{X,t}^2 &= \omega_X + \varphi_X \varepsilon_{X,t-1}^2 + \phi_X \sigma_{X,t-1}^2.
\end{aligned} \tag{4.2}$$

Diagnostic tests for ARCH effects described in Chapter 3 will be applied based on the OLS and WLS models with interaction variables in this chapter.

The VAR and WVAR models of Chapter 2 and the BEKK model of Chapter 3 are not used in this chapter in conjunction with the model at (4.2). This is because incorporation of the dummy variables would require the development of purpose built software which is beyond the scope of the thesis. Given the general similarity of results from the OLS, WLS and GARCH models, this is not considered to be a major limitation.

4.4 Data Description

This chapter utilises the same high frequency quotes and trades dataset of Chapters 2 & 3 in a 10-year period from May 2001 and May 2011 sourced from the HKEx.

To examine how speed of price adjustment affected under different market conditions when stocks become shortable or non-shortable, three different approaches are applied in this chapter. First, the models with interaction variables are used to investigate the sensitivity of speed of price adjustment under the current market condition and the one-day lagged market condition. Daily market return is calculated by the Hang Seng daily market index sourced from Datastream. Secondly, the estimation results based on Chapters 2 & 3 before and after the event are divided into two categories by crisis and non-crisis periods to provide alternative analysis on how speed of price adjustment reacts to the extreme market condition. By using the recession period from the National Bureau of Economic Research [NBER

henceforth], the crisis period is defined as December 2007 and June 2009¹⁴ which refers to the recent global financial crisis¹⁵. The event is defined in the crisis period if the whole estimation period of the event (60 days before and after the effective date) is in the crisis period. Thirdly, the 60-day data before and after each event for each stock is divided into two groups, transactions on the days with positive and non-positive market returns respectively. For each stock, six models including OLS, WLS, VAR, WVAR, GARCH and BEKK are estimated under up and down markets before and after each event. The first quote return and trade variables on each day is treated as a missing value since the time duration between trades on non-continuous days would be large.

To examine the impact of market capitalisation, trading volume, activities of short sellers including short interest and shorting flow on the speed of price adjustment during addition and deletion events, information on daily market capitalisation, daily trading volume and total shares outstanding is obtained from Datastream. Daily short sales data is sourced from the HKEx.

The analysis based on the current and one-day lagged market return in this chapter contains three models including OLS, WLS, and GARCH. The subgroup analysis based on crisis period, up and down markets, size, trading volume and shorting activities in this chapter contains six models including OLS, WLS, VAR, WVAR, GARCH and BEKK. Time-weighted variables provide alternative solution to heteroscedasticity, and therefore GARCH framework is not implemented with the weighted variables to avoid that the variance dynamics is over specified.

¹⁴NBER recession periods are available on the website of <http://www.nber.org/cycles.html>.

¹⁵The paper by [Boehmer & Wu \(2013\)](#) does not specify the crisis period explicitly and so no direct comparison may be made. The models have been re-estimated using [Beber & Pagano \(2013\)](#) who use the period January 2008 to June 2009 as the crisis period. Not surprisingly the results are the same since the crisis periods differ by one month but the short selling lists are revised quarterly.

4.5 Results

This section presents the results on changes in speed of price adjustment during both addition and deletion events by market conditions, market capitalisation, trading volume and short sales activities. The SS test used in this chapter is defined the same as in Section 2.3.3 and the significance level is at 1%. Section 4.5.1 reports the results by using three approaches dealing with good and bad market conditions. It also reports the diagnostic tests for the ARCH effects introduced in Section 3.2.3 for univariate models interacted with dummy variable for current and one-day lagged market return (OLS and WLS), respectively. Sections 4.5.2, 4.5.3 and 4.5.4 report analyses for the speed of price adjustment by size, trading volume and shorting activities, respectively.

4.5.1 Market Conditions

To investigate the process of price discovery under different market conditions during addition and deletion events, three approaches are used in this section. Section 4.5.1.1 reports the results based on OLS and WLS models interacted with a dummy variable of different market returns to explore the reaction of speed of price adjustment to current and one-day lagged positive and non-positive market returns. Section 4.5.1.2 presents the analysis based on the estimation results from Chapters 2 & 3 broken down by crisis and non-crisis periods. Section 4.5.1.3 provides the estimation results based on the transaction days divided into two groups, days with positive market returns and days with non-positive market returns.

4.5.1.1 Dummy Variable of Current and One-day Lagged Market Return

This section first presents diagnostics tests for ARCH effects for the model at (4.2). It then presents results for current and one-day lagged market return, respectively.

Diagnostic Tests for ARCH effects

For addition events, Table 4.1 reports the diagnostic tests results based on OLS interacted with a dummy variable of current market condition, showing the percentage of event stocks at the corresponding significance level. For intraday data, the significant levels are 0.1%, 1% and 5%. Any results with a significance level greater than 5% are classified as neutral. The *JB* test is used to test for normality of residuals. The Portmanteau *Q* test (*Q* test) and the Lagrange Multiplier Test (*LM* test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for *R* and *X* equations in the model before and after the event. Bold format in the table denotes the greatest percentage value.

[Insert Table 4.1 about here]

Table 4.1 shows that the residuals of nearly all of the sample stocks do not follow normality, which represents a white noise process at a significance level of 1%. At least half of the sample stocks experience ARCH effects at a significance level of 1% as shown by *Q* test and *LM* test. Furthermore, there are more stocks with significant ARCH effects in the *R* equation before and after the addition events compared with the *X* equation. The results in Table 4.2 report the similar results for the deletion events. Although the percentage of event stocks with significant ARCH effects reduces in the *X* equation compared to the addition events, there are still more than approximately 42% of the sample stocks in the *X* equation and 79% in the *R* equation present existence of the ARCH effect at the significance level of 1%.

[Insert Table 4.2 about here]

Tables 4.3 & 4.4 present the corresponding diagnostic test results based on the WLS models. It is noted that there are less stocks experiencing the significant ARCH effects in both *R* and *X* equations during both events when variables in

the models are time-weighted, compared with those under OLS models. There are approximately less than 50% (40%) of event stocks with significant ARCH effect at the level of 1% for addition (deletion) events when time duration is taken into account.

[Insert Tables 4.3 and 4.4 about here]

Furthermore, Tables 4.5, 4.6, 4.7 and 4.8 report similar results for diagnostic tests for ARCH effects for OLS and WLS models interacted with a dummy variable of one-day lagged market return.

[Insert Tables 4.5, 4.6, 4.7 and 4.8 about here]

Based on the results shown above, GARCH framework is applicable to the OLS model interacted with the dummy variable of current market return and one-day lagged market return.

Dummy Variable for Current Market Return

Tables 4.9 shows the p-value for tests of model parameters by OLS interacted with a dummy variable of current market return for addition and deletion events using the model defined at (4.1). The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters and their probabilities, respectively. Differences are calculated by absolute value of model parameters between after and before both events. Bold format in the table denotes significance at the 1% level.

In Table 4.9, there are significantly positive relationships between the dummy variable of current market return and quote returns as well as trades before and after

addition and deletion events. It suggests the average transaction trend of the sample stocks and the current market are moving in the same direction. No significant changes in the dummy variables, $D_{m,R}$ and $D_{m,X}$, during both events indicate that the co-movement between individual stocks and the market is not affected by short sales. Parameters measuring price efficiency interacted with the dummy variable of the current market return stay insignificant before and after both events and this implies that there are no significant differences in the speed of price adjustment between good and bad current market conditions.

[Insert Table 4.9 about here]

Tables 4.10 shows the p-value for tests of model parameters for the WLS model interacted with dummy variable of current market return for addition and deletion events. Table 4.10 reports that the relationship with current market return remains significantly positive for trades before and after both events while the relation will disappear for quote returns when stocks become non-shortable. For the sums of parameters, only the aggregated price impact of trades, B_{Dm} , are found significantly negative before the deletion events. It suggests that the price impact of trades for shortable stocks is greater when the current market condition is bad. However, it is found that the speed differences of price adjustment are not significant under different current market conditions under the time-weighted model.

[Insert Table 4.10 about here]

Table 4.11 presents the corresponding results for the GARCH model. Similar to those under OLS and WLS, the findings do not provide strong evidence to support that current market conditions affect the aggregated parameters, as most of the estimated parameters of the proxies for price efficiency do not change significantly.

[Insert Table 4.11 about here]

In sum, the application of interaction variables with a dummy variable of current market return in the models do not provide strong evidence supporting that the speed of price adjustment is significantly impacted by current market conditions.

Dummy Variable for One-day Lagged Market Return

As discussed in Section 4.3, a dummy variable for current market return used in the models at (4.1) and (4.2) raises foresight issues. Quote returns and trades at an arbitrary time of the day would be associated with the dummy variable computed by the price at the end of the day. To resolve this issue, this section considers the model with interaction variables based on the sign of one-day lagged market return.

Table 4.12 reports the p-value for tests of model parameters by OLS interacted with a dummy variable of one-day lagged market return for both events. D_{mlag} is a dummy variable set to one if the one-day lagged market return is positive and zero otherwise. It is found that the aggregated parameter of interaction variable with trade autocorrelations (Δ_{Dmlag}) before the addition events is 0.009, which is significant at the 1% level. This positively significant parameter Δ_{Dmlag} implies that trade continuity is stronger when the one-day lagged market return is positive and it indicates that stocks have quicker reaction to new information when the market on the previous day was bad.

[Insert Table 4.12 about here]

Table 4.13 shows the corresponding results for WLS. The results show that the aggregated parameter Δ_{Dmlag} after the addition events is 0.013 significant at the 1% level. This positive value implies that the stickiness of trades are stronger when the market return on the previous day is good. The findings support that stocks

have shorter price delay if they are under a bad previous-day market return.

[Insert Table 4.13 about here]

Table 4.14 present the results for the GARCH model. The insignificant aggregated parameters interacted with the dummy variable, A_{Dmlag} and Δ_{Dmlag} , show that the effect of one-day lagged market conditions on the speed of price adjustment is not visible during both events when heteroscedasticity is taken into account.

[Insert Table 4.14 about here]

The overall results based on the dummy variable of current and one-day lagged market return suggest that the speed of price adjustment is not impacted by the current market conditions while it is affected by the one-day lagged market conditions mainly with attribution to the changes in trade continuity. Stocks have stronger trade continuity if the market on the previous day is good which reveals that stocks react more quickly to the news under the bad market.

4.5.1.2 Crisis Period

This section carries on an investigation of the speed of price adjustment in the crisis period. The estimation results from Chapters 2 & 3 are divided into two categories: crisis and non-crisis periods. According to the records from NBER, the crisis period is defined as the period between December 2007 and June 2009.

Table 4.15 shows the tests for sums A and Δ measuring speed of price adjustment for addition stocks under all estimation models in crisis and non-crisis periods. The cell entries of the first panel show the number of observations in each period. The cell entries of other panels show the estimated parameters and their probabilities under each model, respectively. Differences are calculated by absolute value

of model parameters between the corresponding periods. For addition stocks, in the non-crisis period, parameters measuring price efficiency experience a significant decrease implying short sales accelerate the speed of price adjustment. However, results in the crisis period are mixed as trade continuity decreases for all estimation models while autocorrelation in quote returns increases under time-weighted models. Comparing the differences in parameters between crisis and non-crisis periods, the mixed results do not result in a clear picture showing that the speed of price adjustment is significantly different in the crisis period.

[Insert Table 4.15 about here]

Table 4.16 reports the results for deletion stocks. For both periods, the increases in the sums A and Δ reveal that stocks have longer price delay when they are removed from the D-list. Similar to those for the addition events, the differences in the speed of price adjustment between crisis and non-crisis period for deletion stocks stay inconclusive.

[Insert Table 4.16 about here]

To conclude, short sales recover the process of price discovery in the non-crisis period. However the results are mixed for the crisis period as the efficiency-enhancing effect of short sales is not clear under the time-weighted models. In addition, no clear evidence is found to support that there are significant differences in the speed of price adjustment between the crisis and the non-crisis periods. Table F.1 in Appendix F reports the cumulative abnormal returns based on the market model estimated before, during and after the crisis. The period of crisis is as defined above (December 2007 to June 2009). For the market model, an estimation window of $(-240, -60)$, with a minimum length of 180 days, is used. Each panel of the table indicates that the minimum length of time for stocks to be added back after deletion from the D-list.

As shown in Table F.1, the daily estimated cumulative abnormal returns are greater than zero at the 1% level of significance during the crisis period for the event windows of 30 and 60 days. If the stocks had been deleted from the D list following the standard procedures, it is to be expected that the abnormal returns would be zero or negative. The positive cumulative daily abnormal returns in the crisis panel indicate that stocks were not deleted from the D-list during the crisis period because of poor performance. It is therefore considered likely that deletion served as a substitute for a ban on short sales in the Hong Kong stock market. This finding is supported qualitatively by the small number of stocks added to the D-list during the crisis period compared to the number of additions both before and afterwards. The lack of clear effects being associated with the crisis period for addition stocks reported in Table 4.15 may be caused by the small number of stock/events imposed by the market regulator.

4.5.1.3 Up and Down Markets

Based on the method described in Section 4.4, this section provides an alternative examination on the process of price discovery under different market conditions. The 60-day trade-by-trade data before and after each event for each stock is divided into two groups by the sign of the market returns on that day including positive (up) and non-positive (down) returns. For each stock, the six models used in this thesis are estimated based on these two groups, namely up and down markets respectively. Table 4.17 reports that there are significant decreases in the autocorrelations of quote returns and trades indicating that price efficiency is enhanced by short sales in both up and down markets for the addition stocks. Moreover, the differences between up and down markets show that stocks respond more quickly in the down market rather than in the up market. For the deletion events, the similar results are observed in Table 4.18. Price efficiency becomes worse when stocks are removed from the D-list regardless of market conditions. However, no significant differences in the speed of price adjustment are found under the majority of estimation models between the up and the down markets for the deletion stocks.

[Insert Tables 4.17 and 4.18 about here]

To conclude, a faster process of price discovery is associated with short sales no matter whether market return is up or down. In general, stocks adjust their prices more quickly to new information in down markets rather than in up markets. This finding is consistent with the uncertain information hypothesis of [Brown *et al.* \(1988\)](#) that stock price reactions to bad news (down markets) are stronger than good news (up markets). With a higher aversion to downside risk, investors react faster to bad news. For instance, institutional investors are quicker to respond to unfavourable information as they feel that they would be penalized more if they underperform in a bear market than in a bull market ([Sortino & Van Der Meer, 1991](#)). Furthermore, the cost of not adjusting prices downward is higher than the cost of not adjusting prices upward as the former involves building up inventory with overpriced securities to maintain price continuity.

4.5.2 Market Capitalisation

This section explores the effect of market capitalisation on the speed of price adjustment during both events by dividing the estimation results into four-sized quartiles. For each addition and deletion event, stocks are ranked by their average daily market value calculated in the previous 60 days before the effective date. Group values are assigned to each stock ranging from 0 to 3 for quartiles. For instance, quartiles partition the market values into four groups, with the smallest values receiving, by default, a quartile value of 0 and the largest values receiving a quartile value of 3. The formula for calculating group values is:

$$FLOOR(rank * k / (n + 1)),$$

where *FLOOR* is the function that returns the largest integer that is less than or equal to the argument, fuzzed to avoid unexpected floating point results; *rank*

is the value's order rank; k is the number of groups; n is the number of observations having non-missing values of the ranking variables. $MVRANK = 1$ indicates the smallest market size group while $MVRANK = 4$ indicates the largest market size group. Sums of A and Δ are the average values within each size quartile. Table 4.19 presents the results based on the size quartiles for the addition events. Stocks in all size quartiles experience a significant decrease in the autocorrelation of quote returns and trades which indicates that all the stocks regardless their market capitalisations react more quickly when they are allowed for short sales. The differences between the biggest and the smallest firms after the addition events reveal that large firms are faster in incorporation of new information in the market. It is also found that the medium-sized firms are affected most by short sales during the addition events with robustness under all models.

[Insert Table 4.19 about here]

Table 4.20 shows the results of size quartiles for the deletion events. Stocks for all size quartiles have significant increased autocorrelations in quote returns and trades under a majority of models and it implies that there is a greater price delay after stocks being removed from the D-list for short sales. Similar to the results for the addition events, the significant differences between the largest and the smallest show that large firms still respond more quickly when they are removed from the D-list. The results also indicate that small firms are the most affected with robustness under all models during the deletion events.

[Insert Table 4.20 about here]

To sum up, the overall results show that short sales strengthen price efficiency regardless of firm size. The speed of price adjustment of larger firms is quicker than smaller firms for both events. Medium-sized firms are the most influenced during the addition events while small-sized firms are the most affected during the

deletion events.

4.5.3 Trading Volume

This section investigates the effect of trading volume on the speed of price adjustment during both events by dividing estimation results into four volume quartiles. The trading volume is the ratio of the number of shares traded to the total number of shares outstanding. This measure is used as it offsets the size effect on the raw trading volume, as there is a high correlation between size and raw trading volume (Chordia & Swaminathan, 2000). For each addition and deletion event, stocks are ranked by their average daily trading volume calculated in the previous 60 days before the effective date. $VORANK = 1$ indicates the lowest trading volume group while $VORANK = 4$ indicates the highest trading volume group. Sums of A and Δ are the average values within each volume quartile. In Table 4.21 for addition stocks, results under the majority of models show a significant decrease in the autocorrelation of quote returns and trades although there are a few mixed evidences in changes under GARCH. The first and second largest volume quartiles exhibit significant changes under all models and it reveals that stocks traded more frequently are more likely to be affected by short sales than those traded infrequently during the addition events. The comparison between the most and the least traded groups indicates that, in general, stocks with more trading activities spend less time to adjust their prices under the unweighted models.

[Insert Table 4.21 about here]

Table 4.22 shows the results within four volume quartiles for deletion stocks. The sample stocks within all volume quartiles experience a significant increase in the autocorrelations of the corresponding variables indicating that price efficiency gets worse when stocks become non-shortable. Contrary to those in the addition events, stocks from the top and the bottom of the quartile are affected most during the deletion events. Moreover, it also supports that the speed of price adjustment is

faster when stocks are traded more frequently.

[Insert Table 4.22 about here]

In conclusion, short sales generally accelerate the speed of price adjustment during addition and deletion events. Price efficiency is more likely to be affected for stocks with higher trading volume when they are added to the D-list while stocks with the highest and the lowest trading volume are impacted most during the deletion events. Stocks traded more frequently respond more quickly in price adjustment to new information and this is robust under the models which are not time-weighted.

4.5.4 Shorting Activity

This section uses short interest (Desai *et al.*, 2002) and shorting flow (Boehmer *et al.*, 2008) to measure the effect of short sellers' activities on the speed of price adjustment. Note that shorting activity quartiles are formed based on the daily short sales data which is only available after the addition events and before the deletion events. It means that the analysis of shorting activity approaches may face foresight issues. Therefore, although the results can offer an explanation of price discovery they are not suitable to be used for forecasting purposes.

4.5.4.1 Short Interest

The short interest is the ratio of the number of shares sold short to the total number of shares outstanding on a daily basis. For each addition and deletion event, stocks are ranked by their short interest calculated in the 60-day period after the addition events and before the deletion events. $SIRANK = 1$ indicates the group for stocks with the lowest short interest and $SIRANK = 4$ indicates the group for stocks with the highest short interest. Sums of A and Δ are the average values

within each short interest ratio quartile.

In Table 4.23 for addition stocks, the results show that stocks under all models experience a significant improvement when they become shortable with significant decreases in autocorrelations in quote returns and trades after the addition events. The differences between the most shorted stocks and the least shorted stocks reveal that addition stocks are more efficient in price discovery if they are more frequently traded by short sellers under all unweighted models.

[Insert Table 4.23 about here]

Table 4.24 reports the results for deletion stocks based on the short selling interest before the event. In general, stocks lose their price efficiency shown as significant increases in autocorrelations in quote returns and trades except for the case under the BEKK model. The mixed results on the price comparison between two short interest quartiles 4 & 1 imply that the correlation between the speed of price adjustment and the level of short interest is unclear for deletion stocks.

[Insert Table 4.24 about here]

4.5.4.2 Shorting Flow

The shorting flow is the ratio of the stock's daily shorting volume to its daily trading volume. This ratio makes shorting activity comparable across stocks with different trading volumes (Boehmer *et al.*, 2008). For each addition and deletion event, stocks are ranked by their average daily shorting flow in the following 60 days after the addition event or in the previous 60 days before the deletion event. SFRANK = 1 indicates the lowest shorting flow while SFRANK = 4 indicates the highest shorting flow. Sums of A and Δ are the average values for each shorting flow quartile. In Table 4.25 for addition stocks, similar to the results from short

interest, it shows that short sales enhance the process of price discovery for addition stocks for each shorting flow quartile under all models with significant decreases in autocorrelations. Again, significant decreases in the parameter autocorrelations are observed between stocks with the highest shorting flow and those with the lowest shorting flow and the findings are robust with all models. It suggests that more shorting flows are associated with a better price discovery for addition stocks.

[Insert Table 4.25 about here]

Table 4.26 reports the results for the deletion events. Except for the results based on the VAR model, stocks becoming non-shortable have slower speed of price adjustment shown by some significant increases in autocorrelations after the event. Consistent with the results for the addition stocks, stocks with higher shorting flows exhibit better price efficiency.

[Insert Table 4.26 about here]

To sum up, the analysis of short sellers' activities including short interest and shorting flow suggests that short sellers' trading activities contributes significantly to the process of price discovery and in general stock prices respond more quickly if short sellers are more active in the market.

4.6 Conclusions

This chapter conducts more detailed empirical analyses by examining how the process of price discovery is impacted under different market conditions, market capitalisation, trading volume and short sellers' activities during both addition and deletion events in the Hong Kong stock market for the 10-year study period.

For market conditions, three different approaches are applied for investigation including models interacted with the dummy variable of current and one-day lagged market return, subgroup analysis by dividing the study period into crisis and non-crisis periods, re-estimation of the models in Chapter 2 & 3 based on the days with positive and non-positive market returns, respectively. Taken together, the results suggest that short sales enhance the speed in incorporation of all available information into prices regardless of the market status. Comparing up and down markets, stock prices adjust more quickly to the information in a bearish market. By looking at the recent global financial crisis, the results show that short sales improve price efficiency during the non-crisis period however the effect of short sales remain unclear during the crisis.

For the size effect, there is a positive relationship between firm size and the corresponding speed of information. Larger-sized firms absorb the news faster than smaller-sized firms. The results are consistent with those from [Lo & MacKinlay \(1990\)](#) that small firms have stronger price stickiness or the tendency to adjust more slowly to new information. Medium-sized addition stocks and small-sized deletion stocks are the most affected groups during these events under all estimation models. For the volume effect, stocks with higher trading volume are faster on information assimilation under unweighted models. Based on two different measures as the proxies for short sellers' activity, the findings suggest that short sellers play an important role in the information efficiency of prices. Higher short interest is associated with shorter price delays under the models without consideration of time duration. More shorting flows are found to be a significant determinant of faster speed of price adjustment to new information. Both proxies suggest that prices incorporate information faster when short sellers are more active.

**Table 4.1: Diagnostic Tests for ARCH effects
(OLS - Addition Events - Current Market Return Dummy Variable)**

Based on the model defined at (4.1) using 60 days of trade data by OLS for the addition events, the table represents the diagnostic tests for ARCH effects with the percentage of the event stocks at the corresponding significance level. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The Jarque-Bera test (JB test) is used to test for normality of residuals. The Ljung-Box test (Q test) and the Engles Lagrange Multiplier ARCH test (LM test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for R and X equations in the model before and after the event. Bold format denotes the greatest percentage value.

	Before Addition (R equation)			Before Addition (X equation)			After Addition (R equation)			After Addition (X equation)						
	0.1%	1%	5%	ntrl	0.1%	1%	5%	ntrl	0.1%	1%	5%	ntrl				
JB	99.60%	0.00%	0.00%	0.40%	99.90%	0.00%	0.00%	0.10%	99.20%	0.30%	0.10%	0.40%	99.70%	0.00%	0.00%	0.30%
$Q(1)$	60.40%	6.40%	4.70%	28.60%	36.60%	8.70%	8.80%	45.90%	60.50%	4.80%	6.50%	28.20%	33.00%	9.40%	8.20%	49.40%
$Q(2)$	71.40%	4.50%	4.10%	20.00%	45.30%	9.10%	6.70%	38.80%	71.50%	4.80%	3.20%	20.40%	39.80%	9.70%	8.90%	41.60%
$Q(3)$	75.90%	3.70%	2.10%	18.40%	49.20%	8.40%	6.50%	35.80%	74.90%	4.40%	3.20%	17.50%	42.60%	9.10%	8.00%	40.40%
$Q(4)$	78.60%	2.50%	1.60%	17.30%	51.20%	8.00%	7.90%	32.90%	77.40%	2.90%	4.30%	15.40%	45.90%	8.10%	6.40%	39.60%
$Q(5)$	80.10%	2.00%	2.30%	15.60%	52.00%	7.60%	6.60%	33.70%	78.70%	3.00%	3.40%	14.90%	48.40%	5.70%	6.30%	39.60%
$Q(6)$	81.70%	2.00%	1.70%	14.60%	53.50%	7.40%	6.80%	32.30%	79.60%	3.20%	2.80%	14.40%	48.60%	5.80%	6.90%	38.70%
$Q(7)$	82.50%	1.60%	2.10%	13.80%	54.80%	7.10%	5.90%	32.20%	80.20%	3.20%	3.10%	13.50%	49.90%	5.30%	6.20%	38.70%
$Q(8)$	82.60%	2.20%	2.20%	13.00%	55.70%	6.80%	5.90%	31.60%	81.10%	3.20%	3.00%	12.70%	49.80%	5.70%	5.80%	38.70%
$Q(9)$	83.30%	2.50%	1.40%	12.70%	56.30%	6.60%	6.20%	30.90%	81.70%	3.00%	2.40%	12.80%	50.20%	6.00%	5.60%	38.20%
$Q(10)$	83.70%	2.20%	1.80%	12.30%	57.50%	5.90%	5.20%	31.40%	82.10%	3.00%	2.70%	12.20%	50.70%	6.40%	5.70%	37.20%
$Q(11)$	83.50%	2.70%	1.60%	12.20%	58.00%	5.70%	5.60%	30.70%	82.30%	2.90%	3.00%	11.80%	51.70%	5.70%	5.40%	37.10%
$Q(12)$	83.30%	3.00%	1.30%	12.40%	58.10%	6.10%	5.50%	30.30%	82.70%	2.70%	2.90%	11.70%	52.40%	5.60%	5.30%	36.80%

Table continued on the following page.

Table 4.1: Continued from previous page

	Before Addition (<i>R</i> equation)			Before Addition (<i>X</i> equation)			After Addition (<i>R</i> equation)			After Addition (<i>X</i> equation)						
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%				
<i>LM</i> (1)	60.40%	6.30%	4.70%	28.70%	36.60%	8.70%	8.80%	45.90%	60.50%	4.90%	6.30%	28.30%	33.00%	9.30%	8.10%	49.60%
<i>LM</i> (2)	70.70%	5.10%	4.00%	20.20%	44.80%	9.10%	7.00%	39.10%	70.80%	5.50%	3.20%	20.50%	39.20%	9.50%	9.10%	42.20%
<i>LM</i> (3)	75.30%	3.90%	2.20%	18.50%	48.50%	8.30%	7.30%	35.90%	74.20%	4.70%	3.30%	17.80%	42.20%	8.50%	8.40%	40.90%
<i>LM</i> (4)	77.80%	2.90%	2.00%	17.40%	50.30%	7.80%	8.10%	33.80%	76.60%	3.20%	4.10%	16.10%	45.00%	7.70%	6.90%	40.40%
<i>LM</i> (5)	78.80%	3.00%	2.20%	16.10%	50.60%	8.10%	6.90%	34.40%	77.40%	3.40%	3.40%	15.80%	46.90%	6.40%	6.30%	40.40%
<i>LM</i> (6)	80.50%	2.80%	1.90%	14.80%	51.70%	7.90%	7.40%	33.00%	78.80%	2.90%	3.50%	14.80%	47.40%	6.50%	6.50%	39.60%
<i>LM</i> (7)	81.30%	2.70%	1.90%	14.20%	52.80%	7.50%	6.10%	33.50%	79.30%	3.00%	3.40%	14.30%	48.70%	5.70%	5.60%	40.10%
<i>LM</i> (8)	81.20%	3.20%	2.30%	13.30%	54.10%	7.40%	5.70%	32.70%	80.10%	3.10%	3.10%	13.60%	48.40%	6.10%	5.40%	40.10%
<i>LM</i> (9)	82.00%	2.90%	2.10%	13.10%	54.30%	6.90%	6.70%	32.10%	80.80%	2.50%	3.20%	13.50%	48.60%	6.10%	5.70%	39.60%
<i>LM</i> (10)	81.60%	3.40%	2.20%	12.70%	55.70%	6.30%	5.70%	32.40%	81.20%	2.40%	3.00%	13.40%	49.20%	6.10%	5.60%	39.10%
<i>LM</i> (11)	82.20%	3.00%	2.30%	12.60%	55.80%	6.50%	6.30%	31.40%	81.10%	3.00%	3.00%	13.00%	50.00%	5.20%	5.80%	38.90%
<i>LM</i> (12)	82.20%	2.90%	2.10%	12.80%	56.00%	6.30%	5.70%	32.00%	81.90%	2.50%	2.90%	12.70%	50.70%	5.70%	5.70%	37.80%

**Table 4.2: Diagnostic Tests for ARCH effects
(OLS - Deletion Events - Current Market Return Dummy Variable)**

Based on the model defined at (4.1) using 60 days of trade data by OLS for the deletion events, the table represents the diagnostic tests for ARCH effects with the percentage of the event stocks at the corresponding significance level. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The Jarque-Bera test (*JB* test) is used to test for normality of residuals. The Ljung-Box test (*Q* test) and the Engles Lagrange Multiplier ARCH test (*LM* test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for *R* and *X* equations in the model before and after the event. Bold format denotes the greatest percentage value.

	Before Deletion (<i>R</i> equation)			Before Deletion (<i>X</i> equation)			After Deletion (<i>R</i> equation)			After Deletion (<i>X</i> equation)						
	0.1%	1%	5%	ntrl	0.1%	1%	5%	0.1%	1%	5%	ntrl	0.1%	1%	5%	ntrl	
<i>JB</i>	98.50%	0.30%	0.60%	0.70%	99.90%	0.00%	0.00%	0.10%	96.20%	0.70%	0.70%	2.40%	99.30%	0.40%	0.10%	0.10%
<i>Q</i> (1)	63.00%	6.30%	6.70%	24.00%	26.40%	8.00%	7.40%	58.20%	58.10%	9.10%	6.30%	26.50%	21.90%	8.60%	9.30%	60.30%
<i>Q</i> (2)	72.10%	5.50%	3.60%	18.80%	32.10%	8.10%	7.00%	52.70%	69.10%	6.00%	4.80%	20.10%	29.50%	7.30%	8.30%	55.00%
<i>Q</i> (3)	75.50%	4.50%	3.40%	16.70%	35.10%	7.90%	6.50%	50.60%	72.70%	4.90%	4.30%	18.10%	32.00%	7.40%	7.30%	53.30%
<i>Q</i> (4)	76.60%	4.10%	3.80%	15.60%	37.20%	7.30%	5.50%	50.10%	74.90%	3.90%	3.80%	17.40%	32.70%	8.60%	5.30%	53.40%
<i>Q</i> (5)	77.60%	3.90%	4.20%	14.30%	38.10%	6.70%	6.50%	48.70%	75.30%	4.10%	3.60%	17.00%	33.70%	8.00%	5.50%	52.90%
<i>Q</i> (6)	78.00%	3.60%	3.50%	14.90%	39.40%	7.40%	4.20%	48.90%	76.00%	3.60%	3.80%	16.50%	35.10%	7.00%	6.50%	51.50%
<i>Q</i> (7)	78.30%	3.50%	3.10%	15.10%	41.00%	5.20%	6.70%	47.10%	76.30%	3.40%	3.20%	17.10%	36.30%	6.00%	7.00%	50.60%
<i>Q</i> (8)	78.30%	3.80%	3.10%	14.90%	41.20%	5.30%	6.00%	47.40%	77.10%	2.90%	3.80%	16.10%	37.00%	6.30%	7.00%	49.60%
<i>Q</i> (9)	78.10%	3.50%	3.10%	15.30%	42.20%	5.50%	5.30%	47.00%	77.00%	3.20%	3.20%	16.50%	36.50%	6.70%	8.10%	48.70%
<i>Q</i> (10)	78.70%	3.20%	3.20%	14.90%	42.90%	4.80%	5.30%	47.00%	76.90%	3.40%	2.40%	17.40%	36.50%	7.40%	7.20%	48.90%
<i>Q</i> (11)	78.70%	2.90%	3.10%	15.30%	43.80%	4.50%	5.30%	46.40%	77.70%	2.80%	2.70%	16.80%	37.70%	5.90%	6.00%	50.40%
<i>Q</i> (12)	79.00%	2.50%	3.10%	15.40%	44.30%	4.30%	4.90%	46.40%	77.70%	2.40%	3.20%	16.70%	38.00%	6.20%	5.30%	50.50%

Table continued on the following page.

Table 4.2: Continued from previous page

	Before Deletion (<i>R</i> equation)			Before Deletion (<i>X</i> equation)			After Deletion (<i>R</i> equation)			After Deletion (<i>X</i> equation)						
	0.1%	1%	5%	ntnl	0.1%	1%	5%	ntnl	0.1%	1%	5%	ntnl				
<i>LM</i> (1)	63.00%	6.30%	6.60%	24.10%	26.40%	7.90%	7.60%	58.20%	57.80%	9.30%	6.50%	26.50%	21.90%	8.40%	9.40%	60.30%
<i>LM</i> (2)	71.40%	5.60%	3.60%	19.40%	31.30%	8.60%	7.00%	53.20%	68.20%	6.50%	4.90%	20.50%	28.20%	7.60%	8.70%	55.50%
<i>LM</i> (3)	74.30%	4.80%	3.90%	17.00%	34.20%	7.90%	7.00%	50.90%	71.20%	6.00%	3.50%	19.20%	30.40%	7.90%	7.70%	54.00%
<i>LM</i> (4)	74.80%	4.90%	4.20%	16.10%	35.80%	7.00%	6.30%	50.90%	71.90%	5.90%	3.90%	18.20%	31.10%	8.80%	5.90%	54.10%
<i>LM</i> (5)	75.70%	4.80%	3.80%	15.70%	36.90%	6.50%	6.60%	50.10%	73.10%	5.20%	3.80%	18.00%	31.60%	9.00%	5.60%	53.90%
<i>LM</i> (6)	75.90%	4.90%	3.10%	16.10%	37.90%	6.90%	6.00%	49.20%	73.80%	5.30%	3.60%	17.30%	33.20%	6.60%	7.90%	52.30%
<i>LM</i> (7)	76.30%	3.90%	3.50%	16.30%	39.60%	4.90%	6.70%	48.80%	73.50%	5.50%	3.60%	17.40%	33.70%	6.50%	8.10%	51.80%
<i>LM</i> (8)	76.40%	4.10%	3.40%	16.10%	40.30%	4.30%	6.60%	48.80%	74.20%	4.60%	4.30%	16.80%	34.40%	6.90%	7.30%	51.50%
<i>LM</i> (9)	75.30%	4.50%	4.10%	16.10%	40.70%	5.90%	5.20%	48.20%	74.20%	4.80%	3.10%	18.00%	34.50%	6.90%	7.90%	50.80%
<i>LM</i> (10)	75.60%	4.80%	3.50%	16.10%	40.70%	6.20%	4.80%	48.40%	73.90%	5.80%	3.40%	17.00%	34.60%	7.20%	7.20%	51.10%
<i>LM</i> (11)	75.90%	4.30%	3.40%	16.40%	41.90%	5.00%	4.50%	48.50%	74.10%	5.20%	3.40%	17.40%	35.90%	6.20%	6.30%	51.60%
<i>LM</i> (12)	76.20%	3.90%	3.50%	16.40%	42.10%	5.00%	3.60%	49.20%	73.50%	5.80%	2.90%	17.80%	35.50%	6.00%	6.90%	51.60%

**Table 4.3: Diagnostic Tests for ARCH effects
(WLS - Addition Events - Current Market Return Dummy Variable)**

Based on the model defined at (4.1) using 60 days of trade data by WLS for the addition events, the table represents the diagnostic tests for ARCH effects with the percentage of the event stocks at the corresponding significance level. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The Jarque-Bera test (*JB* test) is used to test for normality of residuals. The Ljung-Box test (*Q* test) and the Engles Lagrange Multiplier ARCH test (*LM* test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for *R* and *X* equations in the model before and after the event. Bold format denotes the greatest percentage value.

	Before Addition (<i>R</i> equation)			Before Addition (<i>X</i> equation)			After Addition (<i>R</i> equation)			After Addition (<i>X</i> equation)		
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%
<i>JB</i>	100.00%	0.00%	0.00%	100.00%	0.00%	0.00%	99.80%	0.00%	0.10%	100.00%	0.00%	0.00%
<i>Q</i> (1)	32.30%	3.30%	3.30%	61.10%	0.00%	9.10%	30.90%	4.00%	4.80%	30.90%	7.00%	7.00%
<i>Q</i> (2)	37.10%	2.50%	2.20%	58.10%	42.20%	7.40%	35.90%	3.50%	3.40%	36.10%	6.60%	6.60%
<i>Q</i> (3)	38.90%	2.80%	2.60%	55.70%	45.40%	6.50%	38.00%	3.80%	4.00%	39.50%	5.80%	6.20%
<i>Q</i> (4)	40.60%	3.10%	3.00%	53.20%	46.50%	5.50%	40.40%	3.60%	2.90%	41.00%	4.90%	5.70%
<i>Q</i> (5)	41.70%	2.80%	3.70%	51.80%	47.50%	4.60%	41.50%	4.40%	3.50%	42.20%	5.40%	4.80%
<i>Q</i> (6)	42.30%	3.00%	3.70%	50.90%	47.90%	5.40%	43.90%	3.10%	4.70%	42.30%	5.60%	5.10%
<i>Q</i> (7)	42.90%	2.90%	3.50%	50.80%	48.80%	4.70%	44.80%	3.20%	4.50%	43.30%	6.00%	3.90%
<i>Q</i> (8)	43.90%	3.00%	3.40%	49.80%	49.30%	4.30%	45.70%	3.60%	4.80%	44.10%	4.80%	4.60%
<i>Q</i> (9)	44.70%	2.70%	3.50%	49.10%	49.10%	3.90%	46.40%	3.10%	4.70%	44.60%	4.40%	4.70%
<i>Q</i> (10)	45.10%	3.60%	2.80%	48.50%	49.90%	4.50%	46.70%	3.40%	4.80%	45.00%	3.70%	4.80%
<i>Q</i> (11)	45.60%	3.70%	3.10%	47.60%	50.50%	4.40%	47.00%	3.90%	4.00%	45.10%	4.10%	5.00%
<i>Q</i> (12)	45.70%	3.70%	2.40%	48.20%	51.10%	4.90%	47.40%	3.50%	3.90%	45.30%	4.00%	4.60%

Table continued on the following page.

Table 4.3: Continued from previous page

	Before Addition (<i>R</i> equation)			Before Addition (<i>X</i> equation)			After Addition (<i>R</i> equation)			After Addition (<i>X</i> equation)						
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%				
<i>LM</i> (1)	32.30%	3.40%	3.20%	61.10%	35.40%	6.70%	9.10%	48.70%	30.90%	4.00%	4.70%	60.40%	30.90%	7.00%	7.00%	55.20%
<i>LM</i> (2)	36.80%	2.90%	2.20%	58.20%	42.00%	5.40%	7.70%	44.90%	35.70%	3.40%	3.70%	57.20%	36.10%	6.00%	7.20%	50.80%
<i>LM</i> (3)	38.60%	2.70%	3.00%	55.70%	44.80%	5.00%	6.50%	43.70%	37.80%	4.00%	3.70%	54.40%	38.80%	6.00%	6.30%	48.90%
<i>LM</i> (4)	40.10%	3.30%	3.10%	53.50%	46.00%	5.50%	5.30%	43.20%	40.00%	3.60%	3.00%	53.50%	40.10%	5.40%	5.60%	49.00%
<i>LM</i> (5)	41.10%	3.20%	3.40%	52.30%	46.60%	5.50%	4.40%	43.50%	41.00%	3.70%	4.80%	50.60%	41.30%	5.70%	5.00%	47.90%
<i>LM</i> (6)	41.50%	3.60%	3.50%	51.40%	47.30%	5.40%	4.70%	42.70%	43.20%	3.10%	4.80%	48.90%	41.80%	5.50%	4.50%	48.30%
<i>LM</i> (7)	42.50%	3.00%	3.40%	51.10%	48.00%	5.00%	5.20%	41.80%	44.00%	3.10%	4.60%	48.30%	42.90%	5.40%	4.10%	47.60%
<i>LM</i> (8)	43.40%	2.80%	3.70%	50.10%	48.30%	5.00%	4.80%	42.00%	44.90%	3.70%	4.70%	46.70%	43.00%	5.30%	4.20%	47.40%
<i>LM</i> (9)	43.90%	3.00%	3.80%	49.30%	47.90%	6.00%	4.10%	42.00%	45.70%	3.10%	4.80%	46.40%	43.60%	4.20%	5.10%	47.10%
<i>LM</i> (10)	44.40%	3.80%	3.10%	48.70%	49.00%	4.90%	4.50%	41.60%	46.00%	3.30%	4.70%	46.00%	44.00%	4.00%	4.80%	47.10%
<i>LM</i> (11)	45.10%	3.50%	2.90%	48.50%	49.40%	4.70%	4.40%	41.50%	46.50%	3.10%	4.20%	46.20%	43.90%	4.60%	4.50%	47.00%
<i>LM</i> (12)	45.30%	3.10%	2.90%	48.70%	49.50%	4.70%	5.30%	40.50%	47.00%	3.00%	3.80%	46.20%	44.50%	4.50%	4.40%	46.60%

**Table 4.4: Diagnostic Tests for ARCH effects
(WLS - Deletion Events - Current Market Return Dummy Variable)**

Based on the model defined at (4.1) using 60 days of trade data by WLS for the deletion events, the table represents the diagnostic tests for ARCH effects with the percentage of the event stocks at the corresponding significance level. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The Jarque-Bera test (JB test) is used to test for normality of residuals. The Ljung-Box test (Q test) and the Engles Lagrange Multiplier ARCH test (LM test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for R and X equations in the model before and after the event. Bold format denotes the greatest percentage value.

	Before Deletion (R equation)			Before Deletion (X equation)			After Deletion (R equation)			After Deletion (X equation)		
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%
JB	100.00%	0.00%	0.00%	100.00%	0.00%	0.00%	99.30%	0.00%	0.40%	100.00%	0.00%	0.00%
$Q(1)$	21.60%	3.50%	5.20%	27.10%	4.90%	7.20%	22.90%	4.80%	5.80%	27.60%	5.60%	7.70%
$Q(2)$	25.50%	5.00%	4.30%	31.60%	5.60%	6.00%	28.50%	3.60%	4.50%	33.10%	4.20%	5.50%
$Q(3)$	29.60%	5.20%	4.10%	35.10%	4.50%	5.00%	30.70%	3.60%	4.90%	34.90%	4.50%	4.30%
$Q(4)$	31.70%	4.60%	4.60%	35.80%	4.10%	4.30%	31.30%	4.10%	4.50%	35.90%	5.00%	2.70%
$Q(5)$	33.40%	4.10%	4.10%	36.00%	4.50%	4.20%	33.10%	2.70%	5.50%	37.20%	4.10%	3.80%
$Q(6)$	33.70%	4.90%	3.90%	36.70%	4.50%	4.50%	34.40%	3.60%	5.20%	38.10%	3.80%	4.10%
$Q(7)$	35.90%	3.90%	4.80%	37.60%	3.90%	3.90%	35.30%	4.50%	4.20%	38.30%	3.90%	4.80%
$Q(8)$	36.00%	4.50%	4.90%	37.30%	3.80%	4.30%	35.30%	5.00%	4.10%	38.10%	4.50%	4.60%
$Q(9)$	36.90%	4.20%	4.60%	36.70%	4.10%	4.60%	36.00%	5.90%	3.40%	39.30%	4.10%	3.80%
$Q(10)$	37.20%	4.50%	4.30%	37.60%	3.90%	4.50%	36.20%	5.50%	3.50%	40.10%	4.10%	3.90%
$Q(11)$	37.90%	4.50%	4.10%	37.70%	4.30%	3.60%	36.60%	4.60%	3.50%	40.70%	3.80%	3.10%
$Q(12)$	38.30%	4.30%	3.90%	38.70%	3.40%	3.80%	36.50%	4.50%	4.20%	40.40%	3.60%	2.90%

Table continued on the following page.

Table 4.4: Continued from previous page

	Before Deletion (<i>R</i> equation)			Before Deletion (<i>X</i> equation)			After Deletion (<i>R</i> equation)			After Deletion (<i>X</i> equation)						
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%				
<i>LM</i> (1)	21.60%	3.50%	5.00%	69.80%	27.10%	4.90%	7.20%	60.90%	22.90%	4.60%	5.90%	66.60%	27.80%	5.50%	7.70%	59.00%
<i>LM</i> (2)	25.10%	5.30%	3.90%	65.60%	31.10%	5.30%	6.60%	56.90%	28.30%	3.60%	4.50%	63.50%	32.80%	4.10%	5.60%	57.50%
<i>LM</i> (3)	29.20%	5.50%	3.90%	61.40%	34.80%	4.20%	5.20%	55.80%	30.60%	3.50%	5.30%	60.60%	34.50%	4.60%	4.30%	56.50%
<i>LM</i> (4)	30.90%	4.90%	4.80%	59.50%	35.10%	3.80%	4.30%	56.80%	31.00%	3.50%	5.30%	60.20%	35.50%	4.60%	3.50%	56.40%
<i>LM</i> (5)	32.40%	4.30%	4.50%	58.80%	35.10%	4.60%	4.30%	56.00%	32.40%	3.20%	5.00%	59.30%	36.50%	4.50%	3.60%	55.40%
<i>LM</i> (6)	32.80%	4.90%	4.50%	57.80%	35.20%	5.20%	4.20%	55.40%	33.40%	4.30%	4.90%	57.40%	36.60%	4.50%	4.50%	54.40%
<i>LM</i> (7)	34.20%	5.20%	4.60%	56.00%	36.00%	4.50%	3.90%	55.50%	34.50%	5.30%	3.80%	56.40%	36.20%	5.60%	4.50%	53.70%
<i>LM</i> (8)	35.30%	4.60%	4.80%	55.30%	36.20%	3.90%	4.20%	55.70%	35.20%	5.20%	3.90%	55.70%	37.00%	5.00%	4.60%	53.30%
<i>LM</i> (9)	36.30%	4.10%	4.60%	55.00%	35.20%	4.60%	4.10%	56.10%	35.50%	5.60%	3.60%	55.30%	37.90%	4.20%	4.60%	53.30%
<i>LM</i> (10)	36.60%	4.20%	4.50%	54.70%	37.00%	3.60%	4.30%	55.00%	35.30%	4.80%	4.60%	55.30%	39.40%	3.80%	3.60%	53.20%
<i>LM</i> (11)	37.30%	3.80%	4.50%	54.40%	37.00%	3.60%	4.10%	55.30%	35.10%	5.00%	4.10%	55.80%	39.40%	3.50%	3.90%	53.20%
<i>LM</i> (12)	37.00%	4.90%	4.20%	53.90%	37.40%	3.50%	3.60%	55.40%	35.20%	4.80%	5.00%	55.00%	38.30%	4.30%	3.90%	53.40%

**Table 4.5: Diagnostic Tests for ARCH effects
(OLS - Addition Events - One-day Lagged Market Return Dummy Variable)**

Based on the model defined at (4.1) using 60 days of trade data by OLS for the addition events, the table represents the diagnostic tests for ARCH effects with the percentage of the event stocks at the corresponding significance level. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The Jarque-Bera test (*JB* test) is used to test for normality of residuals. The Ljung-Box test (*Q* test) and the Engles Lagrange Multiplier ARCH test (*LM* test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for *R* and *X* equations in the model before and after the event. Bold format denotes the greatest percentage value.

	Before Addition (<i>R</i> equation)			Before Addition (<i>X</i> equation)			After Addition (<i>R</i> equation)			After Addition (<i>X</i> equation)			
	0.1%	1%	5%	ntrl	0.1%	1%	5%	ntrl	0.1%	1%	5%	ntrl	
<i>JB</i>	99.50%	0.10%	0.10%	0.40%	99.80%	0.20%	0.00%	0.00%	99.30%	0.30%	0.20%	0.30%	99.70%
<i>Q</i> (1)	60.00%	6.50%	4.60%	28.90%	36.30%	9.00%	9.40%	45.30%	61.20%	4.50%	5.70%	28.60%	32.40%
<i>Q</i> (2)	71.60%	4.70%	3.20%	20.40%	45.70%	8.20%	7.10%	39.00%	71.70%	4.30%	3.50%	20.50%	39.80%
<i>Q</i> (3)	76.20%	3.80%	2.00%	18.10%	49.20%	8.40%	6.10%	36.30%	74.80%	4.00%	3.30%	17.80%	43.20%
<i>Q</i> (4)	78.30%	2.50%	2.10%	17.10%	51.60%	7.30%	6.80%	34.30%	76.90%	3.50%	3.60%	16.00%	46.30%
<i>Q</i> (5)	79.50%	2.20%	2.70%	15.70%	52.20%	7.00%	6.40%	34.40%	78.50%	3.00%	3.20%	15.30%	48.30%
<i>Q</i> (6)	81.50%	1.70%	2.10%	14.80%	53.50%	7.00%	6.50%	33.00%	79.50%	3.20%	2.40%	14.90%	49.50%
<i>Q</i> (7)	82.30%	1.40%	2.00%	14.20%	55.00%	6.30%	6.40%	32.30%	80.10%	3.00%	2.80%	14.20%	50.30%
<i>Q</i> (8)	83.10%	1.30%	2.00%	13.60%	55.80%	6.40%	6.00%	31.80%	81.30%	3.00%	2.40%	13.40%	50.10%
<i>Q</i> (9)	83.30%	2.00%	1.40%	13.30%	56.30%	6.80%	5.70%	31.20%	82.00%	2.30%	2.90%	12.80%	50.60%
<i>Q</i> (10)	83.20%	2.20%	1.90%	12.70%	56.90%	6.50%	5.20%	31.40%	82.30%	2.40%	2.60%	12.60%	50.90%
<i>Q</i> (11)	83.40%	2.20%	1.80%	12.60%	57.00%	6.60%	5.30%	31.10%	82.50%	2.60%	2.60%	12.30%	51.50%
<i>Q</i> (12)	83.50%	2.60%	1.40%	12.50%	57.70%	6.20%	5.60%	30.50%	82.70%	3.00%	2.20%	12.10%	51.90%

Table continued on the following page.

Table 4.5: Continued from previous page

	Before Addition (<i>R</i> equation)				Before Addition (<i>X</i> equation)				After Addition (<i>R</i> equation)				After Addition (<i>X</i> equation)			
	0.1%	1%	5%	ntnl	0.1%	1%	5%	ntnl	0.1%	1%	5%	ntnl	0.1%	1%	5%	ntnl
<i>LM</i> (1)	59.90%	6.50%	4.70%	28.90%	36.30%	9.00%	9.40%	45.30%	61.10%	4.50%	5.90%	28.50%	32.40%	10.10%	8.40%	49.00%
<i>LM</i> (2)	70.80%	5.50%	3.30%	20.40%	45.00%	8.60%	6.80%	39.60%	71.10%	4.50%	3.90%	20.60%	39.20%	9.80%	9.80%	41.30%
<i>LM</i> (3)	75.70%	3.80%	2.30%	18.20%	48.70%	8.30%	6.50%	36.60%	74.60%	3.90%	3.50%	18.00%	42.10%	8.80%	8.80%	40.30%
<i>LM</i> (4)	78.00%	2.20%	2.50%	17.30%	50.10%	7.90%	7.20%	34.90%	76.40%	3.70%	3.30%	16.60%	44.40%	7.90%	7.80%	39.90%
<i>LM</i> (5)	78.90%	2.40%	2.90%	15.80%	50.80%	7.30%	7.30%	34.70%	77.30%	3.50%	3.50%	15.70%	47.20%	6.50%	6.00%	40.20%
<i>LM</i> (6)	80.50%	2.40%	2.20%	14.90%	51.60%	7.50%	7.30%	33.50%	78.60%	2.90%	3.10%	15.40%	47.50%	6.50%	5.90%	40.10%
<i>LM</i> (7)	81.20%	2.40%	2.20%	14.20%	53.30%	7.10%	5.80%	33.80%	79.20%	3.20%	2.90%	14.70%	47.90%	6.80%	4.90%	40.30%
<i>LM</i> (8)	81.50%	2.60%	2.20%	13.70%	54.60%	5.80%	6.80%	32.80%	79.80%	3.30%	3.00%	13.80%	48.60%	6.30%	5.10%	40.10%
<i>LM</i> (9)	82.40%	2.30%	1.70%	13.50%	54.80%	6.60%	6.00%	32.50%	80.30%	3.30%	2.90%	13.50%	48.40%	6.80%	5.00%	39.80%
<i>LM</i> (10)	82.30%	2.50%	1.90%	13.30%	55.40%	6.50%	5.60%	32.60%	80.60%	3.20%	2.60%	13.60%	49.20%	6.80%	5.30%	38.70%
<i>LM</i> (11)	82.50%	2.30%	2.20%	12.90%	55.80%	6.20%	6.20%	31.80%	81.20%	2.90%	2.90%	13.10%	49.60%	6.10%	6.50%	37.80%
<i>LM</i> (12)	82.30%	2.60%	2.00%	13.10%	56.50%	5.90%	5.50%	32.20%	81.20%	3.30%	2.50%	13.00%	49.80%	6.90%	5.60%	37.60%

**Table 4.6: Diagnostic Tests for ARCH effects
(OLS - Deletion Events - One-day Lagged Market Return Dummy Variable)**

Based on the model defined at (4.1) using 60 days of trade data by OLS for the deletion events, the table represents the diagnostic tests for ARCH effects with the percentage of the event stocks at the corresponding significance level. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The Jarque-Bera test (JB test) is used to test for normality of residuals. The Ljung-Box test (Q test) and the Engles Lagrange Multiplier ARCH test (LM test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for R and X equations in the model before and after the event. Bold format denotes the greatest percentage value.

	Before Deletion (R equation)			Before Deletion (X equation)			After Deletion (R equation)			After Deletion (X equation)					
	0.1%	1%	5%	ntrl	0.1%	1%	5%	0.1%	1%	5%	ntrl	0.1%	1%	5%	ntrl
JB	98.90%	0.40%	0.60%	0.10%	99.90%	0.00%	0.00%	0.10%	96.40%	1.00%	0.40%	99.30%	0.30%	0.00%	0.40%
$Q(1)$	62.50%	7.70%	5.60%	24.20%	26.80%	7.40%	9.40%	56.40%	58.70%	9.10%	6.60%	22.30%	8.40%	9.50%	59.80%
$Q(2)$	73.40%	4.20%	4.30%	18.10%	32.80%	8.40%	7.30%	51.50%	69.00%	5.60%	5.00%	20.30%	8.00%	7.60%	55.70%
$Q(3)$	76.10%	3.60%	3.90%	16.40%	36.40%	6.90%	7.30%	49.40%	72.40%	5.50%	3.50%	18.60%	7.30%	7.60%	53.50%
$Q(4)$	77.50%	3.80%	3.20%	15.50%	38.20%	5.50%	7.30%	49.00%	74.50%	5.00%	3.10%	17.40%	7.40%	5.90%	52.90%
$Q(5)$	78.40%	3.40%	3.20%	15.00%	39.10%	6.20%	6.20%	48.60%	74.90%	4.30%	3.90%	16.80%	7.60%	5.70%	52.70%
$Q(6)$	78.30%	3.60%	3.50%	14.60%	40.60%	7.10%	3.50%	48.70%	75.60%	3.90%	4.20%	16.20%	6.90%	6.40%	51.40%
$Q(7)$	78.30%	3.80%	2.90%	15.00%	41.70%	5.50%	5.20%	47.60%	75.60%	3.60%	4.10%	16.70%	6.60%	7.10%	50.40%
$Q(8)$	78.40%	4.20%	3.40%	14.00%	42.40%	5.00%	5.90%	46.60%	75.80%	4.20%	3.90%	16.10%	6.00%	6.70%	50.40%
$Q(9)$	78.40%	3.90%	3.50%	14.10%	43.70%	3.90%	5.90%	46.50%	76.20%	2.90%	4.50%	16.40%	6.60%	6.60%	49.70%
$Q(10)$	78.00%	4.80%	2.80%	14.40%	44.10%	4.30%	5.30%	46.20%	76.50%	2.90%	3.60%	16.90%	6.60%	8.00%	48.90%
$Q(11)$	77.50%	5.30%	2.80%	14.40%	44.00%	5.20%	4.60%	46.20%	76.20%	4.10%	3.60%	16.10%	6.30%	6.60%	50.00%
$Q(12)$	78.30%	4.50%	3.10%	14.10%	44.80%	5.00%	3.50%	46.60%	76.10%	4.60%	2.70%	16.70%	6.20%	6.30%	50.60%

Table continued on the following page.

Table 4.6: Continued from previous page

	Before Deletion (<i>R</i> equation)				Before Deletion (<i>X</i> equation)				After Deletion (<i>R</i> equation)				After Deletion (<i>X</i> equation)			
	0.1%	1%	5%	ntnl	0.1%	1%	5%	ntnl	0.1%	1%	5%	ntnl	0.1%	1%	5%	ntnl
<i>LM</i> (1)	62.30%	7.70%	5.90%	24.10%	26.80%	7.60%	9.20%	56.40%	58.50%	9.10%	6.70%	25.60%	22.10%	8.50%	9.40%	59.90%
<i>LM</i> (2)	72.10%	4.80%	4.80%	18.30%	31.90%	8.70%	7.60%	51.80%	68.10%	6.20%	4.60%	21.10%	28.00%	8.50%	7.00%	56.40%
<i>LM</i> (3)	74.80%	4.20%	3.20%	17.80%	35.40%	7.00%	7.40%	50.10%	70.60%	6.70%	4.20%	18.50%	30.70%	8.30%	6.70%	54.30%
<i>LM</i> (4)	75.80%	4.20%	3.60%	16.40%	36.70%	6.00%	7.70%	49.60%	72.10%	6.20%	3.80%	17.90%	31.50%	8.30%	6.70%	53.50%
<i>LM</i> (5)	76.30%	4.60%	4.10%	15.00%	38.20%	5.90%	6.20%	49.70%	71.70%	6.20%	4.80%	17.40%	31.70%	8.30%	6.40%	53.60%
<i>LM</i> (6)	76.50%	4.30%	3.80%	15.40%	38.90%	7.10%	4.30%	49.60%	71.80%	6.30%	5.00%	16.80%	33.30%	7.10%	7.40%	52.10%
<i>LM</i> (7)	75.90%	5.00%	3.60%	15.40%	39.90%	6.20%	4.80%	49.20%	73.00%	4.90%	4.30%	17.80%	34.20%	7.00%	7.10%	51.70%
<i>LM</i> (8)	75.60%	5.20%	3.80%	15.40%	40.80%	5.00%	6.00%	48.20%	73.40%	4.30%	4.60%	17.60%	34.90%	6.30%	6.70%	52.10%
<i>LM</i> (9)	75.50%	5.50%	3.60%	15.40%	41.30%	4.90%	6.40%	47.30%	73.40%	4.30%	4.50%	17.80%	34.30%	7.00%	6.90%	51.80%
<i>LM</i> (10)	75.40%	5.60%	3.50%	15.50%	41.30%	6.60%	4.90%	47.20%	73.20%	4.20%	4.90%	17.60%	34.70%	7.30%	7.30%	50.70%
<i>LM</i> (11)	75.80%	4.60%	3.80%	15.80%	42.30%	5.50%	4.80%	47.50%	73.40%	4.80%	4.30%	17.50%	35.30%	6.30%	7.10%	51.30%
<i>LM</i> (12)	75.60%	4.30%	4.20%	15.80%	42.30%	5.90%	4.10%	47.80%	73.00%	4.80%	4.50%	17.80%	35.40%	6.20%	5.50%	52.90%

**Table 4.7: Diagnostic Tests for ARCH effects
(WLS - Addition Events - One-day Lagged Market Return Dummy Variable)**

Based on the model defined at (4.1) using 60 days of trade data by WLS for the addition events, the table represents the diagnostic tests for ARCH effects with the percentage of the event stocks at the corresponding significance level. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The Jarque-Bera test (*JB* test) is used to test for normality of residuals. The Ljung-Box test (*Q* test) and the Engles Lagrange Multiplier ARCH test (*LM* test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for *R* and *X* equations in the model before and after the event. Bold format denotes the greatest percentage value.

	Before Addition (<i>R</i> equation)			Before Addition (<i>X</i> equation)			After Addition (<i>R</i> equation)			After Addition (<i>X</i> equation)		
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%
<i>JB</i>	99.90%	0.00%	0.10%	0.00%	0.00%	0.00%	99.90%	0.00%	0.00%	99.90%	0.00%	0.10%
<i>Q</i> (1)	32.50%	3.00%	3.90%	60.60%	6.40%	9.30%	48.80%	31.40%	4.50%	31.50%	5.90%	7.30%
<i>Q</i> (2)	36.60%	2.80%	2.50%	58.20%	5.80%	7.00%	44.90%	35.70%	3.90%	36.60%	6.40%	6.00%
<i>Q</i> (3)	38.40%	2.90%	3.60%	55.10%	5.00%	5.90%	43.40%	38.50%	4.00%	39.10%	7.30%	4.70%
<i>Q</i> (4)	40.30%	3.00%	3.20%	53.40%	5.40%	6.10%	42.30%	40.00%	4.30%	40.60%	5.80%	5.60%
<i>Q</i> (5)	41.40%	3.10%	3.20%	52.20%	5.20%	4.90%	42.70%	41.50%	4.20%	42.20%	5.70%	5.30%
<i>Q</i> (6)	42.20%	3.00%	3.20%	51.50%	4.80%	4.80%	42.00%	43.10%	3.90%	42.30%	5.50%	5.10%
<i>Q</i> (7)	42.70%	3.00%	3.50%	50.70%	4.80%	5.20%	41.00%	44.20%	3.80%	43.70%	5.60%	4.10%
<i>Q</i> (8)	43.70%	3.30%	3.80%	49.20%	5.30%	4.70%	40.80%	45.30%	3.60%	44.50%	4.50%	4.80%
<i>Q</i> (9)	44.40%	3.20%	3.30%	49.00%	5.50%	3.90%	41.20%	46.10%	3.20%	44.90%	4.40%	4.50%
<i>Q</i> (10)	45.10%	3.60%	3.30%	48.00%	5.30%	4.10%	40.60%	46.60%	3.00%	45.50%	3.30%	5.00%
<i>Q</i> (11)	45.30%	3.90%	3.30%	47.60%	4.20%	4.80%	40.20%	47.00%	2.90%	45.40%	3.80%	4.70%
<i>Q</i> (12)	45.50%	3.90%	2.40%	48.10%	4.20%	5.00%	39.70%	47.20%	3.20%	45.90%	4.10%	3.90%

Table continued on the following page.

Table 4.7: Continued from previous page

	Before Addition (<i>R</i> equation)			Before Addition (<i>X</i> equation)			After Addition (<i>R</i> equation)			After Addition (<i>X</i> equation)						
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%				
<i>LM</i> (1)	32.50%	3.00%	3.90%	60.60%	35.50%	6.40%	9.30%	48.80%	31.40%	4.50%	4.40%	59.80%	31.40%	6.00%	7.30%	55.40%
<i>LM</i> (2)	36.00%	3.20%	2.50%	58.20%	41.80%	6.10%	7.20%	44.90%	35.60%	3.70%	4.20%	56.50%	36.20%	6.40%	6.00%	51.40%
<i>LM</i> (3)	38.30%	3.00%	3.40%	55.40%	45.00%	4.80%	6.60%	43.50%	38.10%	4.30%	3.60%	54.00%	38.40%	7.10%	5.10%	49.50%
<i>LM</i> (4)	40.10%	3.00%	3.30%	53.70%	46.00%	5.00%	6.10%	42.90%	39.20%	4.40%	3.50%	53.00%	39.90%	6.00%	5.50%	48.70%
<i>LM</i> (5)	41.00%	3.00%	3.40%	52.60%	46.70%	4.80%	5.50%	43.00%	40.90%	3.90%	4.70%	50.40%	41.60%	5.20%	5.60%	47.60%
<i>LM</i> (6)	41.50%	3.30%	3.40%	51.80%	47.20%	5.70%	4.30%	42.70%	42.90%	3.30%	5.50%	48.30%	41.90%	5.00%	5.30%	47.80%
<i>LM</i> (7)	42.30%	3.00%	3.70%	51.00%	48.00%	5.10%	5.20%	41.70%	43.70%	3.00%	5.30%	48.00%	42.80%	5.30%	4.70%	47.10%
<i>LM</i> (8)	43.60%	2.40%	4.10%	49.80%	48.20%	5.60%	4.30%	41.90%	44.40%	3.80%	5.20%	46.70%	43.20%	4.80%	4.90%	47.00%
<i>LM</i> (9)	43.60%	3.60%	3.70%	49.10%	48.00%	5.80%	4.40%	41.80%	45.50%	3.10%	5.20%	46.10%	43.80%	4.40%	4.70%	47.00%
<i>LM</i> (10)	43.80%	4.40%	2.90%	48.90%	48.30%	5.70%	4.60%	41.40%	45.50%	3.30%	5.50%	45.70%	43.60%	4.70%	4.60%	47.10%
<i>LM</i> (11)	44.60%	3.90%	2.90%	48.60%	49.30%	4.80%	4.90%	40.90%	46.10%	3.40%	5.20%	45.30%	43.90%	4.90%	3.90%	47.30%
<i>LM</i> (12)	44.60%	4.20%	2.40%	48.70%	49.50%	4.80%	5.40%	40.30%	46.20%	3.70%	4.70%	45.30%	45.00%	4.60%	3.50%	47.00%

**Table 4.8: Diagnostic Tests for ARCH effects
(WLS - Deletion Events - One-day Lagged Market Return Dummy Variable)**

Based on the model defined at (4.1) using 60 days of trade data by WLS for the deletion events, the table represents the diagnostic tests for ARCH effects with the percentage of the event stocks at the corresponding significance level. The significance levels are 0.1%, 1% and 5%. Results at a significance level greater than 5% are classified as neutral. The Jarque-Bera test (*JB* test) is used to test for normality of residuals. The Ljung-Box test (*Q* test) and the Engles Lagrange Multiplier ARCH test (*LM* test) are used to test the significance of ARCH effects. Numbers in the brackets denotes the number of lags of the residual series. Four vertical sections are for *R* and *X* equations in the model before and after the event. Bold format denotes the greatest percentage value.

	Before Deletion (<i>R</i> equation)			Before Deletion (<i>X</i> equation)			After Deletion (<i>R</i> equation)			After Deletion (<i>X</i> equation)			
	0.1%	1%	5%	ntrl	0.1%	1%	5%	ntrl	0.1%	1%	5%	ntrl	
<i>JB</i>	99.90%	0.10%	0.00%	0.00%	100.00%	0.00%	0.00%	98.90%	0.10%	0.30%	0.00%	0.70%	0.00%
<i>Q</i> (1)	21.10%	3.90%	5.60%	69.30%	26.50%	5.00%	7.60%	60.90%	22.10%	5.50%	5.50%	66.90%	27.30%
<i>Q</i> (2)	26.30%	3.90%	5.00%	64.70%	31.50%	5.60%	5.50%	57.40%	28.20%	3.60%	3.80%	64.40%	32.80%
<i>Q</i> (3)	30.10%	4.60%	4.10%	61.20%	34.30%	4.80%	4.50%	56.40%	30.50%	3.50%	4.50%	61.50%	34.60%
<i>Q</i> (4)	32.50%	4.50%	4.30%	58.70%	34.90%	4.60%	4.20%	56.30%	30.30%	5.30%	4.20%	60.20%	35.60%
<i>Q</i> (5)	33.80%	4.20%	4.20%	57.80%	36.30%	3.20%	4.10%	56.40%	31.80%	4.90%	3.90%	59.40%	36.60%
<i>Q</i> (6)	34.60%	4.20%	3.60%	57.60%	36.60%	4.20%	4.20%	55.00%	33.50%	5.70%	3.60%	57.10%	37.00%
<i>Q</i> (7)	36.10%	3.60%	4.50%	55.70%	37.30%	3.50%	4.60%	54.60%	34.30%	5.50%	4.80%	55.50%	37.70%
<i>Q</i> (8)	36.80%	3.50%	4.90%	54.80%	37.00%	3.60%	3.80%	55.60%	35.00%	5.20%	4.80%	55.00%	38.10%
<i>Q</i> (9)	37.30%	3.50%	4.80%	54.50%	36.60%	4.80%	3.50%	55.20%	36.10%	4.50%	4.80%	54.60%	38.70%
<i>Q</i> (10)	37.70%	4.10%	3.90%	54.30%	38.10%	4.10%	3.40%	54.50%	35.70%	5.00%	4.20%	55.00%	40.30%
<i>Q</i> (11)	38.20%	3.60%	4.20%	53.90%	37.80%	4.20%	2.90%	55.00%	36.60%	3.90%	4.50%	55.00%	40.10%
<i>Q</i> (12)	38.50%	3.50%	4.90%	53.10%	39.10%	2.80%	3.20%	54.90%	36.80%	3.40%	5.20%	54.60%	40.30%

Table continued on the following page.

Table 4.8: Continued from previous page

	Before Deletion (<i>R</i> equation)			Before Deletion (<i>X</i> equation)			After Deletion (<i>R</i> equation)			After Deletion (<i>X</i> equation)						
	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%	0.1%	1%	5%				
<i>LM</i> (1)	21.10%	3.90%	5.50%	69.50%	26.50%	5.00%	7.40%	61.10%	22.10%	5.30%	5.30%	67.20%	27.30%	7.30%	7.10%	58.30%
<i>LM</i> (2)	25.90%	4.50%	4.50%	65.10%	30.70%	6.20%	5.30%	57.80%	28.00%	3.40%	4.20%	64.40%	32.50%	4.90%	6.00%	56.60%
<i>LM</i> (3)	29.60%	5.20%	3.90%	61.30%	33.80%	5.00%	4.50%	56.70%	30.10%	3.60%	4.80%	61.50%	34.00%	4.60%	5.20%	56.20%
<i>LM</i> (4)	31.50%	4.80%	4.50%	59.20%	34.00%	4.50%	4.30%	57.10%	29.80%	4.80%	4.30%	61.10%	34.70%	4.80%	4.30%	56.20%
<i>LM</i> (5)	33.10%	4.30%	4.20%	58.40%	35.30%	3.40%	4.30%	57.00%	31.50%	4.50%	4.50%	59.50%	36.10%	4.10%	3.80%	56.00%
<i>LM</i> (6)	33.20%	5.30%	3.80%	57.70%	35.60%	4.20%	4.80%	55.50%	33.10%	5.60%	4.30%	57.00%	36.00%	4.10%	4.60%	55.30%
<i>LM</i> (7)	35.00%	4.20%	4.80%	56.00%	36.10%	4.20%	3.80%	55.90%	34.00%	5.50%	4.50%	56.00%	36.80%	3.80%	4.60%	54.80%
<i>LM</i> (8)	35.40%	4.10%	5.70%	54.80%	36.60%	3.40%	3.60%	56.40%	33.90%	5.70%	5.60%	54.80%	37.10%	4.20%	4.30%	54.30%
<i>LM</i> (9)	36.10%	4.20%	4.60%	55.00%	35.90%	4.80%	2.70%	56.70%	34.60%	5.30%	5.00%	55.00%	38.00%	4.20%	3.10%	54.80%
<i>LM</i> (10)	36.80%	3.80%	4.30%	55.00%	37.00%	4.20%	2.90%	55.90%	34.70%	5.20%	5.50%	54.60%	39.10%	3.60%	2.90%	54.30%
<i>LM</i> (11)	37.10%	4.20%	4.50%	54.20%	36.40%	4.60%	3.10%	55.90%	35.00%	4.20%	5.30%	55.50%	39.50%	3.20%	3.40%	53.90%
<i>LM</i> (12)	37.10%	3.60%	5.60%	53.60%	37.40%	3.90%	3.50%	55.20%	35.30%	4.20%	5.00%	55.50%	38.80%	3.40%	4.10%	53.80%

**Table 4.9: P-Value for Tests of Model Parameters
(OLS - Dummy Variable of Current Market Return)**

The table is based on the model defined at (4.1) using 60 days of trade data by OLS interacted with a dummy variable of current market return. The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters interacted with the corresponding dummy variable and their probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. $D_{m,t}$ is the dummy variable denoted 1 if the market return on that day is positive or 0 if it is non-positive. Bold format denotes significance at the 1% level.

	Addition						Deletion					
	Before		After		Difference		Before		After		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
D_m, R	0.056	0.000	0.102	0.000	0.046	0.021	0.277	0.000	0.143	0.009	-0.134	0.044
$\alpha_1 D_m$	0.000	0.945	0.002	0.507	0.001	0.668	-0.005	0.178	0.003	0.460	-0.002	0.731
$\alpha_2 D_m$	-0.005	0.013	-0.004	0.067	-0.001	0.780	-0.003	0.364	-0.003	0.472	-0.001	0.921
$\alpha_3 D_m$	-0.001	0.581	-0.005	0.024	0.004	0.197	0.002	0.590	0.001	0.741	0.000	0.953
$\alpha_4 D_m$	0.001	0.766	0.001	0.725	0.000	0.968	-0.002	0.533	0.001	0.759	-0.001	0.879
$\alpha_5 D_m$	0.000	0.929	0.003	0.140	0.003	0.279	0.000	0.972	0.001	0.813	0.001	0.867
$\beta_0 D_m$	-0.003	0.876	0.019	0.036	0.016	0.459	-0.052	0.014	0.032	0.480	-0.020	0.684
$\beta_1 D_m$	-0.024	0.433	-0.002	0.830	-0.022	0.475	0.033	0.068	0.432	0.034	0.399	0.051
$\beta_2 D_m$	0.042	0.144	-0.021	0.010	-0.021	0.482	0.045	0.005	-0.212	0.310	0.167	0.426
$\beta_3 D_m$	0.017	0.499	-0.018	0.043	0.001	0.960	-0.021	0.308	-1.451	0.038	1.430	0.041
$\beta_4 D_m$	-0.042	0.125	-0.005	0.541	-0.037	0.187	-0.003	0.850	0.125	0.061	0.122	0.075
$\beta_5 D_m$	-0.009	0.547	0.016	0.092	0.007	0.698	-0.007	0.636	-0.037	0.283	0.030	0.412
D_m, X	0.357	0.000	0.445	0.000	0.089	0.041	0.879	0.000	0.637	0.000	-0.242	0.026
$\gamma_1 D_m$	0.002	0.721	-0.006	0.309	0.004	0.669	-0.011	0.141	-0.004	0.630	-0.007	0.561
$\gamma_2 D_m$	-0.001	0.887	-0.005	0.424	0.004	0.661	-0.019	0.017	0.005	0.566	-0.013	0.268
$\gamma_3 D_m$	-0.009	0.171	-0.003	0.646	-0.006	0.471	-0.006	0.338	0.006	0.406	0.000	0.999
$\gamma_4 D_m$	-0.006	0.275	-0.003	0.643	-0.004	0.661	-0.006	0.381	0.001	0.898	-0.005	0.622
$\gamma_5 D_m$	0.011	0.056	0.000	0.926	-0.010	0.157	-0.004	0.548	0.007	0.300	0.004	0.694
$\delta_1 D_m$	0.002	0.244	0.002	0.403	0.000	0.879	-0.003	0.438	-0.005	0.195	0.002	0.645
$\delta_2 D_m$	-0.004	0.034	0.001	0.660	-0.003	0.236	0.002	0.540	0.001	0.785	-0.001	0.845
$\delta_3 D_m$	0.004	0.073	-0.004	0.035	0.000	0.900	0.004	0.143	-0.003	0.425	-0.001	0.812
$\delta_4 D_m$	0.002	0.256	0.000	0.911	-0.002	0.467	0.002	0.563	-0.003	0.462	0.001	0.855
$\delta_5 D_m$	-0.004	0.058	-0.001	0.640	-0.003	0.340	-0.001	0.858	0.003	0.388	0.003	0.564
$A D_m$	-0.005	0.352	-0.004	0.510	-0.001	0.891	-0.009	0.399	0.004	0.755	-0.005	0.755
$B D_m$	-0.019	0.396	-0.010	0.549	-0.009	0.754	-0.005	0.887	-1.112	0.081	1.108	0.082
ΓD_m	-0.003	0.857	-0.017	0.301	0.014	0.552	-0.045	0.028	0.016	0.501	-0.030	0.341
ΔD_m	0.001	0.853	-0.002	0.556	0.001	0.781	0.005	0.387	-0.007	0.337	0.002	0.861

**Table 4.10: P-Value for Tests of Model Parameters
(WLS - Dummy Variable of Current Market Return)**

The table is based on the model defined at (4.1) using 60 days of trade data by WLS interacted with a dummy variable of current market return. The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters interacted with the corresponding dummy variable and their probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. $D_{m,t}$ is the dummy variable denoted 1 if the market return on that day is positive or 0 if it is non-positive. Bold format denotes significance at the 1% level.

	Addition						Deletion					
	Before		After		Difference		Before		After		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
D_m, R	0.002	0.340	0.007	0.008	0.005	0.101	0.019	0.006	-0.003	0.645	-0.015	0.138
$\alpha_1 D_m$	-0.003	0.097	0.004	0.081	0.001	0.781	0.001	0.722	0.002	0.579	0.001	0.872
$\alpha_2 D_m$	-0.003	0.165	-0.002	0.512	-0.001	0.663	0.003	0.491	-0.007	0.193	0.005	0.495
$\alpha_3 D_m$	0.000	0.981	0.009	0.024	0.009	0.042	0.010	0.003	0.001	0.790	-0.009	0.086
$\alpha_4 D_m$	0.002	0.295	0.001	0.751	-0.001	0.661	0.005	0.110	-0.001	0.817	-0.004	0.457
$\alpha_5 D_m$	0.002	0.350	-0.001	0.687	0.000	0.961	-0.006	0.060	0.003	0.473	-0.003	0.580
$\beta_0 D_m$	-0.010	0.370	-0.089	0.011	0.079	0.029	-0.022	0.361	-0.053	0.346	0.031	0.610
$\beta_1 D_m$	0.002	0.841	-0.008	0.375	0.006	0.696	-0.025	0.107	0.077	0.074	0.052	0.258
$\beta_2 D_m$	0.007	0.390	-0.006	0.614	-0.001	0.930	-0.006	0.673	0.089	0.012	0.083	0.032
$\beta_3 D_m$	0.004	0.668	0.011	0.340	0.007	0.673	-0.036	0.018	0.057	0.211	0.020	0.672
$\beta_4 D_m$	-0.012	0.300	0.024	0.020	0.012	0.422	-0.004	0.765	0.031	0.389	0.027	0.484
$\beta_5 D_m$	-0.010	0.382	0.093	0.004	0.083	0.017	-0.011	0.450	0.038	0.205	0.027	0.424
D_m, X	0.097	0.000	0.098	0.000	0.000	0.994	0.163	0.000	0.108	0.000	-0.055	0.113
$\gamma_1 D_m$	0.018	0.145	0.005	0.660	-0.013	0.428	0.005	0.676	-0.021	0.212	0.016	0.435
$\gamma_2 D_m$	0.009	0.362	-0.001	0.879	-0.008	0.562	-0.001	0.934	-0.016	0.295	0.015	0.414
$\gamma_3 D_m$	0.002	0.873	0.016	0.212	0.013	0.506	0.001	0.924	0.004	0.751	0.003	0.837
$\gamma_4 D_m$	0.006	0.525	0.021	0.075	0.015	0.336	0.002	0.833	-0.017	0.236	0.015	0.389
$\gamma_5 D_m$	0.018	0.151	0.014	0.093	-0.005	0.757	0.000	0.962	0.053	0.117	0.053	0.130
$\delta_1 D_m$	0.004	0.138	0.000	0.903	-0.003	0.352	0.009	0.031	0.010	0.023	0.001	0.908
$\delta_2 D_m$	-0.003	0.209	0.002	0.257	-0.001	0.850	0.002	0.644	-0.007	0.121	0.005	0.403
$\delta_3 D_m$	0.005	0.148	-0.003	0.097	-0.002	0.655	-0.003	0.324	0.007	0.069	0.004	0.469
$\delta_4 D_m$	0.003	0.346	0.002	0.333	-0.001	0.797	-0.006	0.107	-0.001	0.722	-0.004	0.433
$\delta_5 D_m$	-0.009	0.054	-0.002	0.299	-0.007	0.189	-0.002	0.497	-0.010	0.094	0.008	0.264
$A D_m$	-0.003	0.615	0.011	0.071	0.009	0.278	0.013	0.153	-0.002	0.883	-0.012	0.465
$B D_m$	-0.018	0.380	0.025	0.378	0.007	0.829	-0.105	0.008	0.239	0.019	0.134	0.220
ΓD_m	0.055	0.066	0.054	0.045	-0.001	0.981	0.007	0.796	0.004	0.923	-0.003	0.957
ΔD_m	0.000	0.958	-0.001	0.831	0.001	0.916	0.000	0.965	-0.001	0.931	0.000	0.968

**Table 4.11: P-Value for Tests of Model Parameters
(GARCH - Dummy Variable of Current Market Return)**

The table is based on the model defined at (4.2) using 60 days of trade data by GARCH (1, 1) interacted with a dummy variable of current market return. The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters interacted with the corresponding dummy variable and their probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. $D_{m,t}$ is the dummy variable denoted 1 if the market return on that day is positive or 0 if it is non-positive. Bold format denotes significance at the 1% level.

	Addition						Deletion					
	Before		After		Difference		Before		After		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
D_m, R	0.053	0.018	0.036	0.941	-0.017	0.973	0.083	0.366	0.077	0.895	-0.006	0.992
$\alpha_1 D_m$	-0.002	0.436	0.000	1.000	-0.002	0.995	-0.007	0.286	0.004	0.981	-0.003	0.988
$\alpha_2 D_m$	-0.004	0.102	-0.006	0.991	0.001	0.998	-0.008	0.160	-0.002	0.992	-0.005	0.983
$\alpha_3 D_m$	0.000	0.967	-0.005	0.996	0.004	0.996	0.001	0.953	-0.001	0.997	0.001	0.999
$\alpha_4 D_m$	-0.002	0.699	-0.001	0.994	-0.001	0.992	-0.002	0.809	0.001	0.991	0.000	0.998
$\alpha_5 D_m$	0.001	0.844	0.003	0.939	0.002	0.955	0.004	0.360	-0.001	0.992	-0.003	0.976
$\beta_0 D_m$	-0.028	0.300	0.040	0.972	0.012	0.992	-0.047	0.652	0.033	0.969	-0.014	0.987
$\beta_1 D_m$	-0.014	0.782	-0.011	0.982	-0.003	0.995	0.005	0.966	0.502	0.842	0.497	0.843
$\beta_2 D_m$	0.032	0.379	-0.005	0.997	-0.027	0.980	0.038	0.792	-0.171	0.983	0.133	0.987
$\beta_3 D_m$	0.000	0.993	-0.021	0.986	0.020	0.987	-0.014	0.957	-1.472	0.501	1.458	0.509
$\beta_4 D_m$	-0.005	0.899	-0.003	0.998	-0.002	0.999	-0.010	0.972	0.082	0.985	0.071	0.987
$\beta_5 D_m$	0.011	0.767	0.009	0.981	-0.002	0.996	-0.031	0.659	-0.018	0.999	-0.013	0.999
D_m, X	0.253	0.000	0.377	0.000	0.124	0.008	0.757	0.000	0.624	0.000	-0.132	0.454
$\gamma_1 D_m$	0.015	0.002	-0.001	0.930	-0.014	0.106	-0.009	0.513	-0.007	0.618	-0.002	0.901
$\gamma_2 D_m$	0.014	0.008	-0.015	0.072	0.000	0.965	-0.021	0.240	0.012	0.464	-0.009	0.713
$\gamma_3 D_m$	0.008	0.105	-0.002	0.818	-0.006	0.579	-0.015	0.295	0.006	0.734	-0.008	0.715
$\gamma_4 D_m$	0.000	0.999	0.006	0.526	0.006	0.581	0.001	0.949	0.000	0.985	-0.001	0.980
$\gamma_5 D_m$	0.015	0.004	-0.009	0.256	-0.005	0.577	0.004	0.675	0.011	0.546	0.007	0.750
$\delta_1 D_m$	0.002	0.404	0.007	0.081	0.005	0.286	0.009	0.769	-0.006	0.559	-0.003	0.934
$\delta_2 D_m$	-0.004	0.037	-0.001	0.786	-0.003	0.514	0.004	0.884	0.004	0.696	0.000	0.995
$\delta_3 D_m$	0.000	0.902	0.000	0.972	0.000	0.989	0.010	0.663	-0.004	0.871	-0.006	0.867
$\delta_4 D_m$	-0.001	0.574	-0.005	0.337	0.004	0.512	0.000	0.996	-0.005	0.539	0.005	0.905
$\delta_5 D_m$	-0.002	0.385	-0.001	0.893	-0.001	0.796	-0.003	0.939	0.005	0.564	0.003	0.950
$A D_m$	-0.008	0.483	-0.008	0.996	0.001	1.000	-0.012	0.649	0.001	0.998	-0.011	0.979
$B D_m$	-0.005	0.916	0.010	0.989	0.005	0.995	-0.060	0.930	-1.044	0.916	0.983	0.921
ΓD_m	0.052	0.000	-0.021	0.351	-0.031	0.238	-0.040	0.340	0.023	0.648	-0.017	0.794
ΔD_m	-0.005	0.160	0.000	0.970	-0.005	0.617	0.020	0.508	-0.006	0.840	-0.014	0.748

**Table 4.12: P-Value for Tests of Model Parameters
(OLS - Dummy Variable of One-day Lagged Market Return)**

The table is based on the model defined at (4.1) using 60 days of trade data by OLS interacted with a dummy variable of one-day lagged market return. The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters interacted with the corresponding dummy variable and their probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. $D_{mlag,t}$ is the dummy variable denoted 1 if the market return on that day is positive or 0 if it is non-positive. Bold format denotes significance at the 1% level.

	Addition						Deletion					
	Before		After		Difference		Before		After		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
D_{mlag}, R	0.007	0.602	0.000	0.997	-0.007	0.754	-0.031	0.476	-0.155	0.006	0.124	0.079
$\alpha_1 D_{mlag}$	0.000	0.978	-0.003	0.267	0.003	0.431	-0.001	0.787	-0.004	0.296	0.003	0.558
$\alpha_2 D_{mlag}$	-0.001	0.553	-0.001	0.736	-0.001	0.871	-0.002	0.548	-0.006	0.124	0.004	0.482
$\alpha_3 D_{mlag}$	0.001	0.521	-0.004	0.057	0.003	0.388	0.003	0.320	-0.006	0.129	0.002	0.651
$\alpha_4 D_{mlag}$	-0.004	0.057	-0.004	0.081	0.000	0.993	-0.002	0.581	-0.005	0.185	0.003	0.519
$\alpha_5 D_{mlag}$	-0.002	0.116	-0.001	0.403	-0.001	0.688	-0.003	0.216	-0.004	0.231	0.000	0.934
$\beta_0 D_{mlag}$	0.005	0.804	-0.009	0.269	0.005	0.817	-0.064	0.005	0.101	0.078	0.037	0.548
$\beta_1 D_{mlag}$	-0.022	0.108	-0.016	0.023	-0.005	0.719	0.012	0.492	-0.079	0.133	0.067	0.229
$\beta_2 D_{mlag}$	-0.016	0.472	0.009	0.200	-0.006	0.787	0.004	0.798	0.171	0.155	0.168	0.166
$\beta_3 D_{mlag}$	0.044	0.004	0.006	0.426	-0.038	0.028	0.025	0.128	0.069	0.147	0.043	0.385
$\beta_4 D_{mlag}$	-0.073	0.000	0.003	0.703	-0.070	0.000	0.019	0.188	0.001	0.966	-0.018	0.591
$\beta_5 D_{mlag}$	0.111	0.000	0.006	0.588	-0.105	0.000	-0.018	0.268	0.014	0.638	-0.004	0.913
D_{mlag}, X	0.106	0.000	0.018	0.580	-0.088	0.038	-0.057	0.408	-0.027	0.767	-0.030	0.789
$\gamma_1 D_{mlag}$	-0.015	0.023	-0.006	0.329	-0.009	0.314	-0.013	0.090	-0.009	0.261	-0.003	0.758
$\gamma_2 D_{mlag}$	-0.008	0.206	-0.005	0.360	-0.003	0.746	0.007	0.342	-0.010	0.219	0.003	0.775
$\gamma_3 D_{mlag}$	-0.006	0.392	-0.001	0.810	-0.004	0.634	-0.003	0.606	-0.011	0.162	0.007	0.474
$\gamma_4 D_{mlag}$	-0.004	0.522	-0.009	0.117	0.005	0.523	-0.003	0.660	-0.021	0.005	0.018	0.068
$\gamma_5 D_{mlag}$	0.006	0.327	0.004	0.379	-0.001	0.847	0.000	0.985	-0.012	0.065	0.012	0.170
$\delta_1 D_{mlag}$	0.006	0.004	0.002	0.320	-0.003	0.249	-0.002	0.490	0.004	0.333	0.001	0.796
$\delta_2 D_{mlag}$	0.000	0.932	0.002	0.334	0.002	0.519	0.002	0.520	-0.005	0.186	0.003	0.586
$\delta_3 D_{mlag}$	0.002	0.362	0.003	0.195	0.001	0.740	0.002	0.547	0.010	0.007	0.008	0.092
$\delta_4 D_{mlag}$	0.004	0.125	-0.003	0.219	-0.001	0.751	-0.003	0.405	0.007	0.053	0.004	0.352
$\delta_5 D_{mlag}$	-0.001	0.464	0.001	0.597	0.000	0.914	0.007	0.023	-0.005	0.213	-0.002	0.631
$A D_{mlag}$	-0.006	0.338	-0.013	0.055	0.007	0.454	-0.005	0.631	-0.025	0.032	0.020	0.204
$B D_{mlag}$	0.049	0.016	-0.001	0.953	-0.047	0.111	-0.021	0.552	0.277	0.108	0.256	0.147
ΓD_{mlag}	-0.026	0.110	-0.017	0.285	-0.009	0.689	-0.012	0.565	-0.064	0.006	0.052	0.091
ΔD_{mlag}	0.009	0.007	0.005	0.163	-0.004	0.442	0.006	0.312	0.011	0.097	0.005	0.558

**Table 4.13: P-Value for Tests of Model Parameters
(WLS - Dummy Variable of One-day Lagged Market Return)**

The table is based on the model defined at (4.1) using 60 days of trade data by WLS interacted with a dummy variable of one-day lagged market return. The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters interacted with the corresponding dummy variable and their probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. $D_{mlag,t}$ is the dummy variable denoted 1 if the market return on that day is positive or 0 if it is non-positive. Bold format denotes significance at the 1% level.

	Addition						Deletion					
	Before		After		Difference		Before		After		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
D_{mlag}, R	-0.004	0.025	-0.002	0.515	-0.003	0.377	-0.008	0.263	-0.007	0.399	-0.001	0.961
α_1-D_{mlag}	0.000	0.869	-0.001	0.766	0.000	0.964	-0.008	0.038	-0.006	0.159	-0.002	0.756
α_2-D_{mlag}	0.001	0.714	-0.001	0.608	0.000	0.966	-0.006	0.105	0.004	0.546	-0.002	0.727
α_3-D_{mlag}	-0.005	0.116	-0.008	0.002	0.003	0.494	-0.003	0.478	-0.010	0.012	0.007	0.206
α_4-D_{mlag}	0.000	0.939	-0.003	0.217	0.003	0.427	-0.004	0.277	0.010	0.016	0.007	0.221
α_5-D_{mlag}	-0.004	0.056	-0.001	0.584	-0.002	0.432	-0.001	0.696	-0.004	0.321	0.003	0.598
β_0-D_{mlag}	-0.008	0.614	-0.004	0.752	-0.004	0.828	-0.118	0.000	-0.002	0.989	-0.116	0.401
β_1-D_{mlag}	-0.080	0.000	0.004	0.610	-0.076	0.000	0.031	0.131	-0.035	0.444	0.003	0.949
β_2-D_{mlag}	-0.005	0.576	0.051	0.000	0.046	0.000	0.024	0.125	-0.036	0.262	0.012	0.735
β_3-D_{mlag}	0.001	0.952	0.008	0.426	0.007	0.659	-0.005	0.824	0.010	0.781	0.005	0.896
β_4-D_{mlag}	0.010	0.397	-0.004	0.728	-0.006	0.719	0.053	0.021	-0.040	0.460	-0.013	0.821
β_5-D_{mlag}	-0.004	0.743	0.008	0.467	0.004	0.801	0.010	0.548	-0.045	0.343	0.035	0.481
D_{mlag}, X	0.030	0.015	-0.008	0.498	-0.022	0.202	0.002	0.932	-0.030	0.212	0.028	0.404
γ_1-D_{mlag}	-0.011	0.388	0.002	0.854	-0.009	0.610	-0.019	0.088	-0.004	0.848	-0.015	0.483
γ_2-D_{mlag}	-0.019	0.092	0.007	0.473	-0.012	0.399	-0.003	0.793	0.015	0.313	0.012	0.484
γ_3-D_{mlag}	0.000	0.976	-0.016	0.231	0.016	0.435	0.013	0.128	-0.016	0.224	0.003	0.834
γ_4-D_{mlag}	-0.007	0.511	-0.004	0.779	-0.002	0.891	0.012	0.144	0.002	0.897	-0.010	0.491
γ_5-D_{mlag}	0.011	0.430	0.009	0.283	-0.002	0.893	0.017	0.022	-0.013	0.275	-0.004	0.773
δ_1-D_{mlag}	0.007	0.003	0.006	0.021	-0.001	0.753	0.009	0.025	0.011	0.071	0.002	0.788
δ_2-D_{mlag}	-0.001	0.662	0.001	0.792	-0.001	0.872	0.003	0.468	-0.008	0.061	0.005	0.411
δ_3-D_{mlag}	0.003	0.084	0.004	0.078	0.000	0.910	0.002	0.499	-0.007	0.240	0.005	0.469
δ_4-D_{mlag}	0.008	0.200	-0.004	0.055	-0.004	0.580	-0.002	0.555	0.008	0.108	0.006	0.325
δ_5-D_{mlag}	-0.010	0.304	0.007	0.001	-0.003	0.795	0.004	0.169	0.002	0.698	-0.003	0.617
$A-D_{mlag}$	-0.007	0.375	-0.013	0.037	0.006	0.565	-0.022	0.032	-0.007	0.600	-0.015	0.361
$B-D_{mlag}$	-0.086	0.004	0.063	0.035	-0.023	0.589	-0.004	0.914	-0.147	0.354	0.143	0.383
$\Gamma-D_{mlag}$	-0.026	0.403	-0.003	0.924	-0.023	0.586	0.021	0.367	-0.017	0.664	-0.004	0.921
$\Delta-D_{mlag}$	0.008	0.216	0.013	0.002	0.006	0.454	0.017	0.012	0.006	0.464	-0.011	0.296

**Table 4.14: P-Value for Tests of Model Parameters
(GARCH - Dummy Variable of One-day Lagged Market Return)**

The table is based on the model defined at (4.1) using 60 days of trade data by GARCH (1, 1) interacted with a dummy variable of one-day lagged market return. The tests for individual parameters and sums are computed as described in Section 2.3.3. The cell entries show the estimated parameters interacted with the corresponding dummy variable and their probabilities (p-val.), respectively. Differences are calculated by absolute value of model parameters between after and before both events. $D_{mlag,t}$ is the dummy variable denoted 1 if the market return on that day is positive or 0 if it is non-positive. Bold format denotes significance at the 1% level.

	Addition						Deletion					
	Before		After		Difference		Before		After		Difference	
	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.	Coef.	p-val.
D_{mlag}, R	-0.011	0.940	0.019	0.981	0.008	0.992	-0.029	0.574	0.045	0.843	0.016	0.944
$\alpha_1 \cdot D_{mlag}$	0.002	0.615	-0.004	0.908	0.002	0.949	-0.004	0.266	0.002	0.950	-0.003	0.919
$\alpha_2 \cdot D_{mlag}$	-0.002	0.565	0.006	0.819	0.005	0.866	0.000	0.988	0.009	0.733	0.008	0.738
$\alpha_3 \cdot D_{mlag}$	0.000	0.997	-0.002	0.936	0.002	0.965	0.001	0.750	-0.005	0.859	0.004	0.891
$\alpha_4 \cdot D_{mlag}$	-0.003	0.956	-0.002	0.834	-0.001	0.988	0.002	0.605	0.002	0.937	0.001	0.986
$\alpha_5 \cdot D_{mlag}$	-0.001	0.955	-0.001	0.915	-0.001	0.984	0.000	0.886	0.000	0.999	0.000	0.975
$\beta_0 \cdot D_{mlag}$	0.009	0.982	-0.010	0.802	0.001	0.998	-0.048	0.002	0.112	0.284	0.065	0.542
$\beta_1 \cdot D_{mlag}$	0.021	0.961	-0.004	0.925	-0.017	0.968	-0.044	0.049	-0.033	0.787	-0.011	0.929
$\beta_2 \cdot D_{mlag}$	-0.023	0.921	-0.008	0.767	-0.014	0.950	-0.001	0.972	-0.001	0.993	0.000	0.999
$\beta_3 \cdot D_{mlag}$	-0.006	0.984	-0.003	0.935	-0.003	0.992	-0.007	0.713	0.041	0.744	0.034	0.790
$\beta_4 \cdot D_{mlag}$	-0.054	0.889	-0.009	0.799	-0.045	0.908	-0.004	0.819	-0.038	0.702	0.034	0.734
$\beta_5 \cdot D_{mlag}$	0.136	0.885	-0.018	0.664	-0.118	0.901	-0.024	0.160	0.049	0.651	0.025	0.819
D_{mlag}, X	0.100	0.000	0.031	0.469	-0.069	0.149	-0.094	0.383	-0.097	0.499	0.004	0.984
$\gamma_1 \cdot D_{mlag}$	-0.030	0.000	0.000	0.976	-0.030	0.001	-0.001	0.946	0.000	0.993	-0.001	0.964
$\gamma_2 \cdot D_{mlag}$	-0.010	0.047	0.006	0.447	-0.004	0.677	0.019	0.492	-0.020	0.278	0.001	0.972
$\gamma_3 \cdot D_{mlag}$	-0.010	0.038	0.000	0.996	-0.010	0.326	0.008	0.621	-0.011	0.541	0.003	0.900
$\gamma_4 \cdot D_{mlag}$	-0.006	0.243	-0.006	0.505	0.000	0.987	0.016	0.478	-0.024	0.152	0.009	0.761
$\gamma_5 \cdot D_{mlag}$	0.014	0.005	-0.010	0.203	-0.004	0.677	0.006	0.606	-0.025	0.093	0.019	0.328
$\delta_1 \cdot D_{mlag}$	0.004	0.641	0.001	0.904	-0.003	0.778	0.005	0.617	-0.004	0.890	-0.001	0.969
$\delta_2 \cdot D_{mlag}$	0.001	0.892	0.005	0.356	0.004	0.657	-0.001	0.836	0.002	0.946	0.001	0.984
$\delta_3 \cdot D_{mlag}$	0.003	0.606	0.005	0.523	0.002	0.865	0.005	0.745	0.015	0.608	0.010	0.755
$\delta_4 \cdot D_{mlag}$	0.002	0.785	-0.003	0.565	0.001	0.951	-0.002	0.773	0.002	0.931	0.000	0.986
$\delta_5 \cdot D_{mlag}$	-0.004	0.658	0.003	0.656	-0.001	0.923	-0.002	0.845	-0.002	0.919	0.001	0.975
$A \cdot D_{mlag}$	-0.004	0.965	-0.002	0.981	-0.002	0.988	-0.001	0.924	0.007	0.941	0.006	0.949
$B \cdot D_{mlag}$	0.084	0.909	-0.053	0.527	-0.031	0.967	-0.128	0.002	0.131	0.681	0.003	0.994
$\Gamma \cdot D_{mlag}$	-0.042	0.002	-0.010	0.656	-0.033	0.199	0.047	0.380	-0.080	0.111	0.033	0.657
$\Delta \cdot D_{mlag}$	0.006	0.453	0.011	0.253	0.004	0.739	0.005	0.722	0.013	0.770	0.008	0.859

Table 4.15: P-Value for Tests of Model Parameters - Crisis (Addition)

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the addition events are computed as described in Section 2.3.3 for crisis and non-crisis periods. The cell entries of the first panel show the number of observations in crisis and non-crisis periods. The cell entries of other panels show the estimated parameters and probabilities (p-val.), respectively. Differences for crisis and non-crisis periods are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between crisis and non-crisis periods before and after the events are calculated by absolute value of model parameters between crisis and non-crisis periods (Crisis minus Non-crisis) before and after the events. The crisis period is between December 2007 and June 2009. The event is defined as in the crisis period if the whole estimation period of the event (60 days before and after the effective date) is in the crisis period. Crisis = 1 indicates crisis period and crisis = 0 indicates non-crisis period. Bold format denotes significance at the 1% level.

	CRISIS = 0 (1023)				Addition CRISIS = 1 (59)				CRISIS 1 - 0			
	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Diff. in Before	p-val.	Diff. in After	p-val.
	OLS											
A	-0.445	-0.445	-0.001	0.897	-0.533	-0.520	-0.013	0.641	0.087	0.000	0.075	0.000
Δ	0.333	0.324	-0.009	0.003	0.327	0.287	-0.040	0.004	-0.006	0.520	-0.038	0.000
	WLS											
A	-0.238	-0.226	-0.013	0.009	-0.267	-0.363	0.096	0.000	0.029	0.044	0.137	0.000
Δ	0.317	0.305	-0.012	0.000	0.288	0.263	-0.025	0.134	-0.030	0.010	-0.042	0.001
	VAR											
A	-0.410	-0.403	-0.007	0.006	-0.466	-0.485	0.019	0.126	0.056	0.000	0.082	0.000
Δ	0.327	0.319	-0.008	0.000	0.316	0.287	-0.029	0.000	-0.011	0.028	-0.033	0.000
	WVAR											
A	-0.223	-0.211	-0.012	0.000	-0.238	-0.269	0.031	0.006	0.015	0.037	0.058	0.000
Δ	0.315	0.300	-0.015	0.000	0.279	0.246	-0.033	0.000	-0.036	0.000	-0.054	0.000
	GARCH											
A	-0.473	-0.474	0.000	0.943	-0.513	-0.556	0.043	0.458	0.040	0.000	0.083	0.150
Δ	0.382	0.375	-0.007	0.007	0.411	0.309	-0.102	0.000	0.028	0.000	-0.066	0.000
	BEKK											
A	-0.436	-0.421	-0.015	0.000	-0.512	-0.537	0.024	0.153	0.076	0.000	0.115	0.000
Δ	0.373	0.359	-0.014	0.000	0.396	0.302	-0.093	0.000	0.023	0.000	-0.056	0.000

Table 4.16: P-Value for Tests of Model Parameters - Crisis (Deletion)

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the deletion events are computed as described in Section 2.3.3 for crisis and non-crisis periods. The cell entries of the first panel show the number of observations in crisis and non-crisis periods. The cell entries of other panels show the estimated parameters and probabilities (p-val.), respectively. Differences for crisis and non-crisis periods are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between crisis and non-crisis periods before and after the events are calculated by absolute value of model parameters between crisis and non-crisis periods (Crisis minus Non-crisis) before and after the events. The crisis period is between December 2007 and June 2009. The event is defined as in the crisis period if the whole estimation period of the event (60 days before and after the effective date) is in the crisis period. Crisis = 1 indicates crisis period and crisis = 0 indicates non-crisis period. Bold format denotes significance at the 1% level.

	CRISIS = 0 (469)				Deletion CRISIS = 1 (221)				CRISIS 1 - 0			
	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Diff. in Before	p-val.	Diff. in After	p-val.
	OLS											
A	-0.519	-0.552	0.033	0.001	-0.602	-0.636	0.034	0.078	0.082	0.000	0.083	0.000
Δ	0.342	0.345	0.003	0.545	0.311	0.324	0.013	0.218	-0.031	0.000	-0.022	0.018
	WLS											
A	-0.251	-0.267	0.016	0.059	-0.291	-0.279	-0.012	0.440	0.040	0.001	0.012	0.327
Δ	0.319	0.312	-0.007	0.274	0.274	0.290	0.017	0.118	-0.046	0.000	-0.022	0.020
	VAR											
A	-0.453	-0.475	0.022	0.000	-0.522	-0.538	0.017	0.058	0.069	0.000	0.064	0.000
Δ	0.331	0.338	0.007	0.055	0.291	0.311	0.019	0.001	-0.039	0.000	-0.027	0.000
	WVAR											
A	-0.229	-0.238	0.010	0.038	-0.243	-0.256	0.013	0.099	0.014	0.018	0.018	0.012
Δ	0.312	0.308	-0.004	0.325	0.260	0.284	0.024	0.000	-0.051	0.000	-0.024	0.000
	GARCH											
A	-0.550	-0.573	0.023	0.017	-0.616	-0.634	0.018	0.647	0.066	0.000	0.062	0.098
Δ	0.380	0.389	0.009	0.093	0.344	0.366	0.022	0.121	-0.036	0.000	-0.023	0.107
	BEKK											
A	-0.477	-0.492	0.015	0.007	-0.529	-0.540	0.011	0.275	0.052	0.000	0.048	0.000
Δ	0.359	0.363	0.004	0.325	0.330	0.324	-0.006	0.402	-0.029	0.000	-0.038	0.000

**Table 4.17: P-Value for Tests of Model Parameters
Down and Up Markets (Addition)**

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1), and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the addition events are computed as described in Section 2.3.3 for down and up markets. The cell entries show the estimated parameters and probabilities (p-val.), respectively. Differences for down and up markets are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between down and up markets after the addition events are calculated by absolute value of model parameters after the event between down and up markets (Down minus Up). Transaction data before and after the event is divided into up (down) market group if the corresponding daily market return is positive (non-positive). Bold format denotes significance at the 1% level.

	Up Market				Addition Down Market				Down - Up			
	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Diff. in Before	p-val.	Diff. in After	p-val.
	OLS											
A	-0.458	-0.453	-0.005	0.366	-0.446	-0.450	0.004	0.464	-0.013	0.024	-0.003	0.608
Δ	0.333	0.328	-0.005	0.162	0.330	0.324	-0.005	0.202	-0.003	0.417	-0.003	0.369
	WLS											
A	-0.253	-0.238	-0.015	0.003	-0.246	-0.239	-0.006	0.229	-0.007	0.130	0.001	0.788
Δ	0.317	0.308	-0.009	0.030	0.316	0.307	-0.009	0.087	-0.002	0.765	-0.002	0.729
	VAR											
A	-0.427	-0.414	-0.013	0.000	-0.410	-0.416	0.005	0.139	-0.017	0.000	0.001	0.728
Δ	0.320	0.314	-0.006	0.011	0.321	0.312	-0.008	0.001	0.000	0.868	-0.002	0.413
	WVAR											
A	-0.240	-0.220	-0.020	0.000	-0.228	-0.227	-0.001	0.739	-0.012	0.000	0.007	0.023
Δ	0.307	0.298	-0.009	0.000	0.305	0.294	-0.011	0.000	-0.003	0.255	-0.005	0.078
	GARCH											
A	-0.485	-0.482	-0.003	0.555	-0.472	-0.472	0.000	0.982	-0.013	0.114	-0.010	0.061
Δ	0.387	0.382	-0.005	0.467	0.375	0.372	-0.003	0.468	-0.012	0.041	-0.010	0.016
	BEKK											
A	-0.450	-0.440	-0.011	0.005	-0.433	-0.434	0.001	0.830	-0.017	0.000	-0.006	0.164
Δ	0.365	0.357	-0.007	0.006	0.358	0.345	-0.013	0.000	-0.006	0.016	-0.013	0.000

**Table 4.18: P-Value for Tests of Model Parameters
Down and Up Markets (Deletion)**

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1), and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the deletion events are computed as described in Section 2.3.3 for down and up markets. The cell entries show the estimated parameters and probabilities (p-val.), respectively. Differences for down and up markets are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between down and up markets after the addition events are calculated by absolute value of model parameters after the event between down and up markets (Down minus Up). Transaction data before and after the event is divided into up (down) market group if the corresponding daily market return is positive (non-positive). Bold format denotes significance at the 1% level.

	Up Market				Deletion				Down - Up			
	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Diff. in Before	p-val.	Diff. in After	p-val.
	OLS											
A	-0.550	-0.584	0.034	0.002	-0.540	-0.578	0.038	0.000	-0.010	0.318	-0.006	0.607
Δ	0.335	0.339	0.004	0.529	0.333	0.346	0.012	0.048	-0.001	0.841	0.007	0.327
	WLS											
A	-0.261	-0.270	0.008	0.360	-0.282	-0.276	-0.006	0.560	0.020	0.029	0.006	0.539
Δ	0.310	0.310	0.000	0.993	0.303	0.311	0.009	0.269	-0.007	0.334	0.001	0.893
	VAR											
A	-0.476	-0.510	0.034	0.000	-0.477	-0.501	0.024	0.000	0.001	0.918	-0.009	0.176
Δ	0.310	0.318	0.008	0.050	0.311	0.321	0.011	0.015	0.001	0.876	0.003	0.487
	WVAR											
A	-0.237	-0.249	0.012	0.041	-0.256	-0.261	0.005	0.442	0.019	0.001	0.012	0.060
Δ	0.289	0.296	0.007	0.115	0.282	0.292	0.010	0.039	-0.006	0.123	-0.004	0.421
	GARCH											
A	-0.566	-0.591	0.025	0.013	-0.559	-0.591	0.032	0.007	-0.007	0.373	0.000	0.985
Δ	0.368	0.377	0.008	0.418	0.374	0.383	0.009	0.416	0.006	0.478	0.006	0.613
	BEKK											
A	-0.499	-0.521	0.022	0.002	-0.511	-0.512	0.001	0.935	0.013	0.093	-0.008	0.245
Δ	0.339	0.340	0.001	0.849	0.343	0.347	0.004	0.451	0.004	0.413	0.007	0.138

Table 4.19: P-Value for Tests of Model Parameters - MVRANK (Addition)

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the addition events are computed as described in Section 2.3.3 in four-sized quartiles. The cell entries of the first panel show the number of observations in each size quartile. The cell entries of the second panel show the estimated parameters and probabilities (p-val.), respectively. Differences for each size quartile are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between the largest and the smallest size quartiles after the addition events are calculated by absolute value of model parameters after the event between the largest and the smallest size quartiles (Largest minus Smallest). For each addition event, stocks are ranked by their average daily market value in the previous 60 days before the event date. $MVRANK = 1$ indicates the smallest market value and $MVRANK = 4$ indicates the largest market value. Sums of A and Δ are the average within each size quartile. Bold format denotes significance at the 1% level.

	Addition																	
	MVRANK = 1 (258)			MVRANK = 2 (270)			MVRANK = 3 (280)			MVRANK = 4 (264)			MVRANK 4 - 1					
	Before	After	p-val.	Before	After	p-val.	Before	After	p-val.	Before	After	p-val.	Diff. in After	p-val.				
A	-0.488	-0.490	0.002	0.849	-0.479	-0.491	0.012	0.187	-0.456	-0.431	-0.025	0.013	-0.376	-0.380	0.005	0.628	-0.110	0.000
Δ	0.336	0.327	-0.009	0.110	0.347	0.328	-0.019	0.000	0.334	0.314	-0.021	0.001	0.312	0.314	0.002	0.659	-0.013	0.025
A	-0.247	-0.236	-0.011	0.194	-0.246	-0.252	0.006	0.592	-0.244	-0.224	-0.020	0.033	-0.217	-0.222	0.005	0.500	-0.014	0.117
Δ	0.325	0.311	-0.014	0.050	0.326	0.300	-0.026	0.000	0.314	0.304	-0.009	0.165	0.296	0.292	-0.005	0.416	-0.020	0.004
A	-0.453	-0.453	0.000	0.983	-0.445	-0.447	0.002	0.723	-0.408	-0.390	-0.018	0.000	-0.345	-0.340	-0.005	0.245	-0.113	0.000
Δ	0.332	0.324	-0.008	0.024	0.340	0.321	-0.019	0.000	0.327	0.315	-0.012	0.000	0.305	0.308	0.003	0.376	-0.016	0.000
A	-0.253	-0.216	-0.017	0.000	-0.235	-0.228	-0.007	0.095	-0.227	-0.210	-0.017	0.000	-0.198	-0.202	0.004	0.375	-0.014	0.002
Δ	0.321	0.305	-0.016	0.000	0.324	0.294	-0.030	0.000	0.311	0.298	-0.013	0.000	0.294	0.290	-0.004	0.194	-0.015	0.000

Table continued on the following page.

Table 4.19: Continued from previous page

	MVRANK = 1 (258)		MVRANK = 2 (270)		Addition MVRANK = 3 (280)		MVRANK = 4 (264)		MVRANK 4 - 1									
					GARCH													
A	-0.518	-0.519	0.002	0.770	-0.514	-0.516	0.002	0.873	-0.473	-0.463	-0.010	0.405	-0.398	-0.410	0.012	0.149	-0.109	0.000
Δ	0.381	0.375	-0.006	0.192	0.398	0.375	-0.023	0.000	0.383	0.363	-0.020	0.002	0.373	0.369	-0.004	0.361	-0.006	0.189
									BEKK									
A	-0.481	-0.471	-0.010	0.096	-0.472	-0.473	0.001	0.916	-0.443	-0.414	-0.029	0.000	-0.365	-0.351	-0.013	0.008	-0.120	0.000
Δ	0.365	0.357	-0.009	0.034	0.376	0.355	-0.021	0.000	0.387	0.356	-0.031	0.000	0.366	0.353	-0.013	0.000	-0.004	0.336

Table 4.20: P-Value for Tests of Model Parameters - MVRANK (Deletion)

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the deletion events are computed as described in Section 2.3.3 in four-sized quartiles. The cell entries of the first panel show the number of observations in each size quartile. The cell entries of the second panel show the estimated parameters and probabilities (p-val.), respectively. Differences for each size quartile are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between the largest and the smallest size quartiles after the deletion events are calculated by absolute value of model parameters after the event between the largest and the smallest size quartiles (Largest minus Smallest). For each deletion event, stocks are ranked by their average daily market value in the previous 60 days before the event date. $MVRANK = 1$ indicates the smallest market value and $MVRANK = 4$ indicates the largest market value. Sums of A and Δ are the average within each size quartile. Bold format denotes significance at the 1% level.

	Deletion															
	MVRANK = 1 (159)			MVRANK = 2 (177)			MVRANK = 3 (183)			MVRANK = 4 (164)			MVRANK 4 - 1			
	Before	After	p-val.	Before	After	p-val.	Before	After	p-val.	Before	After	p-val.	Diff. in After	p-val.		
A	-0.571	-0.569	0.908	-0.552	-0.597	0.045	0.009	-0.552	-0.589	0.037	0.065	-0.506	-0.560	0.008	-0.008	0.671
Δ	0.332	0.363	0.001	0.336	0.335	-0.002	0.864	0.318	0.324	0.006	0.588	0.341	0.334	-0.006	0.533	0.005
A	-0.266	-0.280	0.014	-0.264	-0.275	0.011	0.467	-0.287	-0.265	-0.022	0.189	-0.236	-0.263	0.027	0.074	-0.017
Δ	0.307	0.342	0.001	0.305	0.306	0.001	0.937	0.287	0.284	-0.003	0.822	0.318	0.288	-0.031	0.024	-0.055
A	-0.505	-0.513	0.008	-0.482	-0.516	0.034	0.000	-0.468	-0.486	0.019	0.058	-0.447	-0.468	0.021	0.025	-0.044
Δ	0.326	0.350	0.000	0.327	0.330	0.003	0.567	0.305	0.318	0.013	0.041	0.314	0.318	0.005	0.467	0.000
A	-0.246	-0.264	0.018	-0.233	-0.243	0.010	0.193	-0.235	-0.232	-0.003	0.707	-0.218	-0.239	0.021	0.018	-0.025
Δ	0.301	0.330	0.000	0.302	0.304	0.003	0.639	0.278	0.279	0.001	0.933	0.299	0.288	-0.010	0.121	-0.041

Table continued on the following page.

Table 4.20: Continued from previous page

	Deletion																	
	MVFRANK = 1 (159)		MVFRANK = 2 (177)		MVFRANK = 3 (183)		MVFRANK = 4 (164)		MVFRANK 4 - 1									
	GARCH																	
A	-0.581	-0.600	0.020	0.142	-0.577	-0.613	0.036	0.423	-0.570	-0.604	0.035	0.170	-0.555	-0.546	-0.008	0.612	-0.054	0.004
Δ	0.363	0.404	0.042	0.000	0.376	0.365	-0.011	0.427	0.348	0.365	0.017	0.227	0.383	0.394	0.010	0.250	-0.010	0.287
	BEKK																	
A	-0.497	-0.524	0.027	0.004	-0.490	-0.528	0.038	0.000	-0.510	-0.497	-0.014	0.203	-0.475	-0.482	0.007	0.521	-0.042	0.000
Δ	0.357	0.378	0.021	0.001	0.361	0.344	-0.016	0.009	0.326	0.335	0.008	0.214	0.355	0.346	-0.010	0.221	-0.032	0.000

Table 4.21: P-Value for Tests of Model Parameters - VORANK (Addition)

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the addition events are computed as described in Section 2.3.3 in four volume ratio quartiles. The cell entries of the first panel show the number of observations in each volume ratio quartile. The cell entries of the second panel show the estimated parameters and probabilities (p-val.), respectively. Differences for each volume ratio quartile are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between the largest and the smallest volume ratio quartiles after the addition events are calculated by absolute value of model parameters after the event between the largest and the smallest volume ratio quartiles (Largest minus Smallest). For each addition event, stocks are ranked by their average daily trading volume divided by the corresponding average daily number of shares outstanding in the previous 60 days before the event date. VORANK = 1 indicates the smallest trading volume ratio and VORANK = 4 indicates the largest trading volume ratio. Sums of A and Δ are the average within each volume ratio quartile. Bold format denotes significance at the 1% level.

	VORANK = 1 (245)				VORANK = 2 (271)				Addition				VORANK = 3 (288)				VORANK = 4 (268)				VORANK 4 - 1				
	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Before	After	Before	After	Diff.	p-val.	Before	After	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	
A	-0.494	-0.489	-0.005	0.730	-0.448	-0.447	-0.001	0.945	-0.448	-0.441	-0.007	0.358	-0.414	-0.420	0.006	0.254	-0.069	0.000							
Δ	0.336	0.331	-0.005	0.592	0.322	0.323	0.001	0.823	0.335	0.315	-0.020	0.000	0.337	0.316	-0.022	0.000	-0.015	0.028							
A	-0.231	-0.238	0.006	0.599	-0.236	-0.208	-0.029	0.013	-0.244	-0.249	0.005	0.498	-0.242	-0.239	-0.003	0.526	0.001	0.914							
Δ	0.304	0.308	0.004	0.698	0.301	0.290	-0.012	0.053	0.323	0.300	-0.023	0.000	0.331	0.310	-0.021	0.000	0.002	0.759							
A	-0.420	-0.421	0.002	0.795	-0.411	-0.393	-0.019	0.000	-0.420	-0.409	-0.012	0.004	-0.400	-0.407	0.007	0.026	-0.015	0.009							
Δ	0.323	0.333	0.010	0.043	0.315	0.313	-0.003	0.434	0.331	0.310	-0.020	0.000	0.335	0.314	-0.022	0.000	-0.020	0.000							
A	-0.208	-0.208	0.000	0.998	-0.213	-0.198	-0.014	0.001	-0.234	-0.216	-0.018	0.000	-0.236	-0.233	-0.003	0.231	0.024	0.000							
Δ	0.303	0.301	-0.002	0.663	0.297	0.282	-0.015	0.000	0.318	0.297	-0.022	0.000	0.329	0.307	-0.023	0.000	0.006	0.206							

Table continued on the following page.

Table 4.21: Continued from previous page

		Addition																
		VORANK = 1 (245)		VORANK = 2 (271)		VORANK = 3 (288)		VORANK = 4 (268)		VORANK 4 - 1								
A	-0.509	-0.508	0.000	0.979	-0.469	-0.484	0.015	0.008	-0.477	-0.464	-0.012	0.334	-0.450	-0.454	0.004	0.176	-0.054	0.000
Δ	0.375	0.376	0.001	0.868	0.390	0.366	-0.024	0.000	0.384	0.367	-0.017	0.000	0.386	0.373	-0.014	0.000	-0.004	0.612
									BEKK									
A	-0.437	-0.424	-0.014	0.085	-0.443	-0.435	-0.008	0.194	-0.448	-0.422	-0.026	0.000	-0.431	-0.428	-0.004	0.292	0.004	0.540
Δ	0.367	0.377	0.010	0.054	0.359	0.346	-0.013	0.001	0.370	0.350	-0.019	0.000	0.400	0.350	-0.050	0.000	-0.027	0.000

Table 4.22: P-Value for Tests of Model Parameters - VORANK (Deletion)

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the deletion events are computed as described in Section 2.3.3 in four volume ratio quartiles. The cell entries of the first panel show the number of observations in each volume ratio quartile. The cell entries of the second panel show the estimated parameters and probabilities (p-val.), respectively. Differences for each volume ratio quartile are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between the largest and the smallest volume ratio quartiles after the deletion events are calculated by absolute value of model parameters after the event between the largest and the smallest volume ratio quartiles (Largest minus Smallest). For each deletion event, stocks are ranked by their average daily trading volume divided by the corresponding average daily number of shares outstanding in the previous 60 days before the event date. $VORANK = 1$ indicates the smallest trading volume ratio and $VORANK = 4$ indicates the largest trading volume ratio. Sums of A and Δ are the average within each volume ratio quartile. Bold format denotes significance at the 1% level.

		Deletion															
		VORANK = 1 (142)			VORANK = 2 (184)			VORANK = 3 (190)			VORANK = 4 (167)			VORANK 4 - 1			
		Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.
A	-0.582	-0.647	0.065	0.010	0.015	0.483	-0.528	-0.566	0.038	0.011	-0.522	-0.546	0.024	0.076	-0.101	-0.101	0.000
Δ	0.340	0.351	0.011	0.409	0.013	0.251	0.332	0.317	-0.015	0.070	0.319	0.341	0.021	0.002	-0.011	-0.011	0.359
A	-0.319	-0.311	-0.008	0.713	0.011	0.479	-0.254	-0.260	0.007	0.607	-0.249	-0.263	0.014	0.223	-0.048	-0.048	0.006
Δ	0.305	0.309	0.003	0.818	0.008	0.533	0.307	0.291	-0.016	0.134	0.296	0.321	0.025	0.004	0.012	0.012	0.305
A	-0.486	-0.524	0.039	0.002	0.003	0.746	-0.459	-0.479	0.021	0.008	-0.486	-0.510	0.024	0.000	-0.014	-0.014	0.173
Δ	0.326	0.334	0.008	0.304	0.028	0.000	0.319	0.314	-0.005	0.383	0.317	0.330	0.013	0.005	-0.005	-0.005	0.460
WLS																	
VAR																	
A	-0.242	-0.274	0.033	0.003	0.011	0.219	-0.222	-0.234	0.012	0.100	-0.240	-0.256	0.016	0.011	-0.018	-0.018	0.053
Δ	0.294	0.312	0.018	0.028	0.001	0.921	0.296	0.284	-0.012	0.028	0.293	0.311	0.018	0.000	-0.001	-0.001	0.874
WVAR																	

Table continued on the following page.

Table 4.22: Continued from previous page

	Deletion																	
	VORANK = 1 (142)			VORANK = 2 (184)			VORANK = 3 (190)			VORANK = 4 (167)			VORANK 4 - 1					
	GARCH																	
A	-0.607	-0.655	0.048	0.099	-0.570	-0.569	-0.001	0.961	-0.559	-0.586	0.027	0.527	-0.554	-0.570	0.016	0.039	-0.085	0.000
Δ	0.358	0.383	0.024	0.056	0.381	0.383	0.003	0.838	0.361	0.369	0.007	0.602	0.367	0.392	0.025	0.000	0.009	0.400
	BEKK																	
A	-0.477	-0.542	0.066	0.000	-0.480	-0.486	0.006	0.578	-0.494	-0.476	-0.018	0.035	-0.522	-0.538	0.016	0.039	-0.005	0.654
Δ	0.351	0.340	-0.011	0.192	0.356	0.363	0.007	0.352	0.341	0.342	0.001	0.822	0.350	0.351	0.002	0.754	0.011	0.128

Table 4.23: P-Value for Tests of Model Parameters - SIRANK (Addition)

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the addition events are computed as described in Section 2.3.3 in four short interest quartiles. The cell entries of the first panel show the number of observations in each short interest quartile. The cell entries of the second panel show the estimated parameters and probabilities (p-val.), respectively. Differences for each short interest quartile are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between the highest and the lowest short interest quartiles after the addition events are calculated by absolute value of model parameters after the event between the largest and the smallest short interest quartiles (Highest minus Lowest). For each addition event, stocks are ranked by their average daily short interest defined as the ratio of daily short volume to the total shares outstanding in the following 60 days after the event date. SIRANK = 1 indicates the lowest short interest ratio and SIRANK = 4 indicates the highest short interest ratio. Sums of A and Δ are the average within each short interest quartile. Bold format denotes significance at the 1% level.

	Addition																		
	SIRANK = 1 (144)			SIRANK = 2 (160)			SIRANK = 3 (177)			SIRANK = 4 (158)			SIRANK 4 - 1						
	Before	After	p-val.	Before	After	p-val.	Before	After	p-val.	Before	After	p-val.	Diff. in After	p-val.					
A	-0.455	-0.457	0.002	0.867	-0.433	-0.411	-0.022	0.062	-0.407	-0.398	-0.009	0.254	-0.394	-0.363	-0.031	0.000	-0.094	0.000	
Δ	0.331	0.321	-0.010	0.203	0.315	0.311	-0.004	0.533	0.323	0.311	-0.012	0.008	0.310	0.304	-0.006	0.233	-0.017	0.012	
	WLS																		
A	-0.237	-0.221	-0.016	0.186	-0.238	-0.214	-0.024	0.018	-0.228	-0.220	-0.007	0.319	-0.232	-0.197	-0.035	0.022	-0.024	0.159	
Δ	0.310	0.288	-0.022	0.024	0.297	0.295	-0.002	0.740	0.315	0.298	-0.017	0.009	0.304	0.301	-0.003	0.613	0.013	0.108	
	VAR																		
A	-0.424	-0.404	-0.020	0.001	-0.405	-0.393	-0.012	0.016	-0.390	-0.385	-0.005	0.165	-0.377	-0.345	-0.032	0.000	-0.059	0.000	
Δ	0.324	0.313	-0.011	0.009	0.310	0.305	-0.006	0.125	0.318	0.307	-0.011	0.000	0.307	0.299	-0.008	0.005	-0.013	0.000	
	WVAR																		
A	-0.225	-0.204	-0.021	0.000	-0.224	-0.211	-0.014	0.003	-0.222	-0.221	-0.001	0.757	-0.223	-0.203	-0.020	0.000	-0.001	0.808	
Δ	0.305	0.284	-0.021	0.000	0.296	0.292	-0.003	0.445	0.310	0.295	-0.014	0.000	0.302	0.296	-0.006	0.045	0.012	0.002	

Table continued on the following page.

Table 4.23: Continued from previous page

	SIRANK = 1 (144)		SIRANK = 2 (160)		SIRANK = 3 (177)		SIRANK = 4 (158)		SIRANK 4 - 1									
	Addition																	
	GARCH					BEKK												
A	-0.494	-0.506	0.011	0.094	-0.460	-0.447	-0.014	0.005	-0.445	-0.419	-0.026	0.000	-0.430	-0.394	-0.036	0.000	-0.112	0.000
Δ	0.400	0.363	-0.037	0.000	0.368	0.364	-0.004	0.360	0.376	0.368	-0.008	0.014	0.363	0.365	0.001	0.730	0.002	0.738
A	-0.451	-0.451	0.000	0.971	-0.428	-0.396	-0.032	0.000	-0.415	-0.382	-0.033	0.000	-0.409	-0.375	-0.035	0.000	-0.077	0.000
Δ	0.378	0.351	-0.027	0.000	0.364	0.330	-0.034	0.000	0.371	0.357	-0.014	0.000	0.350	0.338	-0.012	0.000	-0.013	0.005

Table 4.24: P-Value for Tests of Model Parameters - SIRANK (Deletion)

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the deletion events are computed as described in Section 2.3.3 in four short interest quartiles. The cell entries of the first panel show the number of observations in each short interest quartile. The cell entries of the second panel show the estimated parameters and probabilities (p-val.), respectively. Differences for each short interest quartile are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between the highest and the lowest short interest quartiles after the deletion events are calculated by absolute value of model parameters after the event between the largest and the smallest short interest quartiles (Highest minus Lowest). For each deletion event, stocks are ranked by their average daily short interest defined as the ratio of daily short volume to the total shares outstanding in the following 60 days after the event date. SIRANK = 1 indicates the lowest short interest ratio and SIRANK = 4 indicates the highest short interest ratio. Sums of A and Δ are the average within each short interest quartile. Bold format denotes significance at the 1% level.

		Deletion																	
		SIRANK = 1 (58)			SIRANK = 2 (70)			SIRANK = 3 (74)			SIRANK = 4 (62)			SIRANK 4 - 1					
		Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.		
A	-0.534	-0.605	0.071	0.035	0.035	-0.576	-0.584	0.008	0.815	-0.488	-0.506	0.017	0.342	-0.526	-0.519	-0.007	0.752	-0.008	0.794
Δ	0.321	0.341	0.021	0.208	0.208	0.312	0.334	0.022	0.154	0.317	0.308	-0.010	0.325	0.324	0.330	0.006	0.556	0.003	0.847
		WLS																	
A	-0.246	-0.298	0.053	0.049	0.049	-0.277	-0.235	-0.042	0.052	-0.249	-0.251	0.002	0.907	-0.257	-0.253	-0.005	0.795	0.012	0.614
Δ	0.307	0.311	0.004	0.848	0.848	0.289	0.247	-0.042	0.069	0.304	0.305	0.001	0.939	0.298	0.300	0.003	0.848	-0.009	0.594
		VAR																	
A	-0.547	-0.512	-0.035	0.013	0.013	-0.497	-0.516	0.019	0.176	-0.449	-0.461	0.011	0.180	-0.485	-0.472	-0.013	0.178	-0.063	0.000
Δ	0.294	0.330	0.036	0.000	0.000	0.294	0.303	0.009	0.329	0.313	0.299	-0.014	0.014	0.317	0.325	0.007	0.249	0.023	0.003
		WVAR																	
A	-0.209	-0.279	0.070	0.000	0.000	-0.251	-0.244	-0.006	0.599	-0.238	-0.239	0.001	0.901	-0.241	-0.231	-0.010	0.236	0.031	0.002
Δ	0.294	0.303	0.008	0.369	0.369	0.279	0.274	-0.005	0.576	0.302	0.294	-0.008	0.145	0.291	0.299	0.008	0.210	-0.004	0.635

Table continued on the following page.

Table 4.24: Continued from previous page

	SIRANK = 1 (58)		SIRANK = 2 (70)		Deletion		SIRANK = 3 (74)		SIRANK = 4 (62)		SIRANK 4 - 1							
A	-0.567	-0.572	0.006	0.740	-0.593	-0.610	0.017	0.531	-0.523	-0.566	0.043	0.000	-0.534	-0.573	0.039	0.000	-0.033	0.004
Δ	0.372	0.406	0.034	0.003	0.339	0.399	0.060	0.000	0.363	0.350	-0.013	0.060	0.380	0.370	-0.010	0.229	0.008	0.403
									BEKK									
A	-0.474	-0.542	0.068	0.000	-0.513	-0.493	-0.021	0.160	-0.491	-0.514	0.023	0.017	-0.477	-0.500	0.023	0.043	0.003	0.825
Δ	0.393	0.358	-0.036	0.000	0.320	0.347	0.027	0.006	0.335	0.321	-0.014	0.035	0.347	0.310	-0.037	0.000	-0.047	0.000

Table 4.25: P-Value for Tests of Model Parameters - SFRANK (Addition)

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the addition events are computed as described in Section 2.3.3 in four shorting flow quartiles. The cell entries of the first panel show the number of observations in each shorting flow quartile. The cell entries of the second panel show the estimated parameters and probabilities (p-val.), respectively. Differences for each shorting flow quartile are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between the highest and the lowest shorting flow quartiles after the addition events are calculated by absolute value of model parameters after the event between the largest and the smallest shorting flow quartiles (Highest minus Lowest). For each addition event, stocks are ranked by their average daily shorting flow defined as the ratio of daily short volume to daily trading volume in the following 60 days after the event date. SFRANK = 1 indicates the lowest shorting flow and SFRANK = 4 indicates the highest shorting flow. Sums of A and Δ are the average within each shorting flow quartile. Bold format denotes significance at the 1% level.

		Addition																	
		SFRANK = 1 (148)			SFRANK = 2 (157)			SFRANK = 3 (169)			SFRANK = 4 (155)			SFRANK 4 - 1					
		Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.	Before	After	Diff.	p-val.		
A		-0.448	-0.456	0.008	0.401	-0.420	-0.416	-0.005	0.663	-0.417	-0.384	-0.033	0.003	-0.402	-0.373	-0.030	0.009	-0.083	0.000
Δ		0.333	0.324	-0.009	0.131	0.320	0.303	-0.017	0.003	0.316	0.315	-0.001	0.845	0.310	0.304	-0.006	0.374	-0.020	0.002
WLS																			
A		-0.241	-0.241	0.000	0.961	-0.246	-0.224	-0.022	0.023	-0.228	-0.208	-0.020	0.025	-0.219	-0.181	-0.038	0.020	-0.060	0.000
Δ		0.323	0.309	-0.015	0.050	0.306	0.288	-0.018	0.008	0.297	0.300	0.003	0.683	0.301	0.287	-0.014	0.059	-0.022	0.006
VAR																			
A		-0.428	-0.424	-0.004	0.406	-0.396	-0.401	0.005	0.219	-0.390	-0.360	-0.029	0.000	-0.383	-0.343	-0.040	0.000	-0.081	0.000
Δ		0.331	0.320	-0.010	0.002	0.315	0.298	-0.017	0.000	0.312	0.309	-0.003	0.379	0.303	0.296	-0.006	0.077	-0.024	0.000
WVAR																			
A		-0.238	-0.230	-0.007	0.113	-0.234	-0.223	-0.011	0.008	-0.214	-0.198	-0.016	0.000	-0.211	-0.191	-0.020	0.000	-0.039	0.000
Δ		0.323	0.308	-0.015	0.000	0.303	0.282	-0.021	0.000	0.293	0.298	0.004	0.312	0.296	0.282	-0.014	0.000	-0.026	0.000

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Table 4.25: Continued from previous page

	SFRANK = 1 (148)		SFRANK = 2 (157)		Addition SFRANK = 3 (169)		SFRANK = 4 (155)		SFRANK 4 - 1									
A	-0.491	-0.494	0.002	0.654	-0.454	-0.453	-0.002	0.708	-0.455	-0.407	-0.048	0.000	-0.428	-0.411	-0.017	0.001	-0.083	0.000
Δ	0.395	0.368	-0.026	0.000	0.377	0.365	-0.013	0.001	0.375	0.369	-0.006	0.177	0.361	0.358	-0.002	0.625	-0.010	0.033
									BEKK									
A	-0.468	-0.447	-0.021	0.000	-0.426	-0.416	-0.010	0.061	-0.415	-0.379	-0.036	0.000	-0.394	-0.360	-0.034	0.000	-0.087	0.000
Δ	0.380	0.342	-0.037	0.000	0.377	0.356	-0.022	0.000	0.359	0.342	-0.017	0.000	0.346	0.336	-0.011	0.010	-0.007	0.127

Table 4.26: P-Value for Tests of Model Parameters - SFRANK (Deletion)

Based on the models of OLS, WLS, VAR, WVAR, GARCH (1, 1) and BEKK (1, 1) using 60 days of trade data, the tests for sums of A and Δ measuring speed of price adjustment for the deletion events are computed as described in Section 2.3.3 in four shorting flow quartiles. The cell entries of the first panel show the number of observations in each shorting flow quartile. The cell entries of the second panel show the estimated parameters and probabilities (p-val.), respectively. Differences for each shorting flow quartile are calculated by absolute value of model parameters between after and before for the event (After minus Before). Differences between the highest and the lowest shorting flow quartiles after the deletion events are calculated by absolute value of model parameters after the event between the largest and the smallest shorting flow quartiles (Highest minus Lowest). For each deletion event, stocks are ranked by their average daily shorting flow defined as the ratio of daily short volume to daily trading volume in the following 60 days after the event date. SFRANK = 1 indicates the lowest shorting flow and SFRANK = 4 indicates the highest shorting flow. Sums of A and Δ are the average within each shorting flow quartile. Bold format denotes significance at the 1% level.

	Deletion																	
	SFRANK = 1 (56)			SFRANK = 2 (71)			SFRANK = 3 (74)			SFRANK = 4 (63)			SFRANK 4 - 1					
	Before	After	Diff.	Before	After	Diff.	Before	After	Diff.	Before	After	Diff.	Before	After	Diff.	p-val.		
A	-0.558	-0.629	0.071	0.017	-0.509	-0.523	0.014	0.536	-0.541	-0.549	0.008	0.744	-0.519	-0.518	-0.001	0.976	-0.039	0.190
Δ	0.333	0.331	-0.001	0.929	0.318	0.339	0.021	0.055	0.318	0.323	0.005	0.722	0.304	0.315	0.011	0.415	-0.028	0.048
	OLS																	
A	-0.269	-0.276	0.008	0.712	-0.255	-0.259	0.004	0.820	-0.273	-0.272	-0.002	0.945	-0.233	-0.223	-0.010	0.652	-0.036	0.108
Δ	0.303	0.297	-0.006	0.715	0.306	0.304	-0.002	0.869	0.311	0.298	-0.012	0.495	0.275	0.258	-0.017	0.405	-0.027	0.107
	WLS																	
	VAR																	
A	-0.589	-0.541	-0.049	0.000	-0.452	-0.471	0.020	0.048	-0.478	-0.500	0.021	0.072	-0.467	-0.451	-0.016	0.177	-0.123	0.000
Δ	0.314	0.318	0.004	0.644	0.310	0.326	0.016	0.018	0.302	0.303	0.002	0.819	0.294	0.304	0.010	0.216	-0.020	0.006
	WVAR																	
A	-0.245	-0.260	0.015	0.155	-0.252	-0.248	-0.004	0.667	-0.232	-0.265	0.033	0.001	-0.214	-0.214	0.000	0.978	-0.031	0.001
Δ	0.298	0.298	0.000	0.953	0.303	0.296	-0.007	0.304	0.294	0.290	-0.004	0.575	0.270	0.282	0.013	0.116	-0.028	0.000

Table continued on the following page.

Table 4.26: Continued from previous page

	SFRANK = 1 (56)		SFRANK = 2 (71)		SFRANK = 3 (74)		SFRANK = 4 (63)		SFRANK 4 - 1									
	Deletion																	
	GARCH					BEKK												
A	-0.569	-0.636	0.067	0.000	-0.562	-0.549	-0.013	0.165	-0.572	-0.607	0.034	0.005	-0.508	-0.536	0.028	0.317	-0.061	0.000
Δ	0.368	0.384	0.016	0.100	0.351	0.402	0.052	0.000	0.362	0.365	0.002	0.841	0.370	0.369	-0.002	0.894	0.002	0.789
A	-0.535	-0.563	0.028	0.045	-0.466	-0.515	0.049	0.000	-0.501	-0.492	-0.009	0.452	-0.463	-0.482	0.019	0.171	-0.072	0.000
Δ	0.359	0.340	-0.018	0.045	0.355	0.344	-0.010	0.167	0.338	0.329	-0.008	0.320	0.338	0.320	-0.018	0.046	-0.021	0.013

Appendix F

Other Table

**Table F.1: Cumulative Abnormal Returns
Before, During and After Crisis.**

The table reports the cumulative abnormal returns based on the market model estimated before, during and after the crisis. The period of crisis is defined as the period from December 2007 to June 2009. For the market model, an estimation window of $(-240, -60)$, with a minimum length of 180 days, is used. Each panel of the table indicates that the minimum length of time for stocks to be added back after deletion from the D-list.

Event Windows	Cumulative Daily Abnormal Returns (%)					
	Before Crisis		During Crisis		After Crisis	
	Mean	p-value	Mean	p-value	Mean	p-value
	1-Year					
(0, 10)	2.24	0.01	1.97	0.15	-2.74	0.03
(0, 30)	2.62	0.08	8.35	0.00	-4.20	0.09
(0, 60)	4.34	0.06	18.58	0.00	-3.60	0.32
	2-Year					
(0, 10)	1.74	0.03	0.89	0.40	-3.22	0.04
(0, 30)	2.30	0.13	6.77	0.00	-5.47	0.06
(0, 60)	5.10	0.03	17.66	0.00	-5.68	0.18
	3-Year					
(0, 10)	1.52	0.08	0.45	0.63	-4.09	0.06
(0, 30)	2.78	0.07	5.17	0.00	-7.03	0.04
(0, 60)	5.26	0.05	15.92	0.00	-8.17	0.11

Chapter 5

Conclusions

This thesis investigates whether short sales in the Hong Kong stock market aid price efficiency as measured as the speed of price adjustment to new information using ultra-high frequency trade-by-trade data over 10 years from 2001 to 2011. The selection criteria for short sales in Hong Kong provides a unique opportunity to explore the impact of short sales on price efficiency as the market only allows stocks which meet certain requirements to be added to the list for short selling. Furthermore, this list is revised by the stock exchange every three months. This setting ensures that any differences in the speed of price adjustment in the three months before and after each addition and deletion event can be directly linked to the changes in the eligibility for short sales. The thesis addresses three distinct but related topics. In what follows, it presents a brief summary of the empirical findings and highlights potential areas for future research.

5.1 Summary of Findings

Chapter 2 extends the bivariate VAR model of [Hasbrouck \(1991\)](#) and offers a direct examination of the effect of short sales on price efficiency in the microstructure model for changes in quotes and trade dynamics. In addition to the use of least squares and weighted least squares on an equation by equation basis, a *bona fide*

VAR model is estimated simultaneously with consideration of the cross-sectional correlation of the residuals in the two equations. Therefore, four different estimation methods namely OLS, WLS, VAR and WVAR are included in this chapter.

The main findings of Chapter 2 are as follows. First, there is a subset of stocks with significant changes including increases and decreases in model parameters as the proxies for price efficiency for both addition and deletion events. Based on the results under four estimation models, it is noted that there are more significant changes captured in individual and aggregated parameters during both events when the models are either estimated simultaneously or without consideration of time duration of trades. In other words, it indicates that the VAR and WVAR models are more sensitive in capturing significant changes than their least squares counterparts while time-weighted models (WLS and WVAR) are more conservative.

Secondly, the study of parameter dynamics under four models show that the model parameters, except for those measuring Granger causality, largely remain consistent during both events. It shows that positive/negative parameters which are statistically significant will generally stay in the same significance category with the same sign, even if the estimated values change as a result of the event. Furthermore, it is found that the model parameters are less consistent under the time-weighted models and the consistency is greater for the addition stocks rather than the deletion stocks.

Thirdly, the chapter carries out an investigation on the overall effect of short sales on price efficiency as the significant changes captured in parameters include both increased and decreased values. The Z scores for tests of differences in model parameters show that the autocorrelations in both quote returns and trades decrease significantly during the addition events and the autocorrelations in quote returns increase significantly during the deletion events. The overall results for both events suggest that the speed of price adjustment is faster with the eligibility of shorts sales as a smaller absolute value of autocorrelation in quote returns and trades indicate that stock prices follow a random walk more closely. It is also found that

trade continuity (autocorrelation in quote returns) is affected the most during the addition (deletion) events under four different models. Furthermore, the price impact of trades get stronger when stocks are added to the D-list for short selling under models when the time duration between consecutive transactions is considered.

Fourthly, this chapter also provides some indications of the best model based on the statistical tests on residual correlations and a summary of model estimations under OLS, WLS, VAR and WVAR. Based on residual correlation tests, VAR and WVAR models have better ability to capture residual correlations between quote and trade equations than OLS and WLS models that are estimated equation by equation. Based on the summary of model estimations, more significant improvements (deteriorations) in price efficiency are captured for the addition (deletion) events under VAR and WVAR models than OLS and WLS models. It suggests that VAR and WVAR models are more powerful to explore the differences in the speed of price adjustment during both events. There is no direct standard test to the best of our knowledge to compare between models with and without time-weighted variables. Therefore, the WVAR model is selected to be the preferred model as it is more rigorous theoretically because trades do not occur at regularly spaced time intervals.

The ultra-high frequency based quotations and signed volumes invariably exhibit conditional heteroscedasticity and it is conjectured that heterogeneous variances and covariances could have an effect on the model parameters. Chapter 3 extends the VAR model by conducting the analysis using GARCH and BEKK models. Univariate GARCH (1, 1) models capture dynamic volatility in each equation while the BEKK (1, 1) model considers the co-movement and non-independence of volatility between quote returns and trades.

The findings of Chapter 3 are summarised as follows. First, a number of stocks exhibit significant changes for addition and deletion events under GARCH and BEKK models. Compared with the results from OLS and VAR models, there are

more significant changes captured in all corresponding parameters under the models with consideration of time-varying variances and covariances. It reveals that the models taking conditional heteroscedasticity into account have a deeper insight into the effect of short sales on the model parameters for both events. However, there are still less than 40% (30%) of the stocks affected significantly by changes in the parameters measuring the speed of price adjustment during the addition (deletion) events. Taken together, it means that less than half of the stocks contribute to the overall changes in the speed of price adjustment under all estimation models.

Secondly, the study of parameter dynamics shows that the majority of the model parameters, except for Granger causality, are consistent when stocks change their eligibility of short sales. Moreover, it is interesting to note that the parameter consistency is stronger especially for the autocorrelations in quote returns and trades (the sums, A and Δ) under GARCH model rather than OLS while the consistency is qualitatively the similar between VAR and BEKK.

Thirdly, the overall results from Z scores for tests of difference in model parameters for both events show that gaining (losing) the eligibility of short sales accelerate (hinder) the speed of price adjustment by decreasing (increasing) the autocorrelation in the parameters measuring price efficiency. The findings are consistent with those from Chapter 2. The combined results from Chapters 2 and 3 indicate that trade continuity decreases the most by short sales with robustness under all estimation models. Quote returns are less correlated at significance level 1% under both WVAR and BEKK during the addition events, this finding being consistent with [Boehmer & Wu \(2013\)](#)'s prediction that an efficient price process is associated with a random walk if the quotations are the best estimate of the equilibrium value of the stocks in the market. For the deletion events, autocorrelations in quote returns is strengthened the most and stronger stickiness in quote returns indicates the delayed price incorporation to new information when stocks are removed from the D-list. Trade continuity is not affected significantly during the deletion events. In sum, the results support that the efficiency-enhancing effect of short sales is economically meaningful. Lastly, it is also noted that the models estimated simultaneously or with consideration of heterogeneous variances, that is WVAR and

BEKK models, are more powerful to capture the differences in the speed of price adjustment especially during the addition events.

Chapter 4 carries on an investigation of the determinants of the speed of price adjustment including market conditions, firm size, trading volume and short sellers' activities during both addition and deletion events.

The results of Chapter 4 are as follows. To explore the speed of price adjustment under various market conditions during both events, the chapter introduces three different methods: models with interaction variables for current and one-day lagged market return; comparison between crisis and non-crisis periods; re-estimation based on the positive and non-positive transaction days, respectively. The overall results show that short sales speed up the price adjustment regardless of market conditions. Stocks have a faster speed of incorporation of new information into prices in a down or bad market rather than in an up or good market. By comparing the crisis and the non-crisis periods, it is found that short sales aid price efficiency in the non-crisis period whereas the effect of short sales is unclear during in the crisis.

By dividing the estimation results from Chapters 2 & 3 into four-sized groups based on their 60-day average market value prior to the effective date, it is found that bigger-sized firms have faster speed of price adjustment to new information than smaller-sized ones and the results are robust under all estimation models. It is also interesting to note that the medium-sized (small-sized) firms are the most affected when they are added to (removed from) the D-list for short sales. Comparing stocks with different trading volumes, the results show that stocks with higher trading volume are more likely to be influenced by short sales during the addition events while stocks from the top and the bottom are impacted most during the deletion events. Stocks with more trading volumes are faster to adjust their prices compared with those with lower trading volume. Two proxies for short sales employed to investigate the role of short sellers' activities in the process of price discovery. The findings show that in general stocks with higher short interest have

better price efficiency especially under the models without time durations while stocks with more shorting flow experience a faster speed of price adjustment to new information. These two measures reveal that short sellers contribute significantly to an efficient price discovery and the findings support the hypothesis that short sellers are informed traders with adverse information which can help to correct prices back to their fundamental levels (Boehmer *et al.*, 2008; Desai *et al.*, 2002; Diether *et al.*, 2009).

This thesis investigates the role of short sales in the process of price discovery in the Hong Kong market where the institutional environment provides a rare opportunity with a designated list for short sales. There are 808 stocks with changes of eligibility of short sales over a 10-year period. Among them, there are 72 H-shares and 80 red chips stocks, totally 152 out of 808 stocks (approximately 20%) with mainland Chinese background. H shares refer to the shares of companies incorporated in mainland China that are traded on the HKEx. Red chips stocks are the stocks of mainland China companies incorporated outside mainland China and listed on the HKEx. The businesses of red chips stocks are based in mainland China and controlled, either directly or indirectly, by the central, provincial or municipal governments of China but listed in Hong Kong to allow overseas investment in the companies.

The China Securities Regulatory Commission (CSRC) introduced a pilot program of margin trading in 2010, allowing 90 stocks quoted on either the Shanghai or Shenzhen stock exchanges on a designated list to be sold short and/or purchased on margin. Since the end of the pilot programme, over 800 quoted Chinese stocks have become eligible for short selling in both Chinese stock markets. Given 20% of event stocks with mainland Chinese background, results from the thesis may provide information about the possible impact of short sales on price efficiency in mainland China and provide guidance for policy making for Chinese regulators to facilitate the further development of the market.

5.2 Areas for Future Research

This study has focused on the effect of short selling on price efficiency. The lifting of short sales constraints contributes to the speed of price adjustment by the participation of pessimistic informed investors. The resolution of uncertainty about the fundamentals tends to improve stock liquidity by decreasing the bid-ask spread (Diamond & Verrecchia, 1987). Short sellers could, however, also impact stock liquidity in the opposite way. Short sellers will lower stock liquidity because they, as informed traders, are more likely to be liquidity demanders rather than liquidity suppliers in the market (Boehmer *et al.*, 2008). Moreover, a large volume of informed trading based on private information will force market makers to increase the adverse selection cost of bid-ask spread for compensation, thus resulting in the reduction of stock liquidity (Amihud & Mendelson, 1986). Future research can be conducted on the effect of short sales on liquidity, based upon the models developed in this thesis. Further research to identify more detailed explanations for deletion events during the financial crisis period would also be of interest. This is because the initial investigations reported in Chapter 4 suggest that during the crisis deletion events served as a proxy for a partial short sales ban and so serve a genuine economic function.

This thesis makes several contributions to both the existing literature and financial practice. It also has some limitations which are directions for future research.

First, the dataset used for this study is from May 2001 to May 2011. Although a 10-year intraday dataset is sufficient to examine the effect of short sales on the speed of price adjustment in detail, a more recent dataset would add to the robustness of the research. Specifically, it would permit the investigation of (*i*) possible parameter changes and (*ii*) the extent to which the models developed in this thesis could be used for forecasting purposes.

Secondly, in Chapter 4, the Hang Seng daily market index is used to construct the dummy variable of market condition in the VAR model interacted with model

parameters to investigate the process of price discovery under different market conditions. As discussed in that chapter, the dummy variable of current market return used in the model raises foresight issues as quotations and trades at an arbitrary time of the day are associated with the dummy variable calculated by the price at the end of the day. Although the dummy variable of one-day lagged market return is employed to solve this issue, high frequency market index data would be more appropriate to use with quotes and trades in the model.

Thirdly, the models use a variable based on price quotations. This is similar to returns defined in the conventional manner but is not exactly the same. From a theoretical perspective it would be instructive to determine to what extent these non-standard returns are consistent with standard asset pricing models.

Lastly, results from Chapters 2 & 3 show that the WVAR and BEKK models are more powerful to capture the changes in the speed of price adjustment compared with other models. However, there is no clear evidence to support the exclusive use of one model in preference to another. Furthermore, particularly at the aggregate level, the differences between all six models reported in this thesis are not great. An updated data set would facilitate further analysis of model differences. As noted in Chapter 4, there is a need for development of model software. As noted in Chapter 2, there is a need for development of statistical tests. Distinguishing the differences between the two estimation methods in the microstructure models can contribute to model selection for high frequency trading strategies in future research.

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